U.S. Department of Transportation National Highway Traffic Safety Administration

# Meta-Analysis of Graduated Driver Licensing Laws 

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## 16. Abstract

The objective of the present study was to assess the effectiveness of GDL programs for reducing total, injury, and fatal crashes among drivers 15 to 20 years old by conducting a meta-analysis of GDL research since 2001 that evaluated the effectiveness of GDL systems overall and GDL components individually. The final sample of 14 selected studies represented 13 different States, and three represented GDL programs across most or all U.S. States. Results of the meta-analysis showed that GDL programs as a whole were associated with statistically reliable reductions in traffic crashes outcomes of 16 percent for 16 -year-olds and 11 percent for 17-year-olds, but were not reliably associated with changes in crash outcomes for 18- or 19-year-olds. Unfortunately, the numbers of effect sizes representing the unique effects of individual GDL components and calibrations were small for most of the components, particularly when stratified by the ranges of possible calibrations for those components. Although the exact effectiveness of individual GDL provisions could not be determined, the meta-analysis uncovered no indication that any provision was necessarily counterproductive for the GDL target audience of 16- and 17-year olds. Thus, a reasonable strategy for any State considering passage of a GDL law might involve enumerating the full range of provisions applicable to that State, determining which could be reasonably operationalized given available resources and support from key agencies and organizations, and adopting as comprehensive an approach as possible.

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## EXECUTIVE SUMMARY

Motor vehicle crashes continue to be the leading cause of death for 15- to 20-year-olds (Hoyert \& Jiaquan, 2012). Graduated driver licensing (GDL) programs, which are specialized driver licensing systems for beginner drivers, have been implemented in United States to reduce young teen drivers' exposure to high-risk driving situations while they gain driving experience. GDL programs include three different stages of licensure: (a) a mandatory minimum learner permit period during which new drivers are only allowed to drive under the supervision of a licensed adult; (b) an intermediate period during which the new drivers are allowed to drive unsupervised, but are subject to licensing restrictions regarding passenger ages and the times during which they may drive; and (c) a final stage of unrestricted licensure allowing driving under all conditions. GDL programs in the U.S. hardly represent a single homogeneous intervention; rather, programs vary characteristics such as age and time criteria, lengths of the learner permit and restricted license stages, required hours of supervised practice, and types and lengths of license restrictions included (IIHS, 2012).

While there is a growing body of evidence that supports GDL systems as effective for reducing young driver crashes, little is collectively known about which specific characteristics or provisions of GDL programs, such as minimum learner permit holding periods, and what parameters or calibrations of the provisions are associated with the largest crash reductions. Thus, the objective of this study was to conduct a meta-analysis to systematically synthesize research findings regarding the effectiveness of GDL programs and varied components (e.g., learner entry ages, nighttime driving restrictions) for reducing total, injury, and fatal crashes among drivers aged 15 to 20 years.

Fourteen studies were included in the meta-analysis and the findings suggest that GDL laws create a safety benefit for 16-year-old drivers and potentially have a safety benefit for 17-year-old drivers, although to a lesser extent. Given the diversity in the configuration of GDL provisions among the States, the individual contributions of these characteristics to the overall observed effect was also of interest, however, insufficient studies with suitable information available existed to answer this question. Further, a valid study of individual GDL provisions may be prohibited by practical limitations on the level of experimental control one could obtain. For example, the mere existence of a curfew provision does not mean it was truly operationalized unless it can be shown the affected population was aware of it, adhered to it, and the police actually enforced it at a meaningful level. Few studies quantified these important process factors.

Although the exact effectiveness of individual GDL provisions could not be determined, the meta-analysis uncovered no indication that any provision was necessarily counterproductive for the GDL target audience of 16- and 17-year olds. Thus, a reasonable strategy for any State considering passage of a GDL law might involve enumerating the full range of provisions applicable to that State, determining which could be reasonably operationalized given available resources and support from key agencies and organizations, and adopting as comprehensive an approach as possible.

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## 1. BACKGROUND

Motor vehicle crashes continue to be the leading cause of death for 15- to 20-year-olds (Hoyert \& Jiaquan, 2012). Graduated driver licensing (GDL) programs, specialized driver licensing systems for novice drivers, have been implemented in U.S. States to reduce young teen drivers' exposure to high-risk driving situations while they gain driving experience. GDL programs include three different stages of licensure: (a) a mandatory minimum learner permit period during which new drivers are only allowed to drive under the supervision of a licensed adult; (b) an intermediate period during which the new drivers are allowed to drive unsupervised, but are subject to licensing restrictions regarding passenger ages and the times during which they may drive; and (c) a final stage of unrestricted licensure allowing driving under all conditions. As novice drivers systematically move through these stages and gain more on-road driving experience, the restrictions that limit their exposure to risky driving conditions are gradually removed.

The seven components generally considered to be the core of GDL programs are: (1) learner stage minimum entry ages, (2) required learner permit minimum holding time periods, (3) numbers of required supervised driving hours, (4) minimum intermediate licensing ages, (5) intermediate licensing stage nighttime driving restrictions, (6) intermediate licensing stage passenger driving restrictions, and (7) minimum unrestricted licensing ages. Another less recognized component is the requirement to maintain a crash- and conviction-free driving record in order to progress through the GDL stages, sometimes referred to as "contingent advancement." Various combinations of these licensing components and quantitative/qualitative differences in how they are applied (i.e., calibrated) form the teen driver licensing systems in every U.S. State. GDL programs in the United States hardly represent a single homogeneous intervention. Rather there are many different GDL programs that vary program components including age and time criteria, lengths of the learner permit and restricted license stages, required hours of supervised practice, and types and lengths of license restrictions included (IIHS, 2012).

The overwhelming majority of single-State (i.e., one State or province) GDL studies have found positive safety effects for young teens ranging from 20-40 percent reduction in crash rates (Senserrick \& Haworth, 2005; Shope, 2007). Several multi-State studies of GDL programs have been completed to obtain more global estimates of GDL effectiveness and to avoid some problems associated with limited generalizability of heterogeneous single-State studies by using centralized crash data sources (Chen, Baker, \& Li, 2006; Dee, Grabowski, \& Morrisey, 2005; McCartt, Teoh, Fields, Braitman, \& Hellinga, 2010; Vanlaar et al., 2009). The results from these multi-jurisdiction studies of fatal crashes generally suggest smaller effects that range from 6-19 percent reduction in crash rates. It should be noted that many of these multi-State studies of overall GDL program effects and GDL program components have significant limitations. Particularly, the main criticisms of these studies include weak study design, lack of experimental control over confounding factors, findings subject to statistical artifacts, and findings that appear impractical.

Research has compared crash rates between States with and without GDL showing support for GDL. Specifically, Chen, Baker, and Li (2006) found that the 16-year-old driver-
involved fatal crash rate in States with GDL was 11 percent lower than that compared to States without GDL. Additionally, this study found the 16 -year-old driver-involved fatal crash rate to be 16-21 percent lower in States that included (1) a minimum waiting period of at least three months following attainment of a learner permit, (2) a nighttime restriction, and (3) either a passenger restriction ( $21 \%$ lower), 30 hours of supervised driving in the learner period ( $18 \%$ lower), or both (16\% lower). Additionally, McCartt, Teoh, Fields, Braitman, and Hellinga (2010) conducted a national study of the effectiveness of U.S. GDL laws for reducing fatal crashes among 15-, 16-, and 17-year-olds employing Fatality Analysis Reporting System (FARS) crash data from 1996-2007. The evaluation found the following associations when comparing States with and without certain provisions: (1) 7 percent lower fatal crash rate when licensing was delayed by 6 months and 13 percent lower when delayed by 1 year; (2) 21 percent lower fatal crash rate when no teenage passengers allowed and 7 percent lower with one teenage passenger allowed compared to no passenger restriction; and (3) 20 percent lower fatal crash rate when driving was restricted after 8 pm , 18 percent after $9 \mathrm{pm}, 16$ percent after 10 pm , 14 percent after $11 \mathrm{pm}, 12$ percent after 12 am , and 9 percent after 1 am . A similar effort conducted by Vanlaar et al. (2009) measured changes in fatal crash rates after GDL implementation by analyzing 19922006 data from FARS and the Traffic Accident Information Database (Canadian fatality data). After adjusting for the crash trend of an older age group of drivers, they found a 19.1 percent reduction in 16 -year-old driver fatal crash rates after GDL system implementation. Vanlaar et al. (2009) were also interested in evaluating the contributions of individual provisions. Most notably, the study reported an 88 percent reduction in 16-year-old relative fatal crash risk when passenger restrictions were present during intermediate stages, which is surprisingly large and warrants further examination.

In 2010, the U.S. Government Accountability Office (GAO) issued a report of the National Highway Traffic Safety Administration’s (NHTSA's) teen driver safety programs and research. The report stated that "gaps still exist in researching the effectiveness of specific GDL provisions including specific provisions for minimum entry age, the learner’s permit stage, nighttime and passenger restrictions, bans on electronic devices, driver education, and parental involvement" (U.S. GAO, 2010). The GAO recommended that NHTSA conduct research to fill these gaps in the literature and better inform States regarding specific provision effectiveness. The practical limitations of studying the specific provisions of GDL programs (e.g., inability to ensure implementations in real-world setting) restrict the methods that may be implemented by researchers. Thus, a systematic synthesis of the research findings available in the literature is an appropriate method to yield estimates of GDL program and specific provisions effectiveness in reducing young driver fatal crash rates.

## 2. OBJECTIVE

The objective of the present study was to assess the effectiveness of GDL programs for reducing total, injury, and fatal crash rates among drivers 15 to 20 years old by conducting a meta-analysis of GDL research since 2001 that evaluated the effectiveness of GDL systems overall and the following GDL components.

- Learner entry ages
- Learner permit holding periods
- Supervised driving hours requirements
- Intermediate license entry ages
- Nighttime driving restrictions
- Passenger driving restrictions
- Unrestricted licensure ages


## 3. BRIEF DESCRIPTION OF META-ANALYSIS

Meta-analysis can be thought of as a mathematical synthesis of the research literature (Cooper, Hedges, \& Valentine, 2009). The goal is to combine the results from disparate studies that address the same intervention using an objective, structured protocol to obtain combined effect sizes associated with interventions that are more stable, valid, and generalizable. After scouring the literature for studies on the topic of interest, the candidate studies are scored for relevancy and quality by the research team. For those that meet the inclusion criteria, the study results are converted into a common effect size metric (e.g., rate ratios or Cohen's $d$ ). Other factors that can potentially bias or moderate the size of the effects are coded for each study, such as the age groups studied, level of control for confounding, and the individual components of the GDL programs. The inverse variance weight is calculated to combine the effect sizes (a method of applying greater weight to effect sizes that are based on higher levels of statistical precision). These combined weighted effect sizes are then stratified by the potential moderator variables in a series of analyses (1) to determine which factors are associated with heterogeneity or bias, and (2) to estimate separate weighted effect sizes for these stratification variables (e.g., the GDL components included in the programs or types of crash outcomes). In the current context, the meta-analysis was used to systematically combine the results of all GDL studies conducted since 2001 that meet the inclusion criteria to estimate combined weighted effect sizes associated with overall GDL programs and individual GDL components. One advantage to meta-analysis is that the pooling of effect estimates from a number of individual studies can sometimes reveal significant effects that are missed in individual studies because of inadequate sample size.

## 4. METHOD

The following subsections describe the identification of the studies included in the metaanalysis, the coding of information from those studies, and the analysis approach used. The reader should note that all ratings of relevance and quality were based on the judgments of researchers who are young driver subject matter experts and familiar with statistical techniques appropriate for the analysis of crash data.

### 4.1 Literature Search

The literature search focused on studies of GDL and GDL components published from 2001-2011. Researchers searched for studies from peer-reviewed journals, conference proceedings, monographs, dissertations, and other sources that were not necessarily part of the peer-review literature in order to avoid biasing the results of the meta-analysis towards positive effects. Researchers identified candidate studies by searching for keywords related to GDL and GDL components using MEDLINE, PsycINFO, PsycARTICLES, PsycEXTRA, TRIS Online,

NTIS Bibliographic Database, NHTSA’s Behavioral Safety Research Reports Library, Psychology and Behavioral Sciences Collection, Web of Science, Dissertation Abstracts, Google Scholar, and sources of relevant conference proceedings. The search terms used were wildcard variations as follows.

- Graduated driver licensing; GDL; graduated licensing system; GLS; provisional licensing program; PLP
- Novice driver; young driver; provisional driver; teenage driving; teen drivers
- Licensing age; learner age; license age; driving age; age of licensure
- Learner permit; instruction permit; driving permit; learner stage; permit period
- Supervised driving; driving practice; driving hours; practice hours
- Nighttime restriction; nighttime curfew; driving curfew
- Passenger restriction; teen passengers
- Provisional license; intermediate license; provisional stage; intermediate stage
- Contingent advancement; post-license control; point system

The reference lists of identified manuscripts were further used to identify pertinent literature. Researchers also utilized their network of contacts to identify unpublished research papers or papers under review to the extent such works could be made available. Studies of international GDL programs were excluded.

### 4.2 Relevance Screening

Researchers identified 157 GDL-related research documents in the literature and screened them using the relevance screening tool (RST) created for this project (see Appendix A for the RST form). Appendix B includes a listing of all 157 identified studies and the researchers' final rating of relevance. The purpose of the RST was to help researchers identify the studies that were relevant based on the following criteria: (1) empirical evaluation of overall GDL programs or one or more GDL program components, (2) included police-reported crashes as at least one outcome, (3) examined at least one age from 15 to 20 years old, (4) used United States data, and (5) completed or first published after January 1, 2001.

An informal pilot of the RST was conducted with two senior researchers and differences in interpretation were clarified prior to implementation and independent rating of the studies. The results of their ratings were compared by a third researcher and consensus was reached through three-way discussions. The raters were able to agree on relevance ratings for all 157 of the documents. Of the 157 GDL-related research documents identified:

- Researchers deemed 82 documents not relevant for inclusion for the following reasons.
o 1 did not include a requisite age group separately
o 15 were not crash studies
o 51 were not empirical studies of a GDL program or component
o 5 were not studies of GDL in the United States
o 8 were completed/published before 2001
o 1 study had an updated version
o 1 document did not exist
- Researchers deemed 75 documents relevant for inclusion, but further review revealed that some of the documents actually presented the same or a very similar research study. Therefore, researchers determined that:
o 49 unique studies were represented in the sample; and
o 26 of the documents were other versions of these studies (e.g., earlier monographs of peer-reviewed studies), and any new information from these documents was combined with the 49 unique primary study documents.


### 4.3 Quality Screening

Given practical limitations, all empirical studies of GDL have used quasi-experimental research designs, specifically pretest-posttest nonequivalent group and time series designs, which are at increased risk for confounding and bias than are true experimental designs (Campbell \& Stanley, 1963). An important issue concerning meta-analyses of quasi-experiments is the likelihood of residual confounding of the effect estimates (Colliver, Kucera, \& Verhulst, 2008). Failure to control for potential sources of confounding and sources of bias resulted in a prior meta-analysis of the GDL literature being severely limited (Russell, Vandermeer, \& Hartling, 2011). The earliest U.S. evaluations of GDL (i.e., Foss, Feaganes, \& Rodgman, 2001; Shope, Molnar, Elliott, \& Waller, 2001) used the crash rates of adults in the same States as a counterfactual for what would have been expected to occur to teen crash rates in the absence of the GDL programs. However, these were preliminary studies of GDL and this method was deemed to be appropriate for an early look at the potential effects of GDL programs. This method, however, became common practice for most subsequent evaluations of GDL programs. Of the numerous GDL studies using this method to control for confounding, there are only a handful that actually provided any evidence that the underlying assumption was reasonable: That the crash rates of the adults and teens behaved similarly prior to the GDL intervention, suggesting that changes in adult crash rates are a reasonable expectation for what would have happened to teen crash rates in the absence of GDL.

The method of standardizing pre-post GDL differences in teen crash rates to those for a sample of older drivers as a panacea for all potential sources of historical confounding in teen crash rates has yet to be shown as valid, yet it is the primary, and in most cases, the only method used to control for confounding in studies of GDL. This method is just a disguised pretestposttest nonequivalent groups design (Campbell \& Stanley, 1963), which is subject to numerous threats to research validity. While the approach of using adults as a counterfactual would be expected to model some historical variability in teen crash rates, it assumes that changes in adult crashes are reasonable expected values for changes in teen crashes, which may not be true. The validity of the resulting effect estimates depends to a large extent on the degree to which using changes in adult driver crash rates as the counterfactual expectation actually removed the confounding effects associated with trends, seasonality, changes in other highway-related laws, and other unmeasured historical factors (e.g., fuel prices). The assumption of this method is that adults would not be affected by the GDL programs and so changes in adult crash rates represent a good counterfactual for what would have been expected to occur for each teen age group in the absence of the GDL programs. There is no evidence that it actually removes all the confounding by these factors, and there is ample evidence that it probably does not.

The adults-as-counterfactual method assumes that pre-post GDL changes in adult driver crash rates in each State/province embody all the combined effects associated with trends, seasonality, changes in other highway-related laws, and numerous other unmeasured historical factors that would have affected teen driver crash rates in those States/provinces in the absence of GDL. Furthermore, it assumes that the magnitude of the effects of these confounding factors would have been the same for adults and teens. Among other things for this method to work, the pre-GDL trends in crash rates for the adults in each State/province must be the same as those for the teen age group. This can easily be shown to not be true for all U.S. States. For example, the California GDL program was implemented in July 1998 and Figure 1 shows the California 1990-1997 (pre-GDL) annual per capita driver fatality rates for 16-, 17-, 18-, and 19-year olds, along with those for selected combinations of adult age groups that might be used as counterfactuals for the teens per the adults-as-counterfactual method.


Figure 1. Annual California driver fatal crash involvement rates for 16-, 17-, 18-, and 19-year-olds, along with those for selected combinations of adults, 1986-1997 (pre-GDL).

The trends in California driver fatal crash involvements for the teen age groups in the figure are different than those for the adult age groups during this pre-GDL time period. This is particularly evident from 1995 to 1997, which would be a typical pre-GDL period used to calculate a standardized rate ratio. Hence, the method of standardizing the changes in teen rate ratios to those observed for adults would not have removed all confounding in the California teen rate ratios due to trends, because the adult and teen trends were different. This is likely also true for other North American States/provinces where it is sometimes the case that the pre-GDL teen
and adult driver crash trends moved in opposite directions. When dichotomous (pre versus post) outcome data are analyzed rather than using continuous data (e.g., using multiple snapshots pre versus post), it is not possible to use another approach to adjust for trends in teen crash rates. For the studies with only one pre-GDL data point, is it not possible to remove the effects of preexisting trends because most of the trend estimate is based on data points when the effects of GDL were confounded with the trend estimate. Because of these reasons, it seems likely that many of the GDL effect estimates in the literature are confounded by residual trends in teen crash rates.

The adults-as-counterfactual method also assumes that any effects associated with changes in other highway-related laws (e.g., seat belt laws, speed limits, and alcohol-related driving laws) and unmeasured historical factors (e.g., fuel prices and macroeconomic forces) would be the same for teens as for adult drivers. This is a strong assumption and there is empirical evidence of age-specific differences in effect sizes suggesting that it is not correct for at least some of these confounders (e.g., admin per se laws, primary enforcement seat belt laws, maximum speed limits, fuel prices, and unemployment; Grabowski \& Morrisey, 2004). The method also would not remove the confounding effects associated with other highway-related law changes aimed specifically at teen drivers, such those of minimal legal drinking ages and zero-tolerance laws, given that these laws have been shown to have larger effects on teens than on adult drivers (e.g., McCartt, Hellinga, \& Kirley, 2010; Villaveces, Cummings, Koepsell, Rivara, Lumley, \& Moffat, 2003). To the extent that the use of adults as counterfactuals fails to control for changes in other highway-related laws and unmeasured historical factors, the GDL program effect estimates in the literature based on this method would still be confounded by these factors.

In an attempt to avoid biasing the meta-analysis by including studies with residual confounding and bias, the researchers screened the 49 unique studies meeting the RST criteria using the quality screening tool (QST) created for this project (see Appendix A for the QST form). Researchers used the QST to code the extent to which each study adequately addressed the most serious threats to construct, internal, statistical, and external validity, relying heavily on the classic work of Campbell and Stanley (1963), and the most common threats for traffic safety intervention studies generally and GDL studies in particular. Specifically, for each study the QST assessed the following.

- Construct Validity for Crash Outcomes (DV)
o Are the crash data representative of the entire State (i.e., statewide crashes)?
o Are crashes measured reliably across all time?
o Were enough years of data before and after the intervention analyzed to provide stable estimates of both the pre- and post-intervention periods?
- Construct Validity for GDL Program/Components (IV)
o Did the GDL program/component exposure variables apply to all teens licensed in the State?
- Internal Validity
o Were crash counts adjusted for changes in the underlying population of teens?
o Were the effects of teens transitioning into the program explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the study design?
o Were the effects of age-specific trends explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for trend, was evidence given that the surrogate was likely adequate to remove it?
o Were the effects of seasonality explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for seasonality, was evidence given that the surrogate was likely adequate to remove it?
0 Were the effects of other historical factors (e.g., other traffic safety laws, weather, gas prices, etc.) explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for unmeasured historical factors, was evidence given that the surrogate was likely adequate to remove these effects?
- Statistical Validity
o Was an appropriate analysis conducted using statistical testing or confidence estimates?
o Are estimates of variability/dispersion included or calculable from the data?
o Are estimates of effect size included or calculable from the data?
o Are sample sizes for weighting effect sizes included or easily obtainable?
- Other Serious Threats to Research Validity
o Are there other threats to the validity of the study that could result in bias or confounding of the effect estimates (e.g., combining data across multiple States without adjusting for baseline differences between States teen crash rates)?

As with the RST, researchers first piloted the QST to determine any differences in interpretation and resolved any differences. The two researchers then independently rated all of the remaining documents. The results of their ratings were compared by a third researcher and consensus was reached during three-way discussions. Appendix B includes a table of the 49 studies and researchers' final quality ratings. Studies were rated as being of "High" quality if they adequately addressed all the listed threats to validity. Studies were rated as being "Good" quality if they addressed all the threats, with the possible exceptions of seasonality and other historical factors (e.g., other traffic safety laws, weather, and gas prices). They were rated as being "Moderate" quality if all threats to construct and statistical validity were addressed, but there was the possibility of residual confounding due to threats to internal validity. This typically occurred because studies used the adults-as-counterfactual approach without demonstrating similarity between the pre-GDL crash rates of the teen and adult age groups. Studies that did not, at a minimum, address all threats to construct and statistical validity were rated as being of "Low" quality. The studies were rated based on the highest quality analysis they contained, when different analysis methods and data sources were used. Of the 49 unique studies that met the RST criteria:

- 34 were deemed to be of insufficient quality to be included in the meta-analysis:
o 31 were Low quality (major confounders and other threats not addressed), o 3 were found to not result in effect estimates that could be compared to the others, - 15 were of high enough quality to be included in the meta-analysis:
o 7 were Moderate quality (some major confounders and other threats addressed),
o 2 were Good quality (most major confounders and other threats addressed),
o 6 were High quality (all major confounders and other threats addressed).

Here "quality" refers to the suitability of the study for inclusion in this meta-analysis. Many of the studies that were ultimately rated as Low quality may have been good pieces of research and are useful contributions to the literature, but did not necessarily meet the criteria for inclusion in this meta-analysis.

### 4.4 Selected Studies

Although the original intention was to include in the meta-analysis only studies of High or Good quality to ensure the meta-analysis effect estimates would represent the actual impact of GDL programs or components, those rated as Moderate quality were included due to the small number of Good and High quality studies. The studies coded as Moderate all used the adults-ascounterfactual approach, but the adequacy of this method for actually removing confounding due to trends and other historical factors is not well supported as discussed at length above. Comparisons of effect sizes as a function of study quality ratings were investigated as part of the moderator variable analyses to determine whether the inclusion of Moderate quality studies negatively impacted the overall meta-analysis results. One set of studies (Friedlander et al., 2002, 2004) was coded as being of Moderate quality, but was eventually excluded because additional information necessary to code effect sizes could not be obtained from the authors. Failure to consider or control for preexisting trends in teen crash rates was by far the single most frequent reason that studies were rated as being Low quality and excluded from the meta-analysis. Table 1 lists the 14 studies included in the meta-analysis, their quality ratings, the States studied, and the types of effect sizes that could be coded from each. Table 1 also shows how the documents were grouped to create 14 primary studies in cases where there were multiple documents covering the same or very similar material. Data from the 14 selected studies represented 13 different States, (California, Connecticut, Florida, Georgia, Iowa, Massachusetts, Maryland, Michigan, Minnesota, North Carolina, New Jersey, South Carolina, and Virginia) and three represented GDL programs across most or all U.S. States. The studies contained effect estimates of the following.

- 14 provided effect estimates of overall GDL programs
- 10 provided effect estimates of specific GDL components:
o 1 estimated effects of Learner Entry Ages
o 3 estimated effects of Learner Permit Holding Periods
o 2 estimated effects of Supervised Driving Hours Requirements
o 1 estimated effects of Intermediate License Entry Ages
o 7 estimated effects of Nighttime Driving Restrictions
o 7 estimated effects of Passenger Driving Restrictions
o 1 estimated effects of Unrestricted Licensure Ages.

Table 1. The 14 Primary Studies Included in the GDL Meta-Analysis, Quality Ratings, States Studied, and Types of Effect Sizes Coded

| \# | Ref\# | Quality rating | Authors | Year | Study title | States studied | Types of effect sizes coded |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 072 | High | Chaudhary, Williams, \& Nissen | 2007 | Evaluation and compliance of passenger restrictions in a graduated drivers licensing program. Report No. DOT HS 810 781. NHTSA. | CA, <br> MA, <br> VA | Overall GDL <br> Passenger Driving Restriction |
| 2 | 058-1 | Moderate | Chen, Baker, \& Li | 2006 | Graduated driver licensing programs and fatal crashes of 16-year-old drivers: A national evaluation. Pediatrics, 118, 56-62. | U.S. | Overall GDL <br> Learner Permit Holding Period Nighttime Driving Restriction |
|  | 058-2 |  | Baker, Chen, \& Li | 2007 | Nationwide review of graduated driver licensing. Washington, DC: AAAFT. |  |  |
| 3 | 050-1 | High | Dee, Grabowski, \& Morrisey | 2005 | Graduated driver licensing and teen traffic fatalities. Journal of Health Economics, 24, 571-589. | U.S. | Overall GDL |
|  | 050-2 |  | Morrisey, Grabowski, Dee, Campbell | 2006 | The strength of graduated driver license programs and fatalities among teen drivers and passengers. Accident Analysis and Prevention, 38, 135-141. |  |  |
|  | 050-3 |  | Morrisey \& Grabowski | 2011 | Gas prices, beer taxes and GDL programmes: effects on auto fatalities among young adults in the US. Applied Economics, 43(25), 3645-3654. |  |  |
| 4 | 074 | Good | Foss, Masten, \& Goodwin | 2007 | Long-term effects of graduated driver licensing in North Carolina (Working Paper). Chapel Hill, NC: HSRC | NC | Overall GDL <br> Nighttime Driving Restriction <br> Passenger Driving Restriction |
| 5 | 152-1 | High | Foss, Masten, Goodwin, \& O'Brien | 2012 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program (Report No. DOT HS 811 550). NHTSA. | FL, <br> MN, <br> SC, <br> VA | Overall GDL <br> Learner Permit Holding Period <br> Supervised Driving Hours <br> Passenger Driving Restriction |
|  | 152-2 |  | UNC Highway Safety Research Center | 2009 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program. UNC Highway Safety Research Center |  |  |
|  | 152-3 |  | O’Brien, Foss, Goodwin, \& Masten | 2013 | Supervised hours requirements in graduated driver licensing: Effectiveness and parental awareness. Accident Analysis \& Prevention, 50, 330-335. |  |  |
| 6 | 091 | Moderate | Kirley, Feller, Braver, \& Langenberg | 2008 | Does the Maryland graduated driver licensing law affect both 16 -year-old drivers and those who share the road with them? Journal of Safety Research, 39, 295-301. | MD | Overall GDL |


| \# | Ref\# | Quality rating | Authors | Year | Study title | States studied | Types of effect sizes coded |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 7 | 077-1 | Good | Males | 2007 | California's graduated driver license law: Effect on teenage drivers' deaths through 2005. Journal of Safety Research, 38, 651-659. | CA | Overall GDL <br> Passenger Driving Restriction |
|  | 077-2 |  | Males | 2006 | California's graduated driver license law: Effects on older teenagers. Californian Journal of Health Promotion, 4, 207-221. |  |  |
| 8 | 045-1 | High | Masten \& Hagge | 2004 | Evaluation of California's graduated driver licensing program. Journal of Safety Research, 35, 523-535. | CA | Overall GDL <br> Nighttime Driving Restriction Passenger Driving Restriction |
|  | 045-2 |  | Masten \& Hagge | 2003 | Evaluation of California's Graduated Driver Licensing Program (Report No. 205). California Department of Motor Vehicles. |  |  |
| 9 | 134-1 | High | Masten, Foss, \& Marshall | 2011 | Graduated Driver Licensing and Fatal Crashes Involving 16- to 19-Year-Old Drivers. Journal of the American Medical Association, 306(10), 1098-1103. | U.S. | Overall GDL <br> Learner Entry Age Learner Permit Holding Period Supervised Driving Hours Intermediate License Entry Age Nighttime Driving Restriction Passenger Driving Restriction Unrestricted Licensure Age |
|  | 134-2 |  | Masten | 2011 | National study of teen driver licensing systems and graduated driver licensing program core components (Dissertation). University of North Carolina. |  |  |
| 10 | 147 | High | Neyens, Donmez, \& Boyle | 2008 | The Iowa graduated driver licensing program: effectiveness in reducing crashes in teenage drivers. Journal of Safety Research, 39(4), 383-390. | IA | Overall GDL |
| 11 | 066-1 | Moderate | Rios, Wald, Nelson, Dark, Price, \& Kellerman | 2006 | Impact of Georgia's teenage and adult driver responsibility act. Annals of Emergency Medicine, 47, 361-369. | GA | Overall GDL <br> Nighttime Driving Restriction |
|  | 066-2 |  | Kellermann, Rios, Wald, Nelson, Dark, \& Price | 2007 | Graduated driver licensing in Georgia: The impact of the teenage and adult driver responsibility act (TADRA) (Report No. DOT HS-810-715). NHTSA. |  |  |


| \# | Ref\# | Quality rating | Authors | Year | Study title | States studied | Types of effect sizes coded |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 12 | 157 | Moderate | Rogers, Bentley, Campbell, Borrup, Saleheen, Wang, \& Lapidus | 2011 | Impact of Connecticut's Graduated Driver Licensing System on Teenage Motor Vehicle Crash Rates, Journal of Trauma, 71(5), S527-S530. | CT | Overall GDL |
| 13 | 048-1 | Moderate | Shope \& Molnar | 2004 | Michigan's graduated driver licensing program: Evaluation of the first four years, Journal of Safety Research, 35, 337-344. | MI | Overall GDL <br> Nighttime Driving Restriction |
|  | 048-2 |  | Shope, Molnar, Elliott, \& Waller | 2001 | Graduated driver licensing in Michigan: Early impact on motor vehicle crashes among 16 -year-old drivers. Journal of the American Medical Association, 286, 15931598. |  |  |
|  | 048-3 |  | Elliott \& Shope | 2003 | Use of a Bayesian change point model to estimate effects of a graduated driver's licensing program. Journal of Data Science, 1, 43-63. |  |  |
| 14 | 122-1 | Moderate | Williams, Chaudhary, Tefft, \& Tison | 2010 | Evaluation of New Jersey's graduated driver licensing program. Traffic Injury Prevention, 11, 1-7. | NJ | Overall GDL <br> Nighttime Driving Restriction Passenger Driving Restriction |
|  | 122-2 |  | Williams, Chaudhary, \& Tison | 2010 | Evaluation of New Jersey's Graduated Driver Licensing Program. AAAFTS. |  |  |

### 4.5 Effects and Moderator Variables Coded for the Selected Studies

Researchers coded separate rate ratio effect sizes from the included studies to estimate the impact of overall GDL programs on 15-, 16-, 17-, 18-, 19-, and 20-year-olds, as well as any combinations of these age groups available in the included studies. Separate effect sizes were calculated for total crash outcomes, injury crash outcomes, and fatal crash outcomes, when available. The effect sizes were also coded separately when estimates based on using different rate denominators (e.g., per capita versus per licensed driver) were available. In addition, effect sizes were coded to estimate the unique impact of variations in the key parameters (e.g., start times for nighttime restrictions) of the various GDL components of interest. From this point forward in the report, these variations in key parameters are referred to as "calibrations" of the following individual GDL components.

- Learner entry ages
- Learner permit holding periods
- Supervised driving hours requirements
- Intermediate license entry ages
- Nighttime driving restrictions
- Passenger driving restrictions
- Unrestricted licensure ages

Given the choice of multiple possible time periods upon which to base the effect size estimates, researchers coded the longest post-GDL time period possible to reflect the long-term impact of the GDL programs or program components. The effect sizes were coded using a MetaAnalysis Coding Sheet (MACS) created for the study (see Appendix A for the MACS form). In addition to the effect sizes and information necessary to code estimates of precision (variances), the following data elements were coded for each selected study and/or effect size for use as moderator variables in the meta-analysis.

- Primary Moderator Variables
o Whether the effect size represented the overall effect of a GDL program or the unique effect of a specific component
o The age groups represented by each effect size
o Crash outcome type for each effect size (total, injury, or fatal crashes)
o Denominator choice for each effect size (e.g., per capita, per licensed driver)
- Secondary Moderator Variables
o Program Components/Calibrations, Strength, and Type of Licensing System
- The GDL components and calibrations of the components represented by each effect size
- Program strength ratings for overall effect sizes (McCartt, Teoh, Fields, Braitman, and Hellinga, 2010)
- Teen licensing system types for overall effect sizes (Masten, Foss, \& Marshall, 2011)
o Publication Factors
- Researcher type (e.g., public/government versus academic)
- Publication mode (e.g., peer-review, monograph, conference proceeding) o Methodological Factors
- Type of study design (from QST)
- Comparison/control type (from QST)
- Length of the follow-up period (e.g., 1 year-post GDL implementation)

0 Study Quality and Validity

- Overall study quality ratings (from QST)
- Validity compared to a true experiment (from QST)

Note that the age-specific effect sizes are based on cross-sectional comparisons of drivers involved in crashes of a particular age rather than on longitudinally following a single group of teens licensed under GDL as they grew older. This is important because it means that effect sizes for 18,19 , and 20 year olds reflect the driving of a mixed population consisting of some persons who were actually licensed under the GDL programs as well as persons who were licensed at ages that were not subject to the GDL programs (18 or older in most States). It is therefore not possible to disentangle from the effect sizes of these older age groups the relative contributions of persons being licensed under the GDL programs versus effects due to younger teens delaying licensure until an age when the GDL programs do not apply. In addition, the effect sizes for all age groups include crash contributions of drivers who were completely unlicensed, as crashes were typically categorized in the contributing studies based on the age of the driver, regardless of whether they were licensed to operate motor vehicles. It is possible that implementing GDL programs changed the percentages of teens who delay licensure or who choose to drive unlicensed.

The 156 total, injury, and fatal crash effect sizes coded from the 14 primary studies are presented in Appendix C along with a description of the GDL program components represented by the effect sizes and coding outcomes for selected moderator variables. The effect sizes are organized by the study of origin and are presented alphabetically by author. The notes below the table identify idiosyncrasies that occurred during the coding of the effect sizes and explanations of the abbreviations used in the table.

### 4.6 Analysis

The effect sizes in the form of adjusted rate ratios, estimates of their precision, and coding outcomes for the moderator variables were entered into a Comprehensive Meta-Analysis (version 2.0) database, which is a program specifically created to conduct meta-analyses (Biostat, 2012). Rate ratios are normally distributed on the log scale, so each rate ratio was transformed by taking its natural logarithm (Borenstein, Hedges, Higgins, \& Rothstein, 2009). In cases where a study provided multiple effect sizes, each representing changes in crash rates for a single post-GDL year, the effect sizes were first combined using weighted meta-analysis techniques to create a single estimate across the years. For 10 of the 14 primary studies it was necessary to obtain additional data from the authors in order to calculate some or all of the effect sizes. For example, in order to convert the effect sizes in studies that used an autoregressive integrated moving average (ARIMA) model of time series data to rate ratios that could be combined with the other studies it was necessary to obtain estimates of the pre-GDL mean crash rates, which were frequently not available in the original manuscripts. In other cases estimates of
precision or comparison/control rate ratios for creating adjusted rate ratios needed to be obtained from the authors. The challenges encountered when coding effect sizes and other coding-related notes for each study are presented in Appendix C.

As a first step, funnel plots were used to check for bias (e.g., publication bias). The funnel plots were examined to determine how the effect sizes distributed as a function of each age group by type of GDL intervention (overall GDL versus specific GDL component) combination. Where there was consistency apparent among the estimates stratified by these primary moderator variables, it was deemed appropriate to combine the effect sizes into weighted estimates.

Meta-analyses are based on either a fixed-effects or random-effects model. In some cases, a mixed-effects model is appropriate, as was employed in this study. Fixed-effect models assume that all effect sizes from the included studies are estimates of the same population parameter effect size (i.e., that all estimates represent the same average change in crash rates). When there is systematic heterogeneity in the effect size estimates, they cannot be seen to represent the same average effect size and random effects models are more appropriate. Random-effects models assume that the effect sizes from the studies are estimates of distributions of different population parameter effect sizes (Borenstein, Hedges, \& Rothstein, 2007). Random-effects models include both within- and between-study sources of error in the model, which tend to result in lower weighted effect size estimates and wider confidence intervals. However, inferences based on random-effects models can be generalized to studies beyond just those included in the metaanalysis, which is not the case for fixed-effect models (Hedges \& Vevea, 1998). The approach used in the present study was mixed effects analysis whereby random effects models were used to combine studies within subgroups (e.g., strata formed by moderator variables) and fixed effect models were used to compare subgroups and yield overall effect estimates.

In the next step the effect sizes for each type of GDL intervention (overall GDL versus specific GDL component) were combined for each age group, but stratified by crash outcome types and denominator choices (primary moderator variables). This resulted in estimates of the overall effects of the various GDL interventions for each age group, along with separate estimates for each crash outcome type and denominator choice. The effect sizes were weighted by the inverse of their variances to give greater weight in the combined estimates to effect sizes that were based on higher levels of statistical precision. The statistical significance of the combined weighted effect sizes was assessed using one-sample z-tests testing the null hypothesis that the population effect size parameter is equal to zero (at $\alpha=.05$ ) and by using the 95 percent confidence intervals for the weighted effect sizes. The effect sizes were also expressed as percent changes ( $\Delta \%$ ) to make the results accessible to a larger audience. There is no minimum number of effect sizes required for meta-analysis, but if the number of studies is too small, the resulting effect size can be unstable, and vary depending on which studies are included (Rosenthal, 1995). In general only effect estimates based on at least two independent studies were considered to be meaningful and worthy of discussion, because those based on a single effect size do not provide information about the consistency of findings across multiple replications.

In the next step, primary moderator variable analyses were conducted using the $Q_{B}$ statistic to determine whether the combined effect sizes systematically varied across strata
formed by the crash outcome types and denominator choices (Borenstein, Hedges, Higgins, \& Rothstein, 2009; Borenstein \& Rothstein, 1999; Cooper, Hedges, \& Valentine, 2009; Hedges \& Pigott, 2004). This statistic indicates whether the effect estimates across strata formed by a moderator variable significantly differ from each other, analogous to between-groups ANOVA. The $Q_{B}$ statistic has an approximate chi square distribution with $k-1$ degrees of freedom, $k$ being equal to the number of strata formed by the moderator variable. A non-significant $p$ value for the $Q_{B}$ statistic indicates homogeneity of effect sizes across strata formed by the moderator variable, meaning that it is appropriate to combine the effect sizes across those strata (Borenstein, Hedges, Higgins, \& Rothstein, 2009; Cooper, Hedges \& Valentine, 2009; Hedges \& Olkin, 1985; Hedges \& Pigott, 2004). Heterogeneity was also assessed by using Forrest plots of the 95 percent confidence intervals of the effect sizes to identify outliers and patterns among the effect sizes. When the effect sizes did not differ reliably across crash outcome types or denominator choices (at $\alpha=.05$ ), the effect sizes were then combined across those strata to create more robust combined effect size estimates. When crash outcome types or denominator choices were found to reliably moderate the effect sizes, the effect sizes were not combined across strata.

Next, sensitivity analyses were conducted to test whether a single study exerted a particularly strong influence on the combined effect estimates and to test for publication bias. The former was assessed by omitting each study, one at a time, from the calculation of the combined effect estimates and determining whether these one-off estimates fell within the confidence interval of the original combined estimate based on all studies. Publication bias was assessed using the trim-and-fill method (Duval, 2005). Using this nonparametric procedure, the funnel plots of effect sizes were reviewed to establish that the tails of the data points were relatively symmetric, indicating the likely absence of publication bias.

Finally, secondary moderator variable analyses were conducted for the subset of the overall GDL effect estimates that provided the largest homogenous sample. These were conducted similarly to the primary moderator variable analyses in that combined effect estimates were created for each stratum formed by the secondary moderator variable, and these were then compared using the $Q_{B}$ statistic to determine if the effect sizes reliably differed across the strata formed by the moderator variable (using $\alpha=.05$ ).

## 5. RESULTS

### 5.1 Description of effect sizes

The coded effect sizes covered three different broad crash outcome categories based on the information available in each manuscript: Total crash outcomes ( $25.0 \%$ ), injury crash outcomes ( $41.0 \%$ ), and fatal crash outcomes ( $34.0 \%$; Table 2). The preferred metric for coded effect sizes were rates of driver involvements in total, injury, or fatal crashes. Driver crash involvements were coded for 100 percent of the total crash outcome effect sizes, 87.5 percent of the injury outcome effect sizes, and 73.6 percent of the fatal outcome effect sizes. If effect sizes for driver crash involvements were not available in a manuscript, then the next closest approximation was coded instead. The most common of these were effect sizes representing changes in rates of driver injuries, fatalities, or injuries/fatalities, meaning that the events only included drivers who were actually injured/killed in the crash, rather than all those involved in a crash, injured or not. The most common denominator used for the rates was the age-specific
population at risk, which best captures any overall public health effect of GDL programs (73.1\% of effect sizes). Other rate denominators used were numbers of age-specific licensed drivers and proportional incidence rates (e.g., nighttime or passenger crashes / total crashes), the latter being used only to estimate the specific impact of nighttime or passenger restriction components. Differences in effect sizes resulting from differences in the denominators used for the rates were investigated as part of the moderator variable analyses.

Table 2. Crash Metric and Denominators of Coded Effect Sizes

| Outcome category Specific crash metric (denominator) | $N$ | \% |
| :---: | :---: | :---: |
| Total crash outcomes | 39 | 25.0 |
| Driver crash involvements (per capita) | 17 | 68.0 |
| Driver crash involvements (per licensed driver) | 12 | 48.0 |
| Driver crash involvements (proportional incidence) | 10 | 40.0 |
| Injury crash outcomes | 64 | 41.0 |
| Driver fatal/injury crash involvements (per capita) | 47 | 73.4 |
| Driver fatal/injury crash involvements (proportional incidence) | 6 | 9.4 |
| Driver non-fatal injury crash involvements (per capita) | 2 | 3.1 |
| Driver single-vehicle fatal/injury crash involvements (proportional incidence) | 1 | 1.6 |
| Traffic fatalities/injuries (per capita) | 6 | 9.4 |
| Driver serious injuries (per capita) | 1 | 1.6 |
| Driver serious injuries (per licensed driver) | 1 | 1.6 |
| Fatal crash outcomes | 53 | 34.0 |
| Driver fatal crash involvements (per capita) | 32 | 60.4 |
| Driver fatal crash involvements (proportional incidence) | 7 | 13.2 |
| Driver fatalities (per capita) | 7 | 13.2 |
| Driver fatalities (per licensed driver) | 1 | 1.9 |
| Driver fatalities (proportional incidence) | 4 | 7.5 |
| Traffic fatalities (per capita) | 2 | 3.8 |
| Total | 156 | 100.0 |

The age group most represented among the effect sizes was 16 -year-olds ( $37.2 \% ; n=58$ ), followed by 17-year-olds ( $21.2 \% ; n=33$ ), mixed teen ages ( $17.9 \%$; $n=28$ ), 18-year-olds ( $12.8 \% ; n=20$ ), 19 -year-olds ( $10.3 \%$; $n=16$ ), and 20 -year-olds $(0.6 \% ; n=1)$. None of the studies examined 15 -year-olds as a standalone age group. A majority of the effect sizes represented the overall impact of GDL programs as a whole (59.0\%; $n=92$ ). The remainder represented the specific effect of a GDL program component: (1) passenger driving restrictions (22.4\%; $n=35$ ); (2) nighttime driving restrictions (11.5\%; $n=18$ ); (3) learner permit holding periods (4.5\%; $n=7$ ); (4) learner entry ages (1.3\%; $n=2$ ); and (5) supervised driving hours requirements ( $1.3 \% ; n=2$ ). No effect sizes represented the specific effects of intermediate license entry ages or unrestricted licensure entry ages. Note that the effect sizes for 18- to 20-year-olds reflect a mixed population of persons, some of whom were originally licensed under GDL and others who acquired a license at an age at which GDL requirements are not applied.

Because the effect sizes from the source studies are based on cross-sectional age cohorts rather than on longitudinally comparing cohorts of drivers licensed under GDL to those who were not, it is not possible to disentangle the relative contribution of these populations to the effect estimates for these age groups, but it is important to keep this in mind when interpreting the results.

Thirteen different individual States were represented in the effect sizes (California, Connecticut, Florida, Georgia, Iowa, Massachusetts, Maryland, Michigan, Minnesota, North Carolina, New Jersey, South Carolina, and Virginia), along with the U.S as a whole. A majority of the effects sizes came from studies that used time series designs ( $71.8 \%$; $n=112$ ); the remainder came from studies that used simple pre-post designs ( $28.2 \%$; $n=44$ ). The covariate/comparison groups used in the studies from which the effect sizes were coded were other contemporaneous age groups (59.0\%; $n=92$ ), other States ( $14.1 \%$; $n=22$ ), or both ( $16.7 \% ; n=26$ ). In terms of overall study quality, $71.8 \%$ of the effect sizes came from studies judged to be of high 51.3 percent ( $n=80$ ) or good 20.5 percent ( $n=32$ ) experimental validity, with 28.2 percent ( $n=44$ ) coming from studies judged as having only moderate validity. Twothirds ( $67.3 \% ; n=105$ ) of the effect sizes came from studies conducted by authors affiliated with academic institutions and the remainder ( $32.7 \%$; $n=51$ ) were from federal or State government research studies. Peer-reviewed journals were the source for 47.4 percent ( $n=74$ ) of the effect sizes, with the remainder ( $52.6 \%$; $n=82$ ) coming from research reports/monographs.

### 5.2 Overall GDL Effect Sizes Stratified by Crash Types and Denominator Choices

The overall effects of GDL programs $\left(R R_{\mathrm{w}}\right)$ on total, injury, and fatal crash outcomes are tabulated in the following sections for individual ages (i.e., single years of age) and then for various age groupings as a function of the crash numerator and denominator choices. Nine of the overall GDL effect sizes were excluded from the estimates in the tables to avoid biasing the weighted estimates. These excluded effect sizes were redundant with others from the same study and were the less-desirable outcome choice, such as traffic injuries/fatalities (i.e., anyone injured in the crash) instead of driver-based injuries/fatalities. Hence the estimates in the following sections are based on 83 ( $90.2 \%$ ) of the 92 overall GDL effect sizes. The shaded rows in the tables represent the net weighted effect size for the overall impact of GDL across all included effect sizes for an age group. The $Q_{B}$ statistics for these combined effects represent tests of strata heterogeneity across the various numerator and denominator choices used for the effect sizes, which is analogous to conducting between-subjects ANOVAs using these combined numerator and denominator choices as the levels of an independent variable (Borenstein, Hedges, Higgins, \& Rothstein, 2009; Cooper, Hedges \& Valentine, 2009; Hedges \& Pigott, 2004). All $Q_{B}$ statistics are based on a mixed effects analysis whereby a random effects model is used to combine studies within each subgroup and a fixed effect model is used to compare subgroups and yield the overall effect estimate (Borenstein, Hedges, Higgins, \& Rothstein, 2009). Separate combined effect sizes are also shown for total, injury, and fatal crash outcomes, where data were available. The $Q_{B}$ statistics for these combined effects test whether the estimates differed across the denominator choice strata (i.e., per capita versus per licensed driver), again analogous to conducting a between-subjects ANOVA with the denominator choice as the independent variable. The effect sizes expressed as percent changes ( $\Delta_{\%}$ ) are shown in the eighth column of the tables.

### 5.2.1 Individual Age Overall GDL Effect Sizes Stratified by Crash Types and Denominator Choices

Table 3 shows the overall GDL effect estimates for individual ages stratified by crash outcome type and denominator choice.

Table 3. Individual Age Overall GDL Effect Sizes Stratified by Crash Types and Denominator Choices

| Age group Outcome type Denominator | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | Z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 28 | 0.84 | 0.81 | 0.88 | -7.38 | . 000 | -15.7 | 7.57 | 4 | . 109 |
| Total crashes | 9 | 0.91 | 0.84 | 0.98 | -2.63 | .009* | -9.5 | 1.58 | 1 | . 209 |
| Per capita | 5 | 0.83 | 0.71 | 0.97 | -2.37 | .018* | -16.9 |  |  |  |
| Per licensed driver | 4 | 0.93 | 0.85 | 1.01 | -1.69 | . 091 | -7.1 |  |  |  |
| Injury crashes | 13 | 0.81 | 0.76 | 0.87 | -6.13 | .000* | -18.6 | 0.01 | 1 | . 933 |
| Per capita ${ }^{\text {a }}$ | 12 | 0.81 | 0.76 | 0.87 | -6.08 | .000* | -18.7 |  |  |  |
| Per licensed driver | 1 | 0.83 | 0.52 | 1.32 | -0.78 | . 433 | -17.0 |  |  |  |
| Fatal crashes | 6 | 0.79 | 0.70 | 0.88 | -4.01 | .000* | -21.3 | - | - | - |
| Per capita | 6 | 0.79 | 0.70 | 0.88 | -4.01 | .000* | -21.3 |  |  |  |
| 17-year-olds | 14 | 0.89 | 0.87 | 0.91 | -9.26 | . 000 | -10.7 | 29.98 | 3 | .000* |
| Total crashes | 5 | 0.89 | 0.87 | 0.92 | -8.84 | .000* | -10.8 | 29.71 | 1 | .000* |
| Per capita | 2 | 0.82 | 0.79 | 0.86 | -9.96 | .000* | -17.5 |  |  |  |
| Per licensed driver | 3 | 0.95 | 0.92 | 0.98 | -2.94 | .003* | -5.0 |  |  |  |
| Injury crashes | 5 | 0.91 | 0.83 | 1.00 | -2.00 | .046* | -8.7 | - | - | - |
| Per capita | 5 | 0.91 | 0.83 | 1.00 | -2.00 | .046* | -8.7 |  |  |  |
| Fatal crashes | 4 | 0.88 | 0.78 | 1.00 | -2.00 | .045* | -11.8 | - | - | - |
| Per capita | 4 | 0.88 | 0.78 | 1.00 | -2.00 | .045* | -11.8 |  |  |  |
| 18-year-olds | 12 | 1.00 | 0.97 | 1.02 | -0.44 | . 662 | -0.5 | 4.92 | 3 | . 178 |
| Total crashes | 4 | 0.99 | 0.97 | 1.02 | -0.43 | . 665 | -0.5 | 1.03 | 1 | . 310 |
| Per capita | 2 | 0.95 | 0.86 | 1.04 | -1.09 | . 275 | -5.3 |  |  |  |
| Per licensed driver | 2 | 1.00 | 0.97 | 1.02 | -0.17 | . 862 | -0.2 |  |  |  |
| Injury crashes | 4 | 0.96 | 0.90 | 1.03 | -1.19 | . 236 | -3.9 | - | - | - |
| Per capita | 4 | 0.96 | 0.90 | 1.03 | -1.19 | . 236 | -3.9 |  |  |  |
| Fatal crashes | 4 | 1.08 | 0.98 | 1.19 | 1.58 | . 115 | +8.1 | - | - | - |
| Per capita | 4 | 1.08 | 0.98 | 1.19 | 1.58 | . 115 | +8.1 |  |  |  |
| 19-year-olds | 11 | 1.00 | 0.98 | 1.01 | -0.72 | . 470 | -0.4 | 10.41 | 3 | .015* |
| Total crashes | 3 | 0.99 | 0.98 | 1.01 | -1.03 | . 302 | -0.8 | 8.61 | 1 | .003* |
| Per capita | 2 | 0.98 | 0.96 | 1.00 | -2.37 | .018* | -2.0 |  |  |  |
| Per licensed driver | 1 | 1.03 | 1.00 | 1.06 | 2.02 | .044* | +3.1 |  |  |  |
| Injury crashes | 4 | 1.00 | 0.98 | 1.02 | -0.18 | . 858 | -0.2 | - | - | - |
| Per capita | 4 | 1.00 | 0.98 | 1.02 | -0.18 | . 858 | -0.2 |  |  |  |
| Fatal crashes | 4 | 1.03 | 0.97 | 1.10 | 1.10 | . 269 | +3.5 | - | - | - |
| Per capita | 4 | 1.03 | 0.97 | 1.10 | 1.10 | . 269 | +3.5 |  |  |  |
| 20-year-olds | 1 | 0.94 | 0.80 | 1.10 | -0.78 | . 434 | -6.1 | - | - | - |
| Fatal crashes | 1 | 0.94 | 0.80 | 1.10 | -0.78 | . 434 | -6.1 | - | - | - |
| Per capita | 1 | 0.94 | 0.80 | 1.10 | -0.78 | . 434 | -6.1 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses.
${ }^{\text {a }}$ Four effect estimates were excluded from the 16 -year-old per capita injury estimate to maintain independence; one was based on non-fatal driver injuries and the other three were based on traffic injuries. ${ }^{*} p<.05$.

Effects for 16-year-olds. For 16-year-olds, GDL programs overall were associated with a 16 percent reduction $\left(R R_{\mathrm{w}}=0.84\right.$ [95\% CI $\left.\left.=0.81-0.88\right]\right)$ in crashes combined across all outcome types and denominator choices. With regard to the three major categories of crash outcomes, GDL programs overall were associated with reductions of 9 percent for total crashes ( $R R_{\mathrm{w}}=0.91$ [95\% CI $=0.84-0.98])$, 19 percent for injury crashes $\left(R R_{\mathrm{w}}=0.81\right.$ [95\% CI $\left.\left.=0.76-0.87\right]\right)$, and 21 percent for fatal crashes ( $R R_{\mathrm{w}}=0.79$ [95\% CI $\left.=0.70-0.88\right]$ ). The effect estimates for 16-yearolds did not reliably vary across crash outcomes or denominator choices ( $p>.05$ ). Nonetheless, the reduction in per capita 16 -year-old total crashes (17\%) was larger and statistically reliable ( $p$ $<.05$ ), whereas that for per licensed driver total crashes (7\%) was not ( $p>.05$ ). Effect sizes based on per capita rates would capture effects associated with delayed licensing in addition to other exposure-reducing factors of GDL programs, whereas those based on per licensed driver rates would underestimate changes in crashes that result from reduced licensure (McKnight, Peck, \& Foss, 2002).

Effects for 17-year-olds. For 17-year-olds, a smaller reduction (11\%) across all crash outcome types combined was found $\left(R R_{\mathrm{w}}=0.89\right.$ [ $\left.95 \% \mathrm{CI}=0.87-0.91\right]$ ). For the three major categories of crash outcomes GDL programs overall were associated with reductions of 11 percent for total crashes $\left(R R_{\mathrm{w}}=0.89\right.$ [95\% CI $\left.\left.=0.87-0.92\right]\right)$, 9 percent for injury crashes $\left(R R_{\mathrm{w}}=\right.$ 0.91 [ $95 \% \mathrm{CI}=0.83-1.00]$ ), and 12 percent for fatal crashes ( $R R_{\mathrm{w}}=0.88$ [ $\left.95 \% \mathrm{CI}=0.78-1.00\right]$ ). The heterogeneity analyses indicated that GDL programs overall were associated with larger reductions in 17-year-old crash rates per capita (17\%) than those based on per licensed driver rates (5\%; $p<.05$ ).

Effects for 18-year-olds. With regard to 18-year-olds, no reliable change was found associated with GDL programs overall for crashes combined across all outcome types ( $R R_{\mathrm{w}}=$ 1.00 [95\% CI $=0.97-1.02]$ ), or for total crashes ( $R R_{\mathrm{w}}=0.99$ [ $\left.95 \% \mathrm{CI}=0.97-1.02\right]$ ), injury crashes $\left(R R_{\mathrm{w}}=0.96\right.$ [95\% CI $\left.\left.=0.90-1.03\right]\right)$, or fatal crashes $\left(R R_{\mathrm{w}}=1.08\right.$ [95\% CI $\left.\left.=0.98-1.19\right]\right)$. The heterogeneity analyses did not indicate that there was reliable variation across outcome types or denominator choices ( $p>.05$ )

Effects for 19-year-olds. Reliable reductions associated with GDL overall were also not found for 19-year-olds across all crash outcome types combined ( $R R_{\mathrm{w}}=1.00[95 \% \mathrm{CI}=0.98$ $1.01]$ ), or for total crashes ( $R R_{\mathrm{w}}=0.99$ [95\% CI $\left.=0.98-1.01\right]$ ), injury crashes $\left(R R_{\mathrm{w}}=1.00\right.$ [95\% $\mathrm{CI}=0.98-1.02])$, or fatal crashes $\left(R R_{\mathrm{w}}=1.03\right.$ [95\% CI $\left.=0.97-1.10\right]$ ). Although the heterogeneity analyses suggested that GDL programs are associated with a 2 percent decrease in 19-year-old per population total crash rates, but a small increase in total per licensed driver crash rates, the number of effect sizes involved in this comparison is very small.

Effects for 20-year-olds. Only one effect size was available for 20-year-olds, which prohibits drawing any conclusions about pooled potential effects of GDL programs on this age group.

The reader is reminded that any rate increases or decreases for 18 or older drivers represent (a) the effects of GDL provisions imposed at age 15 to 17 that no longer apply to these older age groups, and (b) the crash rates of drivers who chose to avoid GDL by delaying licensure until 18 or older. Significant increases in crash rates among these older age groups
would therefore suggest that any crash reductions observed at age 15 to 17 are at least partially offset by delayed negative effects.

### 5.2.2 Age Group Overall GDL Effect Sizes Stratified by Crash Types and Denominator Choices

Table 4 shows the overall GDL effect estimates for combined age groups (e.g., 15-17) stratified by crash outcome type and denominator choice. The numbers of effect sizes available for these age groups were generally small, so the discussion here is limited to cases where at least two effect sizes were available to create combined effect estimates.

## Table 4. Age Group Overall GDL Effect Sizes Stratified by Crash Types and Denominator Choices

| Age group Outcome type Denominator | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | Z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 15- to 17-year-olds | 6 | 0.95 | 0.91 | 0.98 | -2.85 | .004* | -5.4 | 7.24 | 2 | .027* |
| Total crashes | 1 | 1.06 | 0.97 | 1.17 | 1.30 | . 195 | +6.4 | - | - | - |
| Per capita | 1 | 1.06 | 0.97 | 1.17 | 1.30 | . 195 | +6.4 |  |  |  |
| Injury crashes | 4 | 0.92 | 0.84 | 1.01 | -1.77 | . 077 | -8.1 | - | - | - |
| Per capita ${ }^{\text {a }}$ | 4 | 0.92 | 0.84 | 1.01 | -1.77 | . 077 | -8.1 |  |  |  |
| Fatal crashes | 1 | 0.93 | 0.88 | 0.97 | -3.25 | .001* | -7.5 | - | - | - |
| Per capita ${ }^{\text {b }}$ | 1 | 0.93 | 0.88 | 0.97 | -3.25 | .001* | -7.5 |  |  |  |
| 16-17-year-olds | 1 | 0.86 | 0.82 | 0.91 | -5.61 | .000* | -14.0 | - | - | - |
| Fatal crashes | 1 | 0.86 | 0.82 | 0.91 | -5.61 | .000* | -14.0 | - | - | - |
| Per capita | 1 | 0.86 | 0.82 | 0.91 | -5.61 | .000* | -14.0 |  |  |  |
| 16-19-year-olds | 3 | 1.04 | 0.97 | 1.12 | 1.23 | . 220 | +4.4 | 0.95 | 1 | . 329 |
| Fatal crashes | 3 | 1.04 | 0.97 | 1.12 | 1.23 | . 220 | +4.4 | 0.95 | 1 | . 329 |
| Per capita | 2 | 1.02 | 0.93 | 1.11 | 0.35 | . 725 | +1.6 |  |  |  |
| Per licensed driver | 1 | 1.09 | 0.98 | 1.22 | 1.53 | . 127 | +9.0 |  |  |  |
| 18-19-year-olds | 6 | 1.09 | 1.06 | 1.12 | 5.57 | .000* | +9.0 | 7.61 | 2 | .022* |
| Total crashes | 1 | 1.19 | 1.08 | 1.32 | 3.46 | .001* | +19.5 | - | - | - |
| Per capita | 1 | 1.19 | 1.08 | 1.32 | 3.46 | .001* | +19.5 |  |  |  |
| Injury crashes | 4 | 1.02 | 0.95 | 1.09 | 0.55 | . 582 | +1.8 | - | - | - |
| Per capita | 4 | 1.02 | 0.95 | 1.09 | 0.55 | . 582 | +1.8 |  |  |  |
| Fatal crashes | 1 | 1.10 | 1.06 | 1.14 | 5.13 | .000* | +10.0 | - | - | - |
| Per capita | 1 | 1.10 | 1.06 | 1.14 | 5.13 | .000* | +10.0 |  |  |  |
| 18- to 20-year-olds | 1 | 0.97 | 0.94 | 1.00 | -1.70 | . 090 | -2.9 | - | - | - |
| Fatal crashes | 1 | 0.97 | 0.94 | 1.00 | -1.70 | . 090 | -2.9 | - | - | - |
| Per capita ${ }^{\text {c }}$ | 1 | 0.97 | 0.94 | 1.00 | -1.70 | . 090 | -2.9 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses.
${ }^{\text {a }}$ Three effect estimates for 15 - to 17 -year-olds were excluded from this estimate to maintain independence; all were for traffic injuries.
${ }^{\mathrm{b}}$ One effect estimate for 15 - to 17 -year-olds was excluded from this estimate to maintain independence; it was for traffic fatalities.
${ }^{\text {c }}$ One effect estimate for 18 - to 20-year-olds was excluded from this estimate to maintain independence; it was for traffic fatalities. * $p<.05$.

Effects for 15- to 17-year-olds. For 15- to 17-year-olds, GDL programs overall were associated with a 5 percent reduction $\left(R R_{\mathrm{w}}=0.95\right.$ [95\% $\left.\left.\mathrm{CI}=0.91-0.98\right]\right)$ in crashes combined across all outcome types and denominator choices. This overall reduction was driven by a suggestive reduction in injury crashes $(p=.08)$ and a reliable reduction in fatal crashes ( $p<.05$ ), given that the heterogeneity analysis indicated that there was reliable variation among effects by crash outcome type ( $p<.05$ ). No net change associated with GDL overall was found for 16-19-year-old fatal crashes $\left(R R_{\mathrm{w}}=1.04\right.$ [95\% CI $\left.\left.=0.97-1.12\right]\right)$.

Effects for 18- and 19-year-olds. When only 18- and 19-year-olds were considered, GDL programs overall were found to be associated with a 9 percent increase combined across all crash outcome types ( $R R_{\mathrm{w}}=1.09$ [ $\left.95 \% \mathrm{CI}=1.06-1.12\right]$ ). The heterogeneity analysis indicated that there was variation in effects among the crash outcome types ( $p<.05$ ), with no reliable increase associated with GDL overall being found for injury crashes.

### 5.3 Overall GDL Effect Sizes Stratified by IIHS GDL Program Strength Ratings and Teen Licensing System Types

Although GDL programs overall were found to be associated with reductions in 16- and 17-year-old crash outcomes, determining whether the reductions varied systematically as a function of the strength and types of the GDL programs was explored using moderator variable analyses. This was done by coding each of the programs represented by the overall GDL effects using two different methods from the literature and then comparing combined estimates across rating categories. The first rating scheme used was that from the IIHS, as described in McCartt, Teoh, Fields, Braitman, and Hellinga (2010). Briefly, this method assigns points to each program based on the components present (e.g., learner permit holding period) and also the calibrations of those components (e.g., $<3$ months $=0$ points, $3-5$ months $=1$ point, and $6+$ months $=2$ points). Based on the summed points the programs are rated as being Good, Fair, Marginal, or Poor.

The other method used was the licensing system rating scheme described in Masten, Foss, and Marshall (2011) in which programs are classified as GDL programs only if 16 -year-old drivers are required to hold a learner permit for at least 3 months, followed by an intermediate period with either a nighttime restriction starting before $1 \mathrm{a} . \mathrm{m}$. or a passenger restriction allowing no more than 1 passenger younger than 18 years. GDL programs are further classified as having only one or both of these restrictions during the intermediate licensing phase. For both rating systems the GDL programs were coded as they existed during the first post-intervention time point in each study; changes to the program after that point were not taken into account. Only per capita effect sizes were used for these moderator analyses because the overall GDL effect sizes sometimes differed as a function of denominator choice and the numbers of per capita effects sizes were larger. To insure independence of the estimates given that some studies contributed more than one per capita effect size for a particular GDL program (e.g., total and injury crash effect sizes), a single average per capita effect size was used for each program from each study. The $Q_{B}$ statistics for these analyses test whether the strength ratings or program types reliably moderate the effect sizes. Only results based on at least two independent effect estimates are discussed below.

### 5.3.1 Overall GDL Effect Sizes Stratified by IIHS GDL Program Strength Ratings

Table 5 shows the per capita effect sizes stratified by the IIHS ratings GDL program strength. The IIHS ratings of GDL program strength were not found to be a reliable moderator of the effect sizes for either 16- or 17-year-olds ( $p>.05$ ). GDL programs rated as Fair were associated with a 28 percent reduction in crashes for 16 -year-olds, whereas those rated as Good were associated with only an 18 percent reduction ( $p<.05$ ). None of the individual IIHS ratings were reliably associated with changes in 17-year-old crash rates ( $p>.05$ ).

The IIHS ratings were found to be a reliable moderator of the effect sizes for both 18and 19-year-olds ( $p<.05$ ). However, the heterogeneity for 18 -year-olds was based on an effect estimate from a single study combined across different GDL programs (hence the IIHS rating of 'varies'), which differed from the others. For 19-year-olds GDL programs with Fair IIHS ratings were found to be associated with a 3 percent reduction in crashes, whereas those with Good ratings were associated with a 6 percent increase ( $p<.05$ ).

Table 5. Overall GDL Effect Sizes Stratified by IIHS GDL Program Ratings

| Age group Rating | Test of effect size |  |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $N_{i}$ |  | 95\% CI |  | z | $p$ | $\Delta_{\%}$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.83 | 0.79 | 0.87 | -7.77 | .000* | -17.4 | 2.97 | 3 | . 396 |
| Marginal | 1 | 0.88 | 0.78 | 1.01 | -1.86 | . 063 | -11.7 |  |  |  |
| Fair | 5 | 0.72 | 0.59 | 0.87 | -3.33 | .001* | -28.0 |  |  |  |
| Good | 7 | 0.82 | 0.77 | 0.88 | -5.94 | .000* | -17.7 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.91 | 0.99 | -2.33 | .020* | -4.7 | 1.93 | 3 | . 587 |
| Marginal | 1 | 0.97 | 0.89 | 1.05 | -0.84 | . 398 | -3.0 |  |  |  |
| Fair | 3 | 0.84 | 0.69 | 1.03 | -1.66 | . 098 | -16.0 |  |  |  |
| Good | 3 | 0.92 | 0.80 | 1.05 | -1.23 | . 218 | -8.0 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.78 | . 437 | +0.8 | 12.36 | 3 | .006* |
| Marginal | 1 | 0.98 | 0.88 | 1.10 | -0.33 | . 742 | -1.9 |  |  |  |
| Fair | 3 | 0.99 | 0.97 | 1.02 | -0.49 | . 626 | -0.5 |  |  |  |
| Good | 3 | 1.03 | 0.89 | 1.19 | 0.43 | . 667 | +3.2 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 11.49 | 3 | .009* |
| Marginal | 1 | 1.00 | 0.90 | 1.12 | 0.08 | . 934 | +0.5 |  |  |  |
| Fair | 3 | 0.97 | 0.95 | 0.99 | -2.72 | .007* | -2.9 |  |  |  |
| Good | 3 | 1.06 | 1.00 | 1.11 | 2.13 | .033* | +5.7 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. IIHS ratings of GDL strength were coded using the method described in McCartt, Teoh, Fields, Braitman, and Hellinga (2010). Programs were coded as they existed during the first post-GDL time period.
*p $<.05$.

### 5.3.2 Overall GDL Effect Sizes Stratified by Teen Licensing System Types

Table 6 shows the per capita effect sizes stratified by the Masten, Foss, and Marshall (2010) teen licensing system types. The teen licensing system classification was found to be a reliable moderator of effect sizes for 16-, 18-, and 19-year-olds ( $p<.05$ ), but not those for 17-year-olds ( $p>.05$ ). However, while all the teen licensing system classifications were associated with decreased crashes for 16 -year-olds ( $p<.05$ ), the results counter intuitively indicate that stronger GDL programs with two license restrictions during the intermediate stage were associated with smaller reductions in 16-year-old crashes (18\%) than weaker GDL programs with only one such restriction (23\%). Neither type of GDL program was reliably associated with changes in 17-year-old crashes $(p>.05)$. The heterogeneity for 18 -year-olds was again based on a single effect estimate differing from the others. Among 19-year-olds GDL programs with one intermediate license restriction were associated with a 3 percent reduction in crashes, whereas those with two such restrictions were associated with a 6 percent increase in crashes ( $p<.05$ ).

Table 6. Overall GDL Effect Sizes Stratified by Teen Licensing System Types

| Age group Category | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $4 \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.80 | 0.77 | 0.84 | -8.75 | .000* | -19.6 | 8.12 | 3 | .044* |
| Learner stage only | 1 | 0.64 | 0.53 | 0.76 | -5.10 | .000* | -36.4 |  |  |  |
| GDL 1 restriction | 5 | 0.77 | 0.66 | 0.91 | -3.15 | .002* | -22.6 |  |  |  |
| GDL 2 restrictions | 7 | 0.82 | 0.77 | 0.88 | -5.94 | .000* | -17.7 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.94 | 0.90 | 0.98 | -2.90 | .004* | -6.2 | 3.91 | 3 | . 272 |
| Learner stage only | 1 | 0.82 | 0.70 | 0.96 | -2.46 | .014* | -18.2 |  |  |  |
| GDL 1 restriction | 3 | 0.91 | 0.81 | 1.03 | -1.54 | . 124 | -8.8 |  |  |  |
| GDL 2 restrictions | 3 | 0.92 | 0.80 | 1.05 | -1.23 | . 218 | -8.5 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.78 | . 437 | +0.8 | 12.34 | 3 | .006* |
| Learner stage only | 1 | 0.98 | 0.84 | 1.14 | -0.27 | . 789 | -2.0 |  |  |  |
| GDL 1 restriction | 3 | 0.99 | 0.97 | 1.02 | -0.51 | . 611 | -0.5 |  |  |  |
| GDL 2 restrictions | 3 | 1.03 | 0.89 | 1.19 | 0.43 | . 667 | +3.2 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 11.47 | 3 | .009* |
| Learner stage only | 1 | 0.93 | 0.80 | 1.08 | -0.95 | . 344 | -7.1 |  |  |  |
| GDL 1 restriction | 3 | 0.97 | 0.95 | 0.99 | -2.54 | .011* | -2.7 |  |  |  |
| GDL 2 restrictions | 3 | 1.06 | 1.00 | 1.11 | 2.13 | .033* | +5.7 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. Licensing system ratings were coded using the method described in Masten, Foss, and Marshall (2011). Programs were coded as they existed during the first post-GDL time period.
*p $<.05$.

### 5.4 Overall GDL Effect Sizes Stratified by GDL Component Calibrations

The specific components (e.g., learner permit holding period lengths and required supervised driving hours) and calibrations of the components (e.g., 6 months and 40 hours, respectively) of the GDL programs represented by the overall GDL effect sizes were also coded as they existed during the first post-intervention time period. This section presents the overall GDL effect sizes stratified by the calibrations of each GDL program component. Again only the per capita effect estimates were used and single average effect sizes were used to represent studies with multiple per capita effect sizes (e.g., total and fatal) for a particular GDL program. Note that while this exercise has some merit, the estimates are confounded by the fact that other GDL components and calibrations also varied besides those shown in a particular table, which could account for the observed differences in the effect sizes. With regard to the nighttime and passenger restriction components, note that the effect sizes presented here are not based on the specific types of crashes targeted by these restrictions; the latter are presented later in the section on component-specific effect sizes. Only results based on independent effect sizes from at least two different sources are discussed.

### 5.4.1 Overall GDL Effect Sizes Stratified by Learner Entry Age

Table 7 shows the GDL per capita effect sizes stratified by the learner entry ages of the GDL programs represented by the effect sizes. Learner entry age calibrations were a reliable moderator of the effect sizes for 17- and 18-year-olds ( $p<.05$ ), but not those for 16- or 19-yearolds ( $p>.05$ ). Focusing on calibrations with composite effect sizes based on at least two independent studies, 16 -year-old crashes were 20 percent lower under GDL programs with a learner entry age of 15 years, 24 percent lower under programs with an entry age of $151 / 2$ years, and 18 percent lower under programs with an entry age of 16 years ( $p<.05$ ). Crashes of 17-yearolds were 10 percent lower under programs with a learner entry age of 15 years, $(p<.05)$. An entry age of 15 years was not associated with a change in crashes for 18- or 19-year-olds ( $p>$ .05).

Table 7. Overall GDL Effect Sizes Stratified by Learner Entry Age

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.80 | 0.76 | 0.84 | -9.02 | .000* | -20.3 | 2.22 | 5 | . 818 |
| 14 years, 9 months | 1 | 0.79 | 0.70 | 0.89 | -3.81 | .000* | -21.0 |  |  |  |
| 15 years | 7 | 0.80 | 0.70 | 0.91 | -3.32 | .001* | -20.4 |  |  |  |
| 15 years, 6 months | 2 | 0.76 | 0.68 | 0.84 | -5.21 | .000* | -24.3 |  |  |  |
| 15 years, 9 months | 1 | 0.72 | 0.52 | 1.00 | -1.99 | .047* | -28.1 |  |  |  |
| 16 years | 2 | 0.82 | 0.73 | 0.91 | -3.66 | .000* | -18.5 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.91 | 0.98 | -2.97 | .003* | -5.3 | 7.87 | 3 | .049* |
| 14 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 15 years | 5 | 0.90 | 0.82 | 0.99 | -2.10 | .036* | -10.0 |  |  |  |
| 15 years, 6 months | 1 | 0.98 | 0.92 | 1.05 | -0.54 | . 588 | -1.7 |  |  |  |
| 15 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 16 years | 1 | 0.81 | 0.71 | 0.93 | -3.10 | .002* | -18.6 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.02 | 1.00 | 1.06 | 1.63 | . 103 | +2.5 | 11.91 | 3 | .008* |
| 14 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 15 years | 5 | 1.01 | 0.97 | 1.05 | 0.36 | . 719 | +0.8 |  |  |  |
| 15 years, 6 months | 1 | 0.99 | 0.92 | 1.06 | -0.33 | . 743 | -1.2 |  |  |  |
| 15 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 16 years | 1 | 0.92 | 0.82 | 1.04 | -1.38 | . 167 | -8.1 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 1.00 | 0.98 | 1.02 | -0.07 | . 948 | -0.1 | 5.89 | 3 | . 117 |
| 14 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 15 years | 5 | 0.98 | 0.95 | 1.01 | -1.43 | . 151 | -2.0 |  |  |  |
| 15 years, 6 months | 1 | 1.05 | 0.99 | 1.12 | 1.58 | . 115 | +5.2 |  |  |  |
| 15 years, 9 months | 0 |  |  |  |  |  |  |  |  |  |
| 16 years | 1 | 1.04 | 0.91 | 1.18 | 0.57 | . 568 | +3.9 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses.
Note. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. *p $<.05$.

### 5.4.2 Overall GDL Effect Sizes Stratified by Learner Permit Holding Period

Table 8 shows the GDL per capita effect sizes stratified by the learner permit holding periods of the GDL programs represented by the effect sizes. Learner permit holding period calibrations were a reliable moderator of the effect sizes for all of the age groups ( $p<.05$ ). Based on composite effect sizes representing two or more two independent estimates, 16 -year-old crashes were lower under GDL programs with learner permit holding periods of 6 months or longer ( $p<.05$ ). GDL programs with longer holding periods were associated with even larger 16 -year-old crash reductions, ranging from 12 percent for those lasting 6 months to 40 percent for those lasting 12 months. Programs with learner permit holding periods of 6 months were not associated with changes in crash rates for 17-, 18-, or 19-year olds ( $p>.05$ ). However, crashes of 17-year-olds were 23 percent lower and those of 19-year-olds were 3 percent lower under GDL programs with permit holding periods lasting 12 months ( $p<.05$ ), though those for 18-year-olds were not different ( $p>.05$ ).

Table 8. Overall GDL Effect Sizes Stratified by Learner Permit Holding Period

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | $z$ | $p$ | $\Delta_{\%}$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.83 | 0.80 | 0.86 | -10.89 | .000* | -17.1 | 39.05 | 4 | .000* |
| 4 months | 1 | 0.72 | 0.52 | 1.00 | -1.99 | .047* | -28.1 |  |  |  |
| 6 months | 8 | 0.87 | 0.84 | 0.91 | -6.46 | .000* | -12.5 |  |  |  |
| 9 months | 2 | 0.76 | 0.68 | 0.84 | -5.21 | .000* | -24.3 |  |  |  |
| 12 months | 2 | 0.60 | 0.53 | 0.67 | -8.45 | .000* | -40.1 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.92 | 0.98 | -2.96 | .003* | -5.0 | 10.42 | 3 | .015* |
| 4 months | 0 |  |  |  |  |  |  |  |  |  |
| 6 months | 4 | 0.94 | 0.87 | 1.01 | -1.60 | . 110 | -6.0 |  |  |  |
| 9 months | 1 | 0.98 | 0.92 | 1.05 | -0.54 | . 588 | -1.7 |  |  |  |
| 12 months | 2 | 0.77 | 0.67 | 0.88 | -3.73 | .000* | -23.2 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.04 | 1.12 | . 263 | +1.3 | 11.49 | 3 | .009* |
| 4 months | 0 |  |  |  |  |  |  |  |  |  |
| 6 months | 4 | 1.02 | 0.93 | 1.11 | 0.32 | . 747 | +1.5 |  |  |  |
| 9 months | 1 | 0.99 | 0.92 | 1.06 | -0.33 | . 743 | -1.2 |  |  |  |
| 12 months | 2 | 1.00 | 0.97 | 1.02 | -0.32 | . 752 | -0.4 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 9.63 | 3 | .022* |
| 4 months | 0 |  |  |  |  |  |  |  |  |  |
| 6 months | 4 | 1.00 | 0.97 | 1.04 | 0.14 | . 891 | +0.3 |  |  |  |
| 9 months | 1 | 1.05 | 0.99 | 1.12 | 1.58 | . 115 | +5.2 |  |  |  |
| 12 months | 2 | 0.97 | 0.94 | 0.99 | -2.75 | .006* | -3.3 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. ${ }^{*} p<.05$.

### 5.4.3 Overall GDL Effect Sizes Stratified by Supervised Driving Hours

Table 9 shows the GDL per capita effect sizes stratified by the supervised driving hours requirements of the GDL programs represented by the effect sizes. Supervised driving hours calibrations were a reliable moderator of the effect sizes for 18- and 19-year-olds ( $p<.05$ ), but not those for 16 - or 17 -year-olds ( $p>.05$ ). Based on calibrations with at least two independent studies contributing to the composite effect size, 16-year-old crashes were 29 percent lower under GDL programs without a minimum number of required supervised driving hours, 21 percent lower under those requiring 40 hours of supervised driving, and 15 percent lower under those requiring 50 hours ( $p<.05$ ). Crashes of 17 -year-olds were 16 percent lower and those of 19-year-olds were 3 percent lower under programs without minimum required hours of supervised driving, ( $p<.05$ ), but those for 18 -year-olds were not different ( $p>.05$ ). Programs with 40 hours of required supervised driving were not associated with changes in crash rates for $17-$, 18- or 19-year-olds ( $p>.05$ ).

Table 9. Overall GDL Effect Sizes Stratified by Supervised Driving Requirement

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.82 | 0.78 | 0.86 | -8.06 | .000* | -17.8 | 2.21 | 4 | . 697 |
| None Required | 4 | 0.71 | 0.54 | 0.94 | -2.36 | .018* | -28.7 |  |  |  |
| 12 hours | 1 | 0.82 | 0.73 | 0.91 | -3.50 | .000* | -18.4 |  |  |  |
| 40 hours | 4 | 0.79 | 0.72 | 0.87 | -4.70 | .000* | -20.6 |  |  |  |
| 50 hours | 4 | 0.85 | 0.78 | 0.93 | -3.72 | .000* | -14.8 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.96 | 0.93 | 1.00 | -2.22 | .027* | -3.8 | 3.40 | 3 | . 334 |
| None Required | 4 | 0.84 | 0.72 | 0.98 | -2.23 | .026* | -16.3 |  |  |  |
| 12 hours | 0 |  |  |  |  |  |  |  |  |  |
| 40 hours | 2 | 0.98 | 0.93 | 1.03 | -0.95 | . 343 | -2.3 |  |  |  |
| 50 hours | 1 | 0.96 | 0.73 | 1.26 | -0.29 | . 771 | -3.9 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.75 | . 454 | +0.7 | 20.99 | 3 | .000* |
| None Required | 4 | 0.99 | 0.97 | 1.01 | -0.72 | . 471 | -0.8 |  |  |  |
| 12 hours | 0 |  |  |  |  |  |  |  |  |  |
| 40 hours | 2 | 0.99 | 0.93 | 1.05 | -0.45 | . 650 | -1.4 |  |  |  |
| 50 hours | 1 | 1.24 | 1.07 | 1.44 | 2.90 | .004* | +24.5 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 10.44 | 3 | .015* |
| None Required | 4 | 0.97 | 0.95 | 0.99 | -2.59 | .010* | -2.7 |  |  |  |
| 12 hours | 0 |  |  |  |  |  |  |  |  |  |
| 40 hours | 2 | 1.04 | 0.98 | 1.10 | 1.41 | . 159 | +4.0 |  |  |  |
| 50 hours | 1 | 1.10 | 0.97 | 1.24 | 1.48 | . 138 | +9.8 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. *p $<.05$.

### 5.4.4 Overall GDL Effect Sizes Stratified by Intermediate License Entry Age

Table 10 shows the GDL per capita effect sizes stratified by the intermediate license entry ages of the GDL programs represented by the effect sizes. Intermediate license entry age calibrations were a reliable moderator of the effect sizes for 18 -year-olds ( $p<.05$ ), but not those for 16-, 17-, or 19 -year-olds ( $p>.05$ ). Few calibrations had composite effect sizes based on two or more independent studies, but among them 16-year-old crashes were 22 percent lower under GDL programs with an intermediate license entry age of 16 years and 24 percent lower under programs with an entry age of $161 / 4$ years ( $p<.05$ ). While the crash rates of 17-, 18-, and 19-year-olds were not reliably different under programs with an intermediate license entry age of 16 years ( $p>.05$ ), there was some suggestion that 17 -year-old crashes may be 13 percent lower under such programs ( $p=.08$ ).

Table 10. Overall GDL Effect Sizes Stratified by Intermediate License Entry Age

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.81 | 0.77 | 0.85 | -8.36 | .000* | -19.2 | 4.36 | 6 | . 629 |
| 15 years, 6 months | 1 | 0.88 | 0.78 | 1.01 | -1.86 | . 063 | -11.7 |  |  |  |
| 16 years | 7 | 0.78 | 0.68 | 0.90 | -3.58 | .000* | -21.7 |  |  |  |
| 16 years, 1 month | 1 | 0.72 | 0.52 | 1.00 | -1.99 | .047* | -28.1 |  |  |  |
| 16 years, 3 months | 2 | 0.76 | 0.68 | 0.84 | -5.21 | .000* | -24.3 |  |  |  |
| 16 years, 6 months | 1 | 0.82 | 0.73 | 0.91 | -3.50 | .000* | -18.4 |  |  |  |
| 17 years | 1 | 0.80 | 0.54 | 1.20 | -1.06 | . 288 | -19.6 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.92 | 0.99 | -2.80 | .005* | -4.7 | 8.11 | 4 | . 088 |
| 15 years, 6 months | 1 | 0.97 | 0.89 | 1.05 | -0.84 | . 398 | -3.4 |  |  |  |
| 16 years | 4 | 0.86 | 0.73 | 1.02 | -1.72 | . 085 | -13.5 |  |  |  |
| 16 years, 1 month | 0 |  |  |  |  |  |  |  |  |  |
| 16 years, 3 months | 1 | 0.98 | 0.92 | 1.05 | -0.54 | . 588 | -1.7 |  |  |  |
| 16 years, 6 months | 0 |  |  |  |  |  |  |  |  |  |
| 17 years | 1 | 0.81 | 0.71 | 0.93 | -3.10 | .002* | -18.6 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.03 | 1.00 | 1.06 | 1.67 | . 096 | +2.6 | 12.07 | 4 | .017* |
| 15 years, 6 months | 1 | 0.98 | 0.88 | 1.10 | -0.33 | . 742 | -1.9 |  |  |  |
| 16 years | 4 | 1.01 | 0.97 | 1.06 | 0.54 | . 589 | +1.4 |  |  |  |
| 16 years, 1 month | 0 |  |  |  |  |  |  |  |  |  |
| 16 years, 3 months | 1 | 0.99 | 0.92 | 1.06 | -0.33 | . 743 | -1.2 |  |  |  |
| 16 years, 6 months | 0 |  |  |  |  |  |  |  |  |  |
| 17 years | 1 | 0.92 | 0.82 | 1.04 | -1.38 | . 167 | -8.1 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 1.01 | 0.98 | 1.03 | 0.45 | . 653 | +0.6 | 4.85 | 4 | . 303 |
| 15 years, 6 months | 1 | 1.00 | 0.90 | 1.12 | 0.08 | . 934 | +0.5 |  |  |  |
| 16 years | 4 | 0.98 | 0.95 | 1.02 | -1.10 | . 273 | -1.9 |  |  |  |
| 16 years, 1 month | 0 |  |  |  |  |  |  |  |  |  |
| 16 years, 3 months | 1 | 1.05 | 0.99 | 1.12 | 1.58 | . 115 | +5.2 |  |  |  |
| 16 years, 6 months | 0 |  |  |  |  |  |  |  |  |  |
| 17 years | 1 | 1.04 | 0.91 | 1.18 | 0.57 | . 568 | +3.9 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. All estimates for the 17 -year-old intermediate license age are based on New Jersey, which is the only State to have a minimum intermediate age of 17 years. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. * $p<.05$.

### 5.4.5 Overall GDL Effect Sizes Stratified by Nighttime Driving Restriction

Table 11 shows the per capita GDL effect sizes stratified by the intermediate stage nighttime driving restriction calibrations of the GDL programs represented by the effect sizes. Nighttime driving restriction calibrations were a reliable moderator of the effect sizes for all of the age groups ( $p<.05$ ). However, the only calibration with at least two independent studies contributing to the composite effect size was for GDL programs with a nighttime restriction starting at 12 a.m. Under such programs 16-year-old crashes were 19 percent lower and 19-yearold crashes were 6 percent higher ( $p<.05$ ), and the rates for 17- and 18-year-olds were not different ( $p>.05$ ).

Table 11. Overall GDL Effect Sizes Stratified by Nighttime Driving Restriction

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | $z$ | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.83 | 0.81 | 0.86 | -10.95 | .000* | -16.8 | 40.80 | 5 | .000* |
| 6 p.m. | 1 | 0.88 | 0.78 | 1.01 | -1.86 | . 063 | -11.7 |  |  |  |
| 9 p.m. | 1 | 0.57 | 0.48 | 0.67 | -6.80 | .000* | -43.2 |  |  |  |
| 11 p.m. | 1 | 0.90 | 0.86 | 0.96 | -3.60 | .000* | -9.5 |  |  |  |
| 12 a.m. | 9 | 0.81 | 0.77 | 0.86 | -7.65 | .000* | -18.6 |  |  |  |
| 1 a.m. | 1 | 0.64 | 0.53 | 0.76 | -5.10 | .000* | -36.4 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.96 | 0.93 | 0.99 | -2.96 | .003* | -4.2 | 15.68 | 5 | .008* |
| 6 p.m. | 1 | 0.97 | 0.89 | 1.05 | -0.84 | . 398 | -3.4 |  |  |  |
| 9 p.m. | 1 | 0.71 | 0.59 | 0.85 | -3.62 | .000* | -29.0 |  |  |  |
| 11 p.m. | 1 | 0.98 | 0.94 | 1.02 | -0.84 | . 401 | -1.8 |  |  |  |
| 12 a.m. | 3 | 0.92 | 0.80 | 1.05 | -1.23 | . 218 | -8.5 |  |  |  |
| 1 a.m. | 1 | 0.82 | 0.70 | 0.96 | -2.46 | .014* | -18.2 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.78 | . 437 | +0.8 | 12.41 | 5 | .030* |
| 6 p.m. | 1 | 0.98 | 0.88 | 1.10 | -0.33 | . 742 | -1.9 |  |  |  |
| 9 p.m. | 1 | 1.00 | 0.97 | 1.02 | -0.27 | . 786 | -0.4 |  |  |  |
| 11 p.m. | 1 | 0.99 | 0.96 | 1.03 | -0.39 | . 699 | -0.7 |  |  |  |
| 12 a.m. | 3 | 1.03 | 0.89 | 1.19 | 0.43 | . 667 | +3.2 |  |  |  |
| 1 a.m. | 1 | 0.98 | 0.84 | 1.14 | -0.27 | . 789 | -2.0 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 12.32 | 5 | .031* |
| 6 p.m. | 1 | 1.00 | 0.90 | 1.12 | 0.08 | . 934 | +0.5 |  |  |  |
| 9 p.m. | 1 | 0.97 | 0.94 | 0.99 | -2.63 | .008* | -3.2 |  |  |  |
| 11 p.m. | 1 | 0.99 | 0.94 | 1.03 | -0.61 | . 542 | -1.4 |  |  |  |
| 12 a.m. | 3 | 1.06 | 1.00 | 1.11 | 2.13 | .033* | +5.7 |  |  |  |
| 1 a.m. |  | 0.93 | 0.80 | 1.08 | -0.95 | . 344 | -7.1 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. *p $<.05$.

### 5.4.6 Overall GDL Effect Sizes Stratified by Passenger Driving Restriction

Table 12 shows the per capita GDL effect sizes stratified by the intermediate stage passenger driving restriction calibrations of the GDL programs represented by the effect sizes. Passenger driving restriction calibrations were a reliable moderator of the effect sizes for 16-, 18 -, and 19-year-olds ( $p<.05$ ), but not for 17 -year-olds ( $p>.05$ ). Based on the composite effect sizes representing two or more two independent estimates, 16-year-old crash rates were 26 percent lower under GDL programs without a passenger restriction, 24 percent lower under those allowing only one teen passenger for a period of 6 months or longer, and 14 percent lower under those disallowing any teen passengers for 6 months or longer ( $p<.05$ ). GDL programs without passenger restrictions were not reliably associated with changes in 17- or 18-year-old crash rates ( $p>.05$ ), but 19-year-old crashes were 3 percent lower under such programs ( $p<.05$ ). While the crash rates of 17-, 18-, and 19-year-olds were not reliably different under programs allowing only one teen passenger for 6 months or longer ( $p>.05$ ), there was some suggestion that 19-year-old crashes were 5 percent higher under such programs ( $p=.09$ ). Given that the heterogeneity of the results and confusing pattern of findings it is important to note that the "No restriction" category does not represent a valid baseline for assessing the causal impact of passenger restrictions; this is better suited for studies which isolate changes in passenger-related crashes, as is presented later in the section on component-specific effect estimates.

Table 12. Overall GDL Effect Sizes Stratified by Passenger Driving Restriction

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.81 | 0.78 | 0.85 | -9.01 | .000* | -18.6 | 13.89 | 5 | .016* |
| No restriction | 4 | 0.74 | 0.60 | 0.92 | -2.74 | .006* | -25.7 |  |  |  |
| 3 pass, $\geq 6$ months | 1 | 0.64 | 0.53 | 0.76 | -5.10 | .000* | -36.4 |  |  |  |
| 2 pass, $\geq 6$ months | 1 | 0.88 | 0.78 | 1.01 | -1.86 | . 063 | -11.7 |  |  |  |
| 1 pass, $\geq 6$ months | 3 | 0.76 | 0.69 | 0.84 | -5.31 | .000* | -24.1 |  |  |  |
| 0 pass, $\geq 6$ months | 4 | 0.86 | 0.80 | 0.92 | -4.06 | .000* | -14.2 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.91 | 0.99 | -2.64 | .008* | -5.3 | 4.48 | 5 | . 483 |
| No restriction | 2 | 0.85 | 0.62 | 1.16 | -1.03 | . 302 | -15.4 |  |  |  |
| 3 pass, $\geq 6$ months | 1 | 0.82 | 0.70 | 0.96 | -2.46 | .014* | -18.2 |  |  |  |
| 2 pass, $\geq 6$ months | 1 | 0.97 | 0.89 | 1.05 | -0.84 | . 398 | -3.4 |  |  |  |
| 1 pass, $\geq 6$ months | 2 | 0.90 | 0.75 | 1.08 | -1.09 | . 275 | -9.7 |  |  |  |
| 0 pass, $\geq 6$ months | 1 | 0.96 | 0.73 | 1.26 | -0.29 | . 771 | -3.9 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.76 | . 445 | +0.7 | 21.58 | 5 | .001* |
| No restriction | 2 | 1.00 | 0.97 | 1.02 | -0.45 | . 650 | -0.5 |  |  |  |
| 3 pass, $\geq 6$ months | 1 | 0.98 | 0.84 | 1.14 | -0.27 | . 789 | -2.0 |  |  |  |
| 2 pass, $\geq 6$ months | 1 | 0.98 | 0.88 | 1.10 | -0.33 | . 742 | -1.9 |  |  |  |
| 1 pass, $\geq 6$ months | 2 | 0.97 | 0.91 | 1.03 | -0.98 | . 329 | -3.1 |  |  |  |
| 0 pass, $\geq 6$ months | 1 | 1.24 | 1.07 | 1.44 | 2.90 | .004* | +24.5 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.21 | . 225 | -1.1 | 12.25 | 5 | .032* |
| No restriction | 2 | 0.97 | 0.95 | 0.99 | -2.61 | .009* | -2.8 |  |  |  |
| 3 pass, $\geq 6$ months | 1 | 0.93 | 0.80 | 1.08 | -0.95 | . 344 | -7.1 |  |  |  |
| 2 pass, $\geq 6$ months | 1 | 1.00 | 0.90 | 1.12 | 0.08 | . 934 | +0.5 |  |  |  |
| 1 pass, $\geq 6$ months | 2 | 1.05 | 0.99 | 1.11 | 1.67 | . 095 | +4.9 |  |  |  |
| 0 pass, $\geq 6$ months | 1 | 1.10 | 0.97 | 1.24 | 1.48 | . 138 | +9.8 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |
| Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effec sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. ${ }^{*} p<.05$. |  |  |  |  |  |  |  |  |  |  |

### 5.4.7 Overall GDL Effect Sizes Stratified by Unrestricted Licensure Age

Table 13 shows the per capita GDL effect sizes stratified by the unrestricted licensure age calibrations of the GDL programs represented by the effect sizes. Unrestricted licensure age calibrations were a reliable moderator of the effect sizes for 18- and 19-year-olds ( $p<.05$ ), but not those for 16 - or 17 -year-olds ( $p>.05$ ). Based on the composite effect sizes comprised of two or more independent studies, 16-year-old crash rates were 29 percent lower under GDL programs with an unrestricted licensure age of $161 / 2$ years, 15 percent lower under programs with an unrestricted age of 17 years, and 22 percent lower under those with an unrestricted license age of 18 years ( $p<.05$ ). The crash rates of 17 - and 18-year-olds were not different under programs with an unrestricted licensure age of $161 / 2$ years ( $p>.05$ ), but those for 19 -year-olds were 3 percent lower under such programs ( $p<.05$ ). Under programs with an unrestricted licensure age of 18 years the crashes of 17 -year-olds were 8 percent lower ( $p<.05$ ), but those for 18- and 19-year-olds were not different ( $p>.05$ ). Overall, no apparent pattern emerged for any age group that suggested a dose-response effect of unrestricted licensing ages.

Table 13. Overall GDL Effect Sizes Stratified by Unrestricted Licensure Age

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.82 | 0.78 | 0.87 | -7.00 | .000* | -17.7 | 2.55 | 4 | . 636 |
| 16 years, 6 months | 2 | 0.71 | 0.46 | 1.09 | -1.55 | . 121 | -29.0 |  |  |  |
| 17 years | 4 | 0.85 | 0.78 | 0.93 | -3.72 | .000* | -14.8 |  |  |  |
| 17 years, 7 months | 1 | 0.72 | 0.52 | 1.00 | -1.99 | .047* | -28.1 |  |  |  |
| 18 years | 6 | 0.78 | 0.70 | 0.88 | -4.24 | .000* | -21.8 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.95 | 0.91 | 0.99 | -2.51 | .012* | -5.3 | 1.34 | 3 | . 720 |
| 16 years, 6 months | 2 | 0.84 | 0.62 | 1.13 | -1.15 | . 251 | -16.2 |  |  |  |
| 17 years | 1 | 0.96 | 0.73 | 1.26 | -0.29 | . 771 | -3.9 |  |  |  |
| 17 years, 7 months | 0 |  |  |  |  |  |  |  |  |  |
| 18 years | 4 | 0.92 | 0.85 | 1.00 | -1.97 | .048* | -7.8 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.01 | 0.99 | 1.03 | 0.75 | . 454 | +0.7 | 21.12 | 3 | .000* |
| 16 years, 6 months | 2 | 1.00 | 0.97 | 1.02 | -0.34 | . 732 | -0.5 |  |  |  |
| 17 years | 1 | 1.24 | 1.07 | 1.44 | 2.90 | .004* | +24.5 |  |  |  |
| 17 years, 7 months | 0 |  |  |  |  |  |  |  |  |  |
| 18 years | 4 | 0.99 | 0.96 | 1.02 | -0.86 | . 390 | -1.3 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 0.99 | 0.97 | 1.01 | -1.32 | . 185 | -1.3 | 8.11 | 3 | .044* |
| 16 years, 6 months | 2 | 0.97 | 0.95 | 0.99 | -2.55 | .011* | -3.0 |  |  |  |
| 17 years | 1 | 1.10 | 0.97 | 1.24 | 1.48 | . 138 | +9.8 |  |  |  |
| 17 years, 7 months | 0 |  |  |  |  |  |  |  |  |  |
| 18 years | 4 | 1.01 | 0.96 | 1.05 | 0.36 | . 721 | +0.8 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies $=$ effect estimates based on multiple State GDL programs so a single rating could not be coded. ${ }^{*} p<.05$.

### 5.4.8 Overall GDL Effect Sizes Stratified by Contingent Advancement Requirements

While not always considered to be a core component of GDL programs, researchers also stratified the per capita GDL effect sizes by whether the GDL programs represented by the effect sizes required that teens maintain crash- and conviction-free driving records in order to advance through the GDL stages (Table 14). This requirement, which is sometimes called contingent advancement was a reliable moderator of the effect sizes 18-year-olds ( $p<.05$ ), but not those for 16 -, 17-, or 19 -year-olds ( $p>.05$ ). Crashes of 16 -year-olds were 20 percent lower under GDL programs that did not have a contingent advancement requirement, and 21 percent lower under those that did have such a requirement ( $p<.05$ ). Crashes for 17-year-olds were 15 percent lower under GDL programs with contingent advancement ( $p<.05$ ), and not different under those without contingent advancement ( $p>.05$ ). Whether the GDL programs had a contingent advancement requirement was not reliably associated with changes in the crash rates of 18 - or 19-year-olds ( $p>.05$ ).

Table 14. Overall GDL Effect Sizes Stratified by Contingent Advancement Requirement

| Age group Component calibration | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.81 | 0.76 | 0.86 | -7.02 | .000* | -19.4 | 0.65 | 2 | . 723 |
| Not required | 3 | 0.80 | 0.72 | 0.89 | -3.93 | .000* | -20.0 |  |  |  |
| Required | 10 | 0.78 | 0.71 | 0.87 | -4.58 | .000* | -21.5 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.96 | 0.93 | 0.99 | -2.26 | .024* | -3.9 | 3.27 | 2 | . 195 |
| Not required | 2 | 0.98 | 0.93 | 1.03 | -0.95 | . 343 | -2.3 |  |  |  |
| Required | 5 | 0.85 | 0.75 | 0.98 | -2.26 | .024* | -14.6 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.03 | 1.00 | 1.06 | 1.79 | . 074 | +2.9 | 9.45 | 2 | .009* |
| Not required | 2 | 0.99 | 0.93 | 1.05 | -0.45 | . 650 | -1.4 |  |  |  |
| Required | 5 | 1.00 | 0.96 | 1.05 | 0.12 | . 905 | +0.3 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 1.00 | 0.98 | 1.03 | 0.33 | . 740 | +0.4 | 3.93 | 2 | . 140 |
| Not required | 2 | 1.04 | 0.98 | 1.10 | 1.41 | . 159 | +4.0 |  |  |  |
| Required | 5 | 0.98 | 0.95 | 1.02 | -1.02 | . 309 | -1.7 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies = effect estimates were based on multiple State GDL programs so a single rating could not be coded. *p $<.05$.

### 5.5 Overall GDL Effect Sizes Stratified by Publication and Methodological Factors

To determine if the overall GDL effect sizes differed as a function of the coded publication and methodological factors of the studies, the per capita effect sizes were stratified separately by a number of different factors. The resulting effect sizes are presented in the following sections.

### 5.5.1 Overall GDL Effect Sizes Stratified by Publication Factors

Table 15 shows the per capita overall GDL effect sizes stratified by the mode of publication and type of researcher. Neither of these factors was a reliable source of heterogeneity among the overall GDL effect sizes ( $p>.05$ ), with the exception of researcher type for 17-yearolds ( $p<.05$ ). This heterogeneity for the 17-year-old effect sizes, however, was based on only one study.

Table 15. Overall GDL Effect Sizes Stratified by Publication Factors

| Age group Factor Level | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 95\% CI |  | Z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds |  |  |  |  |  |  |  |  |  |  |
| Publication mode | 15 | 0.80 | 0.75 | 0.85 | -6.82 | .000* | -20.1 | 0.01 | 1 | . 920 |
| Monograph | 7 | 0.80 | 0.72 | 0.88 | -4.26 | .000* | -20.5 |  |  |  |
| Peer review | 8 | 0.80 | 0.74 | 0.87 | -5.32 | .000* | -19.9 |  |  |  |
| Researcher type | 15 | 0.81 | 0.76 | 0.86 | -6.74 | .000* | -19.3 | 1.14 | 1 | . 285 |
| Academic | 10 | 0.78 | 0.71 | 0.85 | -5.49 | .000* | -22.0 |  |  |  |
| Public | 5 | 0.83 | 0.76 | 0.91 | -4.05 | .000* | -16.5 |  |  |  |
| 17-year-olds |  |  |  |  |  |  |  |  |  |  |
| Publication mode | 8 | 0.93 | 0.87 | 0.98 | -2.50 | .012* | -7.4 | 0.91 | 1 | . 341 |
| Monograph | 4 | 0.94 | 0.88 | 1.02 | -1.55 | . 122 | -5.6 |  |  |  |
| Peer review | 4 | 0.89 | 0.80 | 0.99 | -2.19 | .029* | -11.3 |  |  |  |
| Researcher type | 8 | 0.92 | 0.88 | 0.97 | -3.26 | .001* | -7.6 | 4.20 | 1 | .040* |
| Academic | 7 | 0.94 | 0.90 | 0.99 | -2.28 | .023* | -5.8 |  |  |  |
| Public | 1 | 0.81 | 0.71 | 0.93 | -3.10 | .002* | -18.6 |  |  |  |
| 18-year-olds |  |  |  |  |  |  |  |  |  |  |
| Publication mode | 8 | 1.00 | 0.98 | 1.02 | -0.41 | . 678 | -0.4 | 1.00 | 1 | . 317 |
| Monograph | 4 | 0.99 | 0.97 | 1.01 | -0.58 | . 562 | -0.6 |  |  |  |
| Peer review | 4 | 1.06 | 0.94 | 1.19 | 0.91 | . 360 | +5.6 |  |  |  |
| Researcher type | 8 | 1.01 | 0.97 | 1.05 | 0.58 | . 561 | +1.2 | 2.82 | 1 | . 093 |
| Academic | 7 | 1.03 | 0.98 | 1.07 | 1.12 | . 263 | +2.5 |  |  |  |
| Public | 1 | 0.92 | 0.82 | 1.04 | -1.38 | . 167 | -8.1 |  |  |  |
| 19-year-olds |  |  |  |  |  |  |  |  |  |  |
| Publication mode | 8 | 1.01 | 0.98 | 1.04 | 0.51 | . 610 | +0.7 | 1.52 | 1 | . 218 |
| Monograph | 4 | 0.99 | 0.96 | 1.03 | -0.39 | . 699 | -0.7 |  |  |  |
| Peer review | 4 | 1.03 | 0.98 | 1.08 | 1.28 | . 202 | +2.9 |  |  |  |
| Researcher type | 8 | 1.01 | 0.97 | 1.04 | 0.33 | . 740 | +0.5 | 0.26 | 1 | . 614 |
| Academic | 7 | 1.00 | 0.97 | 1.04 | 0.20 | . 845 | +0.3 |  |  |  |
| Public | 1 | 1.04 | 0.91 | 1.18 | 0.57 | . 568 | +3.9 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence.

* $p<.05$.


### 5.5.2 Overall GDL Effect Sizes Stratified by Methodological Factors

Table 16 shows the per capita overall GDL effect sizes stratified by the research design and type of comparison or control used to remove confounding. Neither of these factors was a reliable source of heterogeneity among the overall GDL effect sizes ( $p>.05$ ), with the exception of research design for 17 -year-olds ( $p<.05$ ). Specifically, effect sizes based on pre-post research designs indicate that GDL programs are reliably associated with decreased crashes for 17-yearolds ( $p<.05$ ), whereas those from time series designs suggest only a small or no reduction ( $p>$ .05).

Table 16. Overall GDL Effect Sizes Stratified by Methodological Factors

| Age group Factor Level | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds |  |  |  |  |  |  |  |  |  |  |
| Research design | 15 | 0.79 | 0.75 | 0.85 | -7.25 | .000* | -20.6 | 2.36 | 1 | . 125 |
| Pre-post | 4 | 0.73 | 0.64 | 0.83 | -4.96 | .000* | -27.0 |  |  |  |
| Time series | 11 | 0.82 | 0.76 | 0.88 | -5.51 | .000* | -18.3 |  |  |  |
| Comparison/control | 15 | 0.80 | 0.75 | 0.86 | -6.66 | .000* | -19.9 | 0.28 | 2 | . 870 |
| Other age groups | 9 | 0.80 | 0.72 | 0.89 | -4.08 | .000* | -20.0 |  |  |  |
| Other States | 3 | 0.81 | 0.74 | 0.90 | -4.09 | .000* | -18.6 |  |  |  |
| Both | 3 | 0.78 | 0.67 | 0.90 | -3.36 | .001* | -22.4 |  |  |  |
| 17-year-olds |  |  |  |  |  |  |  |  |  |  |
| Research design | 8 | 0.93 | 0.89 | 0.97 | -3.36 | .001* | -7.1 | 7.80 | 1 | .005* |
| Pre-post | 2 | 0.82 | 0.74 | 0.90 | -3.96 | .000* | -18.4 |  |  |  |
| Time series | 6 | 0.96 | 0.91 | 1.00 | -1.85 | . 065 | -4.4 |  |  |  |
| Comparison/control | 8 | 0.92 | 0.86 | 0.98 | -2.57 | .010* | -8.2 | 0.06 | 1 | . 814 |
| Other age groups | 6 | 0.92 | 0.86 | 0.99 | -2.22 | .026* | -7.9 |  |  |  |
| Other States | 0 |  |  |  |  |  |  |  |  |  |
| Both | 2 | 0.90 | 0.78 | 1.05 | -1.31 | . 189 | -9.7 |  |  |  |
| 18-year-olds |  |  |  |  |  |  |  |  |  |  |
| Research design | 8 | 1.01 | 0.97 | 1.05 | 0.53 | . 599 | +1.1 | 2.72 | 1 | . 099 |
| Pre-post | 2 | 0.94 | 0.86 | 1.03 | -1.25 | . 213 | -5.8 |  |  |  |
| Time series | 6 | 1.03 | 0.98 | 1.08 | 1.20 | . 229 | +2.9 |  |  |  |
| Comparison/control | 8 | 1.00 | 0.97 | 1.04 | 0.27 | . 788 | +0.5 | 1.08 | 1 | . 299 |
| Other age groups | 6 | 1.00 | 0.96 | 1.04 | -0.07 | . 941 | -0.1 |  |  |  |
| Other States | 0 |  |  |  |  |  |  |  |  |  |
| Both | 2 | 1.06 | 0.95 | 1.19 | 1.07 | . 285 | +6.3 |  |  |  |
| 19-year-olds |  |  |  |  |  |  |  |  |  |  |
| Research design | 8 | 1.01 | 0.97 | 1.04 | 0.33 | . 742 | +0.6 | 0.10 | 1 | . 758 |
| Pre-post | 2 | 0.99 | 0.89 | 1.10 | -0.19 | . 848 | -1.1 |  |  |  |
| Time series | 6 | 1.01 | 0.97 | 1.04 | 0.41 | . 683 | +0.7 |  |  |  |
| Comparison/control | 8 | 1.00 | 0.97 | 1.04 | 0.25 | . 804 | +0.4 | 0.00 | 1 | . 977 |
| Other age groups | 6 | 1.00 | 0.97 | 1.04 | 0.24 | . 811 | +0.5 |  |  |  |
| Other States | 0 |  |  |  |  |  |  |  |  |  |
| Both | 2 | 1.00 | 0.92 | 1.09 | 0.07 | . 942 | +0.3 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. *p $<.05$.

### 5.5.3 Overall GDL Effect Sizes Stratified by Follow-Up Time

Table 17 shows the per capita overall GDL effect sizes stratified by the length of postGDL follow-up time used in the source studies. Length of follow-up time was a reliable source of heterogeneity among the overall GDL effect sizes for 17- and 18-year-olds ( $p<.05$ ), but not those for 16 - or 19-year-olds ( $p>.05$ ). However, it was not possible to interpret the meaning of the effect moderation for 17- and 18-year-olds because too few of the follow-up intervals included more than one effect size. In contrast, the crash reductions for 16-year-olds were strikingly homogeneous across the 6-year follow-up range.

Table 17. Overall GDL Effect Sizes Stratified by Follow-Up Time (years)

| Age group Follow-up years | $N_{i}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 15 | 0.81 | 0.76 | 0.86 | -7.09 | .000* | -19.1 | 0.97 | 5 | . 965 |
| 2 | 2 | 0.81 | 0.67 | 0.97 | -2.28 | .023* | -19.3 |  |  |  |
| 3 | 3 | 0.83 | 0.71 | 0.97 | -2.29 | .022* | -17.1 |  |  |  |
| 4 | 2 | 0.79 | 0.70 | 0.89 | -3.96 | .000* | -20.9 |  |  |  |
| 5 | 3 | 0.78 | 0.66 | 0.92 | -2.93 | .003* | -22.0 |  |  |  |
| 6+ | 3 | 0.74 | 0.54 | 1.03 | -1.76 | . 079 | -25.6 |  |  |  |
| Varies | 2 | 0.83 | 0.75 | 0.92 | -3.66 | .000* | -16.9 |  |  |  |
| 17-year-olds | 8 | 0.96 | 0.94 | 0.99 | -2.91 | .004* | -3.8 | 13.15 | 5 | .022* |
| 2 | 1 | 0.97 | 0.89 | 1.05 | -0.84 | . 398 | -3.4 |  |  |  |
| 3 | 1 | 0.98 | 0.92 | 1.05 | -0.54 | . 588 | -1.7 |  |  |  |
| 4 | 1 | 0.81 | 0.71 | 0.93 | -3.10 | .002* | -18.6 |  |  |  |
| 5 | 3 | 0.81 | 0.69 | 0.94 | -2.81 | .005* | -19.4 |  |  |  |
| 6+ | 1 | 0.98 | 0.94 | 1.02 | -0.84 | . 401 | -1.8 |  |  |  |
| Varies | 1 | 0.96 | 0.91 | 1.01 | -1.54 | . 124 | -4.0 |  |  |  |
| 18-year-olds | 8 | 1.02 | 0.99 | 1.04 | 1.17 | . 240 | +1.5 | 11.72 | 5 | .039* |
| 2 | 1 | 0.98 | 0.88 | 1.10 | -0.33 | . 742 | -1.9 |  |  |  |
| 3 | 1 | 0.99 | 0.92 | 1.06 | -0.33 | . 743 | -1.2 |  |  |  |
| 4 | 3 | 1.03 | 0.91 | 1.17 | 0.48 | . 628 | +3.2 |  |  |  |
| 5 | 1 | 0.98 | 0.84 | 1.14 | -0.27 | . 789 | -2.0 |  |  |  |
| 6+ | 1 | 0.99 | 0.96 | 1.03 | -0.39 | . 699 | -0.7 |  |  |  |
| Varies | 1 | 1.11 | 1.05 | 1.17 | 3.53 | .000* | +10.6 |  |  |  |
| 19-year-olds | 8 | 1.00 | 0.98 | 1.02 | -0.11 | . 914 | -0.1 | 8.87 | 5 | . 114 |
| 2 | 1 | 1.00 | 0.90 | 1.12 | 0.08 | . 934 | +0.5 |  |  |  |
| 3 | 2 | 1.06 | 1.00 | 1.12 | 2.08 | .038* | +6.1 |  |  |  |
| 4 | 2 | 0.97 | 0.94 | 1.01 | -1.51 | . 132 | -2.7 |  |  |  |
| 5 | 1 | 0.93 | 0.80 | 1.08 | -0.95 | . 344 | -7.1 |  |  |  |
| 6+ | 1 | 0.99 | 0.94 | 1.03 | -0.61 | . 542 | -1.4 |  |  |  |
| Varies | 1 | 1.03 | 0.97 | 1.08 | 1.01 | . 312 | +2.8 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Varies $=$ effect estimates were based on multiple State GDL programs so a single rating could not be coded. * $p<.05$.

### 5.5.4 Overall GDL Effect Sizes Stratified by Overall Study Quality and Validity

The overall study quality ratings were based on the extent to which each study adequately addressed the most serious threats to construct, internal, statistical, and external validity. Although the ratings ranged from Low to High, only studies deemed to be of Moderate, Good, or High quality were used in the meta-analysis. The overall study validity rating was a subjective judgment of each study's validity and freedom of ambiguity compared to a perfect randomized design. Again, although the ratings assigned to studies considered for inclusion ranged from $<50 \%$ to $\geq 90 \%$, only studies with ratings of 50 percent or higher were incidentally included in the meta-analysis.

Table 18 shows the per capita overall GDL effect sizes stratified by the overall quality and validity of the source studies. Neither overall study quality nor study validity was found to be reliable sources of heterogeneity among the overall GDL effect sizes for 16-, 18-, or 19-yearolds ( $p>.05$ ). However, both of these factors were found to moderate the effect sizes for 17-year-olds ( $p<.05$ ). Specifically, only studies of moderate quality or 60-74 percent comparability to the validity of a randomized design were found to be reliably associated with a reduction in 17 -year-old per capita crashes ( $p<.05$ ).

Table 18. Overall GDL Effect Sizes Stratified by Study Quality and Validity

| Age group Factor Level | $N_{\text {i }}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | $z$ | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds |  |  |  |  |  |  |  |  |  |  |
| Overall quality | 15 | 0.82 | 0.78 | 0.87 | -7.17 | .000* | -17.5 | 3.37 | 2 | . 186 |
| Moderate | 5 | 0.77 | 0.68 | 0.88 | -3.96 | .000* | -23.0 |  |  |  |
| Good | 2 | 0.66 | 0.47 | 0.92 | -2.48 | .013* | -34.0 |  |  |  |
| High | 8 | 0.84 | 0.79 | 0.89 | -5.74 | .000* | -15.7 |  |  |  |
| Overall validity | 15 | 0.81 | 0.76 | 0.86 | -6.81 | .000* | -19.3 | 3.80 | 3 | . 284 |
| 50-59\% | 1 | 0.79 | 0.70 | 0.89 | -3.81 | .000* | -21.0 |  |  |  |
| 60-74\% | 4 | 0.76 | 0.62 | 0.92 | -2.75 | .006* | -24.3 |  |  |  |
| 75-89\% | 9 | 0.80 | 0.73 | 0.87 | -5.29 | .000* | -20.3 |  |  |  |
| $\geq 90 \%$ | 1 | 0.96 | 0.79 | 1.16 | -0.40 | . 687 | -3.9 |  |  |  |
| 17-year-olds |  |  |  |  |  |  |  |  |  |  |
| Overall quality | 8 | 0.96 | 0.94 | 0.99 | -2.94 | .003* | -3.8 | 12.44 | 2 | .002* |
| Moderate | 2 | 0.82 | 0.74 | 0.90 | -3.96 | .000* | -18.4 |  |  |  |
| Good | 2 | 0.81 | 0.60 | 1.09 | -1.38 | . 167 | -18.8 |  |  |  |
| High | 4 | 0.97 | 0.95 | 1.00 | -1.87 | . 062 | -2.5 |  |  |  |
| Overall validity | 8 | 0.93 | 0.89 | 0.97 | -3.36 | .001* | -7.1 | 7.80 | 1 | .005* |
| 50-59\% | 0 |  |  |  |  |  |  |  |  |  |
| 60-74\% | 2 | 0.82 | 0.74 | 0.90 | -3.96 | .000* | -18.4 |  |  |  |
| 75-89\% | 6 | 0.96 | 0.91 | 1.00 | -1.85 | . 065 | -4.4 |  |  |  |
| $\geq 90 \%$ | 0 |  |  |  |  |  |  |  |  |  |
| 18-year-olds |  |  |  |  |  |  |  |  |  |  |
| Overall quality | 8 | 1.00 | 0.95 | 1.05 | 0.06 | . 948 | +0.2 | 2.70 | 2 | . 260 |
| Moderate | 2 | 0.94 | 0.86 | 1.03 | -1.25 | . 213 | -5.8 |  |  |  |
| Good | 2 | 1.10 | 0.89 | 1.37 | 0.86 | . 390 | +10.0 |  |  |  |
| High | 4 | 1.02 | 0.96 | 1.08 | 0.64 | . 522 | +2.0 |  |  |  |
| Overall validity | 8 | 1.01 | 0.97 | 1.05 | 0.53 | . 599 | +1.1 | 2.72 | 1 | . 099 |
| 50-59\% | 0 |  |  |  |  |  |  |  |  |  |
| 60-74\% | 2 | 0.94 | 0.86 | 1.03 | -1.25 | . 213 | -5.8 |  |  |  |
| 75-89\% | 6 | 1.03 | 0.98 | 1.08 | 1.20 | . 229 | +2.9 |  |  |  |
| $\geq 90 \%$ | 0 |  |  |  |  |  |  |  |  |  |
| 19-year-olds |  |  |  |  |  |  |  |  |  |  |
| Overall quality | 8 | 1.01 | 0.98 | 1.04 | 0.87 | . 386 | +1.2 | 0.18 | 2 | . 912 |
| Moderate | 2 | 0.99 | 0.89 | 1.10 | -0.19 | . 848 | -1.1 |  |  |  |
| Good | 2 | 1.02 | 0.90 | 1.15 | 0.25 | . 803 | +1.5 |  |  |  |
| High | 4 | 1.01 | 0.98 | 1.04 | 0.91 | . 361 | +1.4 |  |  |  |
| Overall validity | 8 | 1.01 | 0.97 | 1.04 | 0.33 | . 742 | +0.6 | 0.10 | 1 | . 758 |
| 50-59\% | 0 |  |  |  |  |  |  |  |  |  |
| 60-74\% | 2 | 0.99 | 0.89 | 1.10 | -0.19 | . 848 | -1.1 |  |  |  |
| 75-89\% | 6 | 1.01 | 0.97 | 1.04 | 0.41 | . 683 | +0.7 |  |  |  |
| $\geq 90 \%$ | 0 |  |  |  |  |  |  |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence.
${ }^{*} p<.05$.

### 5.6 Component-Specific Effect Sizes Stratified by GDL Component Calibrations

Recall that 64 effect sizes coded from the source studies represented the specific effects of individual GDL components rather than the effects of GDL programs as a whole. The available numbers of effect sizes for most of the components were small. None of the effect sizes represented the specific effects of intermediate license entry ages or unrestricted licensure entry ages. Only two effect sizes each represented the specific effects of learner entry ages and supervised driving hours requirements. However, there were 7 effect sizes representing learner permit holding periods, 18 for nighttime driving restrictions, and 35 for passenger driving restrictions. The specific effect sizes coded for these latter three GDL components are summarized in the following sections. Again, only results based on two or more independent effect sizes are discussed.

### 5.6.1 GDL Learner Permit Holding Period Effect Sizes

The effect sizes representing learner permit holding period requirements are shown in Table 19 stratified by the crash outcome type and denominator choice. Those for 16 -year-olds were not found to be heterogeneous across crash outcome types ( $p>.05$ ), and all of those for 17-year-olds represented per capita injury crash outcomes. The combined estimates for learner permit holding periods did not indicate that they were associated with changes in crash outcomes for either 16- or 17-year-olds ( $p>.05$ ).

Table 19. GDL Learner Permit Holding Period Effect Sizes by Crash Types and
Denominator Choices

| Age group Outcome Denominator | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 4 | 1.00 | 0.92 | 1.09 | -0.08 | . 938 | -0.3 | 0.55 | 1 | . 459 |
| Injury crashes | 3 | 0.97 | 0.88 | 1.08 | -0.50 | . 616 | -2.7 | - | - | - |
| Per capita | 3 | 0.97 | 0.88 | 1.08 | -0.50 | . 616 | -2.7 |  |  |  |
| Fatal crashes | 1 | 1.04 | 0.90 | 1.20 | 0.55 | . 582 | +4.1 | - | - | - |
| Per capita | 1 | 1.04 | 0.90 | 1.20 | 0.55 | . 582 | +4.1 |  |  |  |
| 17-year-olds | 3 | 0.95 | 0.87 | 1.04 | -1.15 | . 249 | -5.1 | - | - | - |
| Injury crashes | 3 | 0.95 | 0.87 | 1.04 | -1.15 | . 249 | -5.1 | - | - | - |
| Per capita | 3 | 0.95 | 0.87 | 1.04 | -1.15 | . 249 | -5.1 |  |  |  |

The learner permit holding period effect sizes were next stratified as a function of the specific holding period calibration that they represented, as is shown in Table 20. The specific calibration of the learner permit holding periods was not found to be a source of heterogeneity among the effect sizes for either 16 - or 17 -year-olds ( $p>.05$ ). None of the learner permit holding period calibrations were reliably associated with a change in crashes for either of the age groups ( $p>.05$ ).

Table 20. GDL Learner Permit Holding Period Effect Sizes Stratified by Calibration

| Age group Component calibration | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 16-year-olds | 4 | 1.00 | 0.92 | 1.09 | -0.08 | . 938 | -0.3 | 0.98 | 2 | . 611 |
| 3 mo . | 1 | 0.93 | 0.80 | 1.10 | -0.83 | . 408 | -6.6 |  |  |  |
| $\geq 3 \mathrm{mo}$. | 1 | 1.04 | 0.90 | 1.20 | 0.55 | . 582 | +4.1 |  |  |  |
| 6 mo . | 2 | 1.00 | 0.87 | 1.16 | 0.06 | . 955 | +0.4 |  |  |  |
| 17-year-olds | 3 | 0.96 | 0.88 | 1.04 | -1.00 | . 315 | -4.3 | 0.15 | 1 | . 701 |
| 3 mo . | 1 | 0.97 | 0.87 | 1.07 | -0.65 | . 518 | -3.3 |  |  |  |
| $\geq 3 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 6 mo . | 2 | 0.93 | 0.79 | 1.10 | -0.86 | . 390 | -6.9 |  |  |  |

### 5.6.2 GDL Nighttime Driving Restriction Effect Sizes

The specific effect sizes representing nighttime driving restrictions on crashes during the restricted nighttime hours are shown in Table 21 stratified by the crash outcome type and denominator choice. Nighttime driving restrictions in general were found to be associated with reductions combined across nighttime crash outcome types of 9 percent for 15- to 17-year-olds $\left(R R_{\mathrm{w}}=0.91\right.$ [95\% CI $\left.=0.84-0.97\right]$ ) and 9 percent for 18 -year-olds $\left(R R_{\mathrm{w}}=0.91\right.$ [95\% CI $=0.87-$ $0.96]$ ). They were not associated with a change in the combined nighttime crash outcomes of 17-year-olds ( $p>.05$ ). The effect sizes for 16-year-olds were found to be heterogeneous across nighttime crash outcomes (total versus fatal crashes), indicating that is was not appropriate to combine effects across outcome types for this age group ( $p<.05$ ). The outcome-specific effects sizes for 16-year-olds indicate that nighttime restrictions in general were associated with a 31 percent reduction in total nighttime crashes ( $R R_{\mathrm{w}}=0.69$ [ $\left.95 \% \mathrm{CI}=0.64-0.74\right]$ ), but no change in nighttime fatal crashes ( $R R_{\mathrm{w}}=1.02$ [95\% CI = 0.91-1.15]). Crash outcome types and denominator choices were not found to be reliable sources of heterogeneity for any of the other age groups ( $p>.05$ ).

## Table 21. GDL Nighttime Driving Restriction Effect Sizes by Crash Types and Denominator Choices

| Age group Outcome Denominator | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | z | $p$ | 4\% |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 15- to 17-year-olds | 2 | 0.91 | 0.84 | 0.97 | -2.72 | .007* | -9.4 | 0.00 | 1 | . 965 |
| Total crashes | 1 | 0.91 | 0.82 | 1.00 | -1.99 | .047* | -9.3 | - | - | - |
| Proportional | 1 | 0.91 | 0.82 | 1.00 | -1.99 | .047* | -9.3 |  |  |  |
| Injury crashes | 1 | 0.90 | 0.81 | 1.01 | -1.85 | . 064 | -9.6 | - | - | - |
| Proportional | 1 | 0.90 | 0.81 | 1.01 | -1.85 | . 064 | -9.6 |  |  |  |
| 16-year-olds | 7 | 0.77 | 0.73 | 0.83 | -7.71 | .000* | -22.5 | 30.39 | 2 | .000* |
| Total crashes | 3 | 0.69 | 0.64 | 0.74 | -9.44 | .000* | -31.1 | - | - | - |
| Proportional | 3 | 0.69 | 0.64 | 0.74 | -9.44 | .000* | -31.1 |  |  |  |
| Fatal crashes | 4 | 1.02 | 0.91 | 1.15 | 0.39 | . 698 | +2.4 | 0.63 | 1 | . 428 |
| Per capita | 2 | 1.03 | 0.91 | 1.16 | 0.44 | . 659 | +2.7 |  |  |  |
| Proportional | 2 | 0.51 | 0.09 | 2.90 | -0.76 | . 445 | -49.3 |  |  |  |
| 16-17-year-olds | 1 | 0.90 | 0.82 | 1.00 | -1.99 | .047* | -9.6 | - | - | - |
| Injury crashes | 1 | 0.90 | 0.82 | 1.00 | -1.99 | .047* | -9.6 | - | - | - |
| Proportional | 1 | 0.90 | 0.82 | 1.00 | -1.99 | .047* | -9.6 |  |  |  |
| 17-year-olds | 4 | 0.88 | 0.71 | 1.09 | -1.18 | . 239 | -12.1 | 1.27 | 2 | . 531 |
| Total crashes | 2 | 0.82 | 0.61 | 1.10 | -1.33 | . 184 | -18.4 | - | - | - |
| Proportional | 2 | 0.82 | 0.61 | 1.10 | -1.33 | . 184 | -18.4 |  |  |  |
| Fatal crashes | 2 | 0.95 | 0.70 | 1.30 | -0.32 | . 748 | -4.9 | 0.79 | 1 | . 375 |
| Per capita | 1 | 1.02 | 0.72 | 1.44 | 0.11 | . 911 | +2.0 |  |  |  |
| Proportional | 1 | 0.72 | 0.36 | 1.43 | -0.94 | . 349 | -28.0 |  |  |  |
| 18-year-olds | 3 | 0.91 | 0.87 | 0.96 | -3.29 | .001* | -8.6 | 1.55 | 2 | . 461 |
| Total crashes | 1 | 0.92 | 0.87 | 0.97 | -2.99 | .003* | -8.0 | - | - | - |
| Proportional | 1 | 0.92 | 0.87 | 0.97 | -2.99 | .003* | -8.0 |  |  |  |
| Fatal crashes | 2 | 0.78 | 0.60 | 1.02 | -1.82 | . 068 | -21.7 | 0.17 | 1 | . 683 |
| Per capita | 1 | 0.81 | 0.60 | 1.10 | -1.34 | . 179 | -19.0 |  |  |  |
| Proportional | 1 | 0.72 | 0.43 | 1.18 | -1.30 | . 194 | -28.4 |  |  |  |
| 19-year-olds | 1 | 1.10 | 0.89 | 1.36 | 0.87 | . 386 | +10.0 | - | - | - |
| Fatal crashes | 1 | 1.10 | 0.89 | 1.36 | 0.87 | . 386 | +10.0 | - | - | - |
| Per capita | 1 | 1.10 | 0.89 | 1.36 | 0.87 | . 386 | +10.0 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. Effect sizes represent changes in crashes occurring during restricted hours. *p < 05 .

The nighttime driving restriction effect sizes were next stratified by the start time of the restrictions, as is shown in Table 22. Note that effect sizes for 16 -year-olds are presented separately for total and fatal nighttime crashes because outcome type was found to be a reliable source of heterogeneity for this age group. For the other age groups the various nighttime crash outcomes were combined into composite estimates across crash outcome types. Single mean effect sizes were used when a study contributed more than one effect size for a particular calibration to maintain independence.

Table 22. GDL Nighttime Driving Restriction Effect Sizes Stratified by Calibration

| Age group Component calibration | $N_{\text {i }}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | Z | $p$ | $\Delta_{\%}$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 15- to 17-year-olds | 1 | 0.91 | 0.82 | 1.00 | -1.91 | . 056 | -9.4 | - | - | - |
| 9 p.m. | 0 |  |  |  |  |  |  |  |  |  |
| 12 a.m. | 1 | 0.91 | 0.82 | 1.00 | -1.91 | . 056 | -9.4 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |
| 16-year-olds | 7 | - | - | - | - | - | - | - | - | - |
| Total crashes | 3 | 0.68 | 0.65 | 0.71 | -18.77 | .000* | -32.1 | 0.65 | 1 | . 420 |
| $9 \mathrm{p} . \mathrm{m}$. | 1 | 0.68 | 0.65 | 0.71 | -18.75 | .000* | -32.3 |  |  |  |
| 12 a.m. | 2 | 0.79 | 0.54 | 1.16 | -1.19 | . 235 | -20.7 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |
| Fatal crashes | 4 | 1.02 | 0.89 | 1.17 | 0.25 | . 803 | +1.8 | 0.60 | 2 | . 742 |
| 9 p.m. | 0 |  |  |  |  |  |  |  |  |  |
| 12 a.m. | 2 | 0.52 | 0.09 | 2.93 | -0.75 | . 455 | -48.4 |  |  |  |
| 1 a.m. | 1 | 1.04 | 0.61 | 1.78 | 0.14 | . 888 | +3.9 |  |  |  |
| Any | 1 | 1.02 | 0.88 | 1.18 | 0.28 | . 776 | +2.1 |  |  |  |
| 16-17-year-olds | 1 | 0.90 | 0.82 | 1.00 | -1.99 | .047* | -9.6 | - | - | - |
| 9 p.m. | 0 |  |  |  |  |  |  |  |  |  |
| 12 a.m. | 1 | 0.90 | 0.82 | 1.00 | -1.99 | .047* | -9.6 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |
| 17-year-olds | 3 | 0.95 | 0.91 | 0.99 | -2.18 | .029* | -5.1 | 0.15 | 1 | . 703 |
| 9 p.m. | 1 | 0.95 | 0.91 | 1.00 | -2.11 | .035* | -5.0 |  |  |  |
| 12 a.m. | 2 | 0.89 | 0.63 | 1.25 | -0.67 | . 500 | -11.2 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |
| 18-year-olds | 2 | 0.81 | 0.64 | 1.02 | -1.76 | . 078 | -18.9 | - | - | - |
| 9 p.m. | 0 |  |  |  |  |  |  |  |  |  |
| 12 a.m. | 2 | 0.81 | 0.64 | 1.02 | -1.76 | . 078 | -18.9 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |
| 19-year-olds | 1 | 1.10 | 0.89 | 1.36 | 0.87 | . 386 | +10.0 | - | - | - |
| 9 p.m. | 0 |  |  |  |  |  |  |  |  |  |
| 12 a.m. | 1 | 1.10 | 0.89 | 1.36 | 0.87 | . 386 | +10.0 |  |  |  |
| 1 a.m. | 0 |  |  |  |  |  |  |  |  |  |
| Any | 0 |  |  |  |  |  |  |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Effect sizes represent changes in crashes occurring during restricted hours. ${ }^{*} p<.05$.

The start time of the nighttime driving restrictions was not found to be a reliable moderator of the effects sizes for any age group ( $p>.05$ ). Interestingly, when only independent effect sizes were considered and the effect sizes were stratified by start time, nighttime restrictions in general were found to be associated with about a 5 percent reduction in 17-yearold nighttime crash rates ( $p<.05$ ). Among the age groups with two or more effect sizes for a particular start time very little can be interpreted because of small sample sizes. At best it can be stated that for 16-, 17-, and 18-year-olds, nighttime restrictions with $12 \mathrm{a} . \mathrm{m}$. start times were not reliably associated with reduced nighttime crashes ( $p>.05$ ).

### 5.6.3 GDL Passenger Driving Restriction Effect Sizes

The specific effect sizes representing passenger driving restrictions on crashes prohibited under the passenger restrictions are shown in Table 23 stratified by the crash outcome type and denominator choice.

Passenger driving restrictions in general were found to be associated with reductions combined across all passenger crash outcome types of 9 percent for 15- to 17-year-olds ( $R R_{\mathrm{w}}=$ 0.91 [ $95 \% \mathrm{CI}=0.86-0.98]$ ) and 6 percent for 18 -year-olds ( $R R_{\mathrm{w}}=0.94$ [95\% CI $\left.=0.92-0.96\right]$ ). The effect sizes for 16-, 17-, and 19-year-olds were found to be heterogeneous across crash outcomes, meaning that is was not appropriate to combine effects across outcome types for these age groups ( $p<.05$ ).

For 16-year-olds, the effects sizes for total crash outcomes also varied as a function of the denominator. Specifically, the estimate based on proportional incidence (restricted passenger crashes divided by total crashes) was quite different from those based on per capita and per licensed driver denominators. The total effect estimate based on proportional incidence was therefore not combined with the others. The total passenger crash outcome estimate based on combined per capita and per licensed driver effect sizes (not shown in the table) suggested that passenger restrictions in general were associated with a 6 percent reduction in 16-year-old total passenger crashes $\left(R R_{\mathrm{w}}=0.94\right.$ [ $\left.95 \% \mathrm{CI}=0.90-0.97\right]$ ). As far as the other 16 -year-old crash outcomes were concerned, passenger restrictions in general were not associated with a reliable change in injury passenger crashes ( $R R_{\mathrm{w}}=0.96$ [95\% CI $\left.=079-1.15\right]$ ), but a 16 percent reduction in fatal passenger crashes $\left(R R_{\mathrm{w}}=0.84[95 \% \mathrm{CI}=0.75-0.95]\right)$.

The total passenger crash effect sizes for 17-year-olds also differed across denominator choices, with that based on proportional incidence again being the outlying effect size ( $p<.05$ ). Excluding this effect (not shown in table), passenger restrictions were associated with a 5 percent reduction in total passenger crashes for 17-year-olds ( $R R_{\mathrm{w}}=0.95$ [ $\left.95 \% \mathrm{CI}=0.92-0.98\right]$ ). However, passenger restrictions overall were not reliably associated with a change in 17-yearold injury $\left(R R_{\mathrm{w}}=0.98\right.$ [95\% CI $\left.\left.=0.93-1.04\right]\right)$ or fatal passenger crash outcomes $\left(R R_{\mathrm{w}}=0.93\right.$ [95\% CI $=0.84-1.02]$ ).

The heterogeneity for 19-year-olds was based on the difference between total and fatal crash outcomes - the effect for total passenger outcomes suggested that passenger restrictions were associated with fewer total passenger crashes, but that for fatal passenger crashes suggested no effect-but the total effect was based on only a single effect size.

The passenger driving restriction effect sizes were next stratified by their specific calibrations (numbers of allowed passengers and length), as is shown in Table 24. Note that those for 16-, 17-, and 19-year-olds are presented separately by passenger outcome type because outcome type was found to be a reliable source of heterogeneity for these age groups. Also, the total passenger effect sizes based on proportional incidence were excluded for 16- and 17-yearolds due to heterogeneity. Mean effect sizes were used when a study contributed more than one effect size for a particular calibration to maintain independence.

The specific calibration of the passenger driving restriction was not found to be a reliable moderator of the effects sizes for any age group ( $p>.05$ ). Among the calibrations with two or more independent effect sizes, passenger restrictions allowing no more than one teen passenger that lasts for 6 months or longer were not reliably associated with a change in injury passenger crash rates ( $p>.05$ ), but were reliably associated with a 20 percent reduction in fatal passenger crashes for 16 -year-olds ( $p<.05$ ). The results also suggested, though they did not quite reach the conventional level of statistical reliability, that passenger restrictions allowing no more than one passenger for a period of 6 months or longer were associated with an 11 percent reduction in 17-year-old fatal passenger crashes ( $p=.06$ ), though there was no reliable change in their injury passenger crash rates associated with this calibration ( $p>.05$ ). Restrictions allowing one teen passenger were also associated with a 6 percent decrease in passenger crashes for 18-year-olds ( $p$ < .05). Interestingly, restrictions disallowing any passengers for 6 months or longer were not reliably associated with changes in fatal or injury passenger crash rates for 16-year-olds, passenger crash rates for 18-year-olds, or fatal passenger crash rates for 19-year-olds ( $p>.05$ ).

Table 23. GDL Passenger Driving Restriction Effect Sizes by Crash Types and Denominator Choices

| Age group Outcome Denominator | $N$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $R R_{\text {w }}$ | 95\% CI |  | Z | $p$ | $\Delta \%$ |  |  |  |
|  |  |  | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 15- to 17-year-olds | 2 | 0.91 | 0.86 | 0.98 | -2.70 | .007* | -8.6 | 1.04 | 1 | . 308 |
| Total crashes | 1 | 0.93 | 0.86 | 1.00 | -1.84 | . 066 | -6.8 | - | - | - |
| Proportional | 1 | 0.93 | 0.86 | 1.00 | -1.84 | . 066 | -6.8 |  |  |  |
| Injury crashes | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .026* | -13.9 | - | - | - |
| Proportional | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .026* | -13.9 |  |  |  |
| 16-year-olds | 13 | 0.85 | 0.82 | 0.87 | -10.95 | .000* | -15.1 | 93.73 | 6 | .000* |
| Total crashes | 3 | 0.85 | 0.82 | 0.87 | -10.62 | .000* | -15.4 | 91.72 | 2 | .000* |
| Per capita | 1 | 0.94 | 0.90 | 0.99 | -2.42 | .015* | -5.9 |  |  |  |
| Per licensed driver | 1 | 0.94 | 0.88 | 0.99 | -2.27 | .023* | -6.5 |  |  |  |
| Proportional | 1 | 0.68 | 0.64 | 0.72 | -13.91 | .000* | -32.0 |  |  |  |
| Injury crashes | 6 | 0.96 | 0.79 | 1.15 | -0.48 | . 630 | -4.5 | 0.03 | 1 | . 856 |
| Per capita | 3 | 0.96 | 0.78 | 1.19 | -0.33 | . 739 | -3.6 |  |  |  |
| Proportional | 3 | 0.93 | 0.63 | 1.36 | -0.39 | . 696 | -7.4 |  |  |  |
| Fatal crashes | 4 | 0.84 | 0.75 | 0.95 | -2.91 | . 004 | -15.6 | 0.39 | 1 | . 534 |
| Per capita | 2 | 0.84 | 0.74 | 0.94 | -2.97 | .003* | -16.5 |  |  |  |
| Proportional | 2 | 0.95 | 0.64 | 1.40 | -0.26 | . 795 | -5.0 |  |  |  |
| 16-17-year-olds | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .027* | -14.1 | - | - | - |
| Injury crashes | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .027* | -14.1 | - | - | - |
| Proportional | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .027* | -14.1 |  |  |  |
| 17-year-olds | 10 | 0.94 | 0.92 | 0.96 | -5.34 | .000* | -6.1 | 15.48 | 5 | .008* |
| Total crashes | 3 | 0.93 | 0.91 | 0.95 | -5.39 | .000* | -7.0 | 12.26 | 2 | .002* |
| Per capita | 1 | 0.96 | 0.92 | 1.01 | -1.69 | . 091 | -3.9 |  |  |  |
| Per licensed driver | 1 | 0.95 | 0.91 | 0.98 | -2.81 | .005* | -5.4 |  |  |  |
| Proportional | 1 | 0.85 | 0.80 | 0.90 | -5.53 | .000* | -15.1 |  |  |  |
| Injury crashes | 3 | 0.98 | 0.93 | 1.04 | -0.57 | . 568 | -1.6 | - | - | - |
| Per capita | 3 | 0.98 | 0.93 | 1.04 | -0.57 | . 568 | -1.6 |  |  |  |
| Fatal crashes | 4 | 0.93 | 0.84 | 1.02 | -1.49 | . 135 | -7.1 | 0.09 | 1 | . 765 |
| Per capita | 2 | 0.92 | 0.83 | 1.03 | -1.49 | . 135 | -7.7 |  |  |  |
| Proportional | 2 | 0.96 | 0.75 | 1.24 | -0.30 | . 763 | -3.8 |  |  |  |
| 18-year-olds | 5 | 0.94 | 0.92 | 0.96 | -5.98 | .000* | -6.3 | 1.06 | 2 | . 588 |
| Total crashes | 1 | 0.94 | 0.92 | 0.96 | -6.03 | .000* | -6.5 | - | - | - |
| Per capita | 1 | 0.94 | 0.92 | 0.96 | -6.03 | .000* | -6.5 |  |  |  |
| Fatal crashes | 4 | 0.98 | 0.89 | 1.08 | -0.43 | . 667 | -2.2 | 0.32 | 1 | . 573 |
| Per capita | 2 | 1.00 | 0.89 | 1.12 | -0.08 | . 937 | -0.5 |  |  |  |
| Proportional | 2 | 0.93 | 0.77 | 1.13 | -0.70 | . 481 | -6.7 |  |  |  |
| 19-year-olds | 4 | 0.94 | 0.92 | 0.96 | -5.90 | .000* | -6.1 | 8.82 | 2 | .012* |
| Total crashes | 1 | 0.93 | 0.91 | 0.95 | -6.52 | .000* | -7.1 | - | - | - |
| Per capita | 1 | 0.93 | 0.91 | 0.95 | -6.52 | .000* | -7.1 |  |  |  |
| Fatal crashes | 3 | 1.03 | 0.96 | 1.09 | 0.78 | . 438 | +2.5 | 0.49 | 1 | . 486 |
| Per capita | 2 | 1.02 | 0.95 | 1.09 | 0.57 | . 570 | +1.9 |  |  |  |
| Proportional | 1 | 1.12 | 0.87 | 1.42 | 0.87 | . 382 | +11.5 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. Effect sizes represent changes in crashes prohibited by the passenger restrictions. ${ }^{*} p<.05$.

Table 24. GDL Passenger Driving Restriction Effect Sizes Stratified by Calibration

| Age group Component calibration | $N_{\text {i }}$ | Test of effect size |  |  |  |  |  | Tests of strata heterogeneity |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 95\% CI |  |  | Z | $p$ | $\Delta \%$ |  |  |  |
|  |  | $R R_{\text {w }}$ | Lower | Upper |  |  |  | $Q_{B}$ | $d f$ | $p$ |
| 15- to 17-year-olds | 1 | 0.90 | 0.80 | 1.00 | -2.01 | .044* | -10.4 | - | - | - |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.90 | 0.80 | 1.00 | -2.01 | .044* | -10.4 |  |  |  |
| 16-year-olds | 11 | - | - | - | - | - | - | - | - | - |
| Total crashes | 1 | 0.94 | 0.89 | 0.99 | -2.33 | .020* | -6.2 | - | - | - |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.94 | 0.89 | 0.99 | $-2.33$ | .020* | -6.2 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| Injury crashes | 6 | 0.97 | 0.86 | 1.10 | -0.41 | . 680 | -2.5 | 1.54 | 2 | . 462 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.98 | 0.85 | 1.14 | -0.25 | . 804 | -1.8 |  |  |  |
| $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. | 3 | 1.05 | 0.81 | 1.36 | 0.35 | . 729 | +4.7 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.77 | 0.51 | 1.16 | -1.24 | . 216 | -22.8 |  |  |  |
| Fatal crashes | 4 | 0.83 | 0.76 | 0.91 | -4.03 | .000* | -16.7 | 1.59 | 1 | . 207 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.80 | 0.72 | 0.89 | -4.07 | .000* | -19.7 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.91 | 0.77 | 1.07 | -1.11 | . 265 | -8.9 |  |  |  |
| 16-17-year-olds | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .027* | -14.1 | - | - | - |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.86 | 0.75 | 0.98 | -2.22 | .027* | -14.1 |  |  |  |
| 17-year-olds | 8 | - | - | - | - | - | - | - | - | - |
| Total crashes | 1 | 0.95 | 0.91 | 1.00 | -2.19 | .028* | -4.6 | - | - | - |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.95 | 0.91 | 1.00 | -2.19 | .028* | -4.6 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| Injury crashes | 3 | 0.99 | 0.94 | 1.03 | -0.66 | . 507 | -1.5 | 2.08 | 1 | . 149 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 1 | 1.02 | 0.96 | 1.08 | 0.56 | . 577 | +1.8 |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.95 | 0.90 | 1.02 | -1.49 | . 137 | -4.6 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| Fatal crashes | 4 | 0.93 | 0.85 | 1.01 | -1.70 | . 089 | -7.3 | 0.94 | 1 | . 333 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.89 | 0.79 | 1.00 | -1.90 | . 057 | -11.0 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.97 | 0.85 | 1.10 | -0.46 | . 645 | -3.0 |  |  |  |
| 18-year-olds | 5 | 0.94 | 0.92 | 0.96 | -5.98 | .000* | -6.3 | 0.72 | 1 | . 396 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 3 | 0.94 | 0.92 | 0.96 | -6.04 | .000* | -6.4 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 2 | 0.99 | 0.87 | 1.12 | -0.16 | . 872 | -1.0 |  |  |  |
| 19-year-olds | 4 | - | - | - | - | - | - | - | - | - |
| Total crashes | 1 | 0.93 | 0.91 | 0.95 | -6.52 | .000* | -7.1 | - | - | - |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 1 | 0.93 | 0.91 | 0.95 | -6.52 | .000* | -7.1 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| Fatal crashes | 3 | 1.03 | 0.96 | 1.09 | 0.78 | . 438 | +2.5 | 0.01 | 1 | . 903 |
| 3 pass, $\geq 6 \mathrm{mo}$. | 0 |  |  |  |  |  |  |  |  |  |
| 1 pass, $\geq 6 \mathrm{mo}$. | 1 | 1.03 | 0.94 | 1.13 | 0.60 | . 548 | +3.0 |  |  |  |
| 0 pass, $\geq 6 \mathrm{mo}$. | 2 | 1.02 | 0.94 | 1.11 | 0.51 | . 613 | +2.2 |  |  |  |

Note. Effect estimates for 18 or older reflect both drivers licensed under GDL and others who never had GDL licenses. $N_{\mathrm{i}}=$ Number of independent effect sizes; effect sizes from the same study of the same GDL program were averaged to maintain independence. Total crash effects based on proportional incidence were excluded for 16- and 17-year-olds due to heterogeneity. Effect sizes represent changes in crashes prohibited by the passenger restrictions. ${ }^{*} p<.05$.

## 6. SUMMARY AND DISCUSSION

### 6.1 Study Findings Regarding Effects of GDL Programs Overall

GDL programs as a whole were associated with statistically reliable reductions in traffic crash outcomes of 16 percent for 16 -year-olds and 11 percent for 17 -year-olds, but were not reliably associated with changes in crash outcomes for 18 - or 19-year-olds. The crash reductions for 16- and 17-year-olds were larger for effect sizes based on per capita rates, which would capture effects associated with delayed licensing in addition to other exposure-reducing factors of GDL programs, than those based on per licensed driver rates, which underestimate changes in crashes that result from reduced licensure (McKnight, Peck, \& Foss, 2002). In fact, GDL programs were not reliably associated with reductions in per licensed driver crash rates for either of these age groups. This suggests that factors such as reductions in the percentages of teens licensed in these younger age groups, along with factors that reduce the lengths of time they are allowed to drive independently while these ages (e.g., extended learner permit holding periods), are very important factors driving the crash reductions associated with GDL programs. Although GDL programs were not found to be reliably associated with increases in crash outcomes for older teens as has been suggested by some studies (e.g., Masten, Foss, \& Marshall, 2011), the numbers of effect sizes analyzed for these age groups were small, so this controversy should not be considered resolved.

While the effect sizes for 16 -year-olds were invariant with regard to various publication and methodological factors and overall study quality/validity, this was not the case for 17-yearold effect estimates. Specifically, effect sizes based on time series designs, which are some of the strongest quasi-experimental designs (Campbell \& Stanley, 1963), suggest that GDL programs overall are associated with smaller 17-year-old crash reductions than those based on weaker prepost designs. This finding was mirrored with regard to ratings of overall study quality and validity, with only the effect sizes based on lower quality studies or those with lower ratings of validity suggesting that GDL programs were associated with reductions in crashes for 17-yearolds. It is worth noting that the quality and validity of the included studies was restricted in the meta-analysis by only including those of moderate or higher quality, so it is unknown how effect sizes for either age group would differ in studies with even less rigor.

The IIHS ratings of GDL program strength were not found to be a reliable moderator of the effect sizes for 16- or 17-year-olds. That being said, the effect sizes in the present study did not reflect the full range of potential IIHS ratings (i.e., no effect sizes for poor studies were included), which would tend to attenuate any relationship between the IIHS ratings and crash reductions. The teen licensing system classification used by Masten, Foss, and Marshall (2011) was found to be a reliable moderator of effect sizes for 16 -year-olds, but not 17-year-olds. However, stronger GDL programs with two restrictions during the intermediate licensing phase (i.e., both a nighttime and passenger restriction) were associated with smaller reductions in 16-year-old crashes than those with only one of these restrictions. Effect estimates representing the full range of the teen licensing system classifications used under this method were not present in the analysis, which would again tend to attenuate any relationship of the classification method to crash reductions. Even though the results did not provide much support for either of these
methods of classifying GDL programs, they may still have merit when the complete ranges of ratings and classifications are considered.

### 6.2 Study Findings Regarding Specific Effects of GDL Components

One of the primary goals of this meta-analysis was to provide empirical support for including particular components in GDL programs and provide guidance about what calibrations of the various components appear to be associated with the largest crash reductions. Unfortunately the numbers of effect sizes representing the unique effects of individual GDL components and calibrations were small for most of the components, particularly when stratified by the ranges of possible calibrations for those components. Large enough numbers of effect sizes representing the unique components effects were not available for all of the age-based GDL components and supervised driving hours requirements. The numbers of available componentspecific effect sizes only allowed tentative conclusions to be made with regard to learner permit holding periods and nighttime restrictions, as they were not reflective of the full range of calibrations for these components. In general, learner permit holding periods, as represented by the calibrations for which effect sizes were available, were not associated with reductions in crashes for 16- or 17-year-olds. Nighttime driving restrictions in general were associated with a reduction in at least some types of nighttime crashes for both 16- and 17-year-olds. However, besides suggesting that starting nighttime restrictions at 12 a.m. is not the most ideal time, based on the analyses, little else could be stated about specific calibrations of nighttime restrictions. The numbers of available component-specific effect sizes were the largest for passenger restrictions. Passenger restrictions in general were found to be associated with reductions in some types of passenger-related crashes for both 16- and 17-year-olds. While the evidence was not particularly clear, the results suggest that passenger restrictions allowing no more than one teen passenger that lasts for 6 months or longer may be superior to those that disallow any teen passengers.

Given the limited conclusions that could be drawn based on the analyses of componentspecific effect sizes because of the small numbers of effect sizes available and the limited range of possible calibrations that they represented, one additional set of analyses was undertaken in an attempt to provide additional guidance regarding the specific GDL components calibrations that were associated with the largest crash reductions. In these analyses a subset of the effect sizes representing the overall impact of GDL programs were stratified, one at a time, as a function of the calibrations of the components present in the GDL programs those effect sizes reflected. While these analyses have the potential to provide some useful information about component calibrations that, when present as part of the larger GDL program, are reliably associated with reductions in crashes, these analyses are inherently confounded by the fact that the calibrations of other GDL components also vary and could account for any observed differences in effect sizes. Hence, they should at best be viewed as supporting specific component calibrations only when they complement the findings from the analyses of the component-specific effect sizes, or when other factors suggest a causal relationship such as an apparent dose-response effect. To summarize these findings, there was no evidence supporting specific component calibrations for most of the GDL components. One exception was learner permit holding periods for which there was an apparent dose-response relationship supporting longer holding periods, particularly those lasting 12 months, which was reliably associated with crash reductions for both 16- and 17-year-
olds. While the passenger restriction results somewhat collaborated the findings from the component-specific analyses suggesting that restrictions allowing no more than 1 teen passenger are associated with larger reductions than those disallowing any teen passengers, the results were also supportive of not having a passenger restriction. While there was apparent variation in nighttime restriction effect sizes for both age groups as a function of the start time of the restriction, the results were unclear due to small numbers of effects. There was also no evidence supporting specific ages for any of the three age-based GDL components. With regard to numbers of supervised driving hours and requirements to keep crash- and conviction-free records to advance through GDL stages, there was no evidence that having these components was at all beneficial.

### 6.3 Limitations

A tremendous amount of effort was taken in this study to include only the most rigorous evaluations of GDL programs that removed as many potential sources of confounding from the effect estimates as possible, given the limitations of the quasi-experimental designs. Arguably, the requirements for inclusion were too high resulting in the exclusion of important studies. Given that only 14 of the 49 studies ( $28 \%$ ) deemed to be relevant were eventually included in the meta-analysis, such criticism would not be without merit. The exclusion of specific studies should not be interpreted negatively; rather that the excluded studies did not meet the specific criteria delineated for these factors, which in our opinion, were deemed to be important for producing the least biased and most valid estimates of GDL program and component effects.

Although the original intention was to include in the meta-analysis only studies rated as being of "High" or "Good" quality, those rated as being of "Moderate" quality were also reluctantly included because the number of better quality studies was low. The studies coded as "Moderate" all used the adults-as-counterfactual approach, but the adequacy of this method for actually removing confounding due to trends and other historical factors was not well supported in the manuscripts. Comparisons of effect sizes as a function of study quality ratings was investigated as part of the moderator variable analyses to determine whether the inclusion of "Moderate" quality studies impacted the overall meta-analysis results, and was indeed found to be an important factor with regard to estimates of GDL effects for 17-year-olds. However, excluding the studies coded as "Moderate" would have severely reduced the numbers of available effect sizes for analysis and negatively impacted the extent to which the findings could be generalized to other jurisdictions.

The age-specific effect sizes are based on cross-sectional comparisons of drivers involved in crashes of a particular age rather than on longitudinally following a single group of teens licensed under GDL as they grew older. This is important because it means that effect sizes for 18,19 , and 20 year olds reflect the driving of a mixed population consisting of some persons who were actually licensed under the GDL programs as well as persons who were licensed at ages that were not subject to the GDL programs (18 or older in most States). It is therefore not possible to disentangle from the effect sizes of these older age groups the relative contributions of persons being licensed under the GDL programs versus effects due to younger teens delaying licensure until an age when the GDL programs do not apply. In addition, the effect sizes for all age groups include crash contributions of drivers who were completely unlicensed, as crashes
were typically categorized in the contributing studies based on the age of the driver, regardless of whether they were licensed to operate motor vehicles. It is possible that implementing GDL programs changed the percentages of teens who delay licensure or who choose to drive unlicensed.

### 6.4 Conclusions

This meta-analysis supports a conclusion that GDL laws create a safety benefit for 16-year-old drivers and potentially have a safety benefit for 17-year-old drivers, although to a lesser extent. Given the diversity in the configuration of GDL provisions among the States, an additional interest relates to which of the various provisions contributed to the observed effect. Insufficient studies with suitable information available existed to answer this question. Further, an authoritative study of individual GDL provisions might not even be possible given the practical limitations associated with researching this topic in a real world setting. For example, the mere existence of a curfew provision does not mean it was truly operationalized unless it can be shown the affected population was aware of it, adhered to it, and the police actually enforced it at a meaningful level. Few studies quantified these important process factors.

Although the exact effectiveness of individual GDL provisions could not be determined, the meta-analysis uncovered no indication that any provision was necessarily counterproductive for the GDL target audience of 16- and 17-year olds. Thus, a reasonable strategy for any State considering passage of a GDL law might involve enumerating the full range of provisions applicable to that State, determining which could be reasonably operationalized given available resources and support from key agencies and organizations, and adopting as comprehensive an approach as possible.

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## Appendix A

## RST, QST, and MAC Forms

## Study Descriptive Information

| 1a. Authors: | 1b. Year: |
| :--- | :--- | :--- |
| 2. Title or Citation: |  |
| 3. Source: |  |
| $\square$ Journal Article $\square$ Monograph/Report $\quad \square$ Dissertation/Thesis $\quad \square$ Conference Paper/Presentation |  |
| $\square$ Other (specify): |  |


| Study Inclusion Criteria |  |
| :---: | :---: |
| 4. Is this study an original empirical investigation of overall GDL systems or one or more Key GDL Component(s)? | $\square$ Yes $\square$ No |
| Original Empirical: Analysis of data from an original source. <br> Key Components: 1. Minimum learner stage entry ages; 2. learner permit holding periods; <br> 3. Supervised driving hours requirements; 4. Intermediate stage entry ages; <br> 5. Nighttime driving restrictions; 6. Passenger driving restrictions; <br> 7. Minimum ages of unrestricted licensure <br> Exclude: Re-analyses or summaries of already completed works; studies of only zero-tolerance alcohol laws, cell phone restrictions, speed or roadway restrictions, accelerated post-license controls, driver education/training requirements, choice of vehicles. |  |
| 5. Is at least one of the outcome measures police-reported total crashes, injury crashes, and/or fatal crashes, crash involvements, or driver crash involvements? <br> Exclude: Hospitalizations, self-reported crashes, traffic violations/tickets, licensure, arrests, subsets of crashes such as only at-fault, motorcycle, passengers, or alcohol-involved crashes. | $\square$ Yes $\square$ No |
| 6. Is the population examined one or more ages between 15-20-years-old? | $\square$ Yes $\square$ No |
| Exclude: Studies of statewide crash rates aggregated across all or most age groups. |  |
| 7. Are data from one or more U.S. jurisdictions? | $\square$ Yes $\square$ No |
| Exclude: Studies that do not include data from at least one U.S. jurisdiction. |  |
| 8. Was the study completed or first published after January 1, 2001? | $\square$ Yes $\square$ No |
| Note: Studies completed or first published before 2001 may possibly still be included. |  |
| 9. Initial decision based on relevance criteria: | IncludeInclude; combineExcludePossibly Include |
| Exclude: If "no" answer to any item 4-7 above. Possibly Include: If "yes" to 4-7, but "no" to item 8. |  |

## Comments:



Note: Time series is defined as including 3 or more data points and incorporating a time component. If there are multiple designs used in the study, code the strongest design used.

| Comparison/control type: | $\square$ Internal (other age group) $\square$ External (other state) $\square$ Both $\square$ None |
| :--- | :--- |
|  | $\square$ Other: |

Note: Use "Other" control for different type of crashes by same age group.

## Outcome Metric

Check all that apply:


| 2. Are crashes measured reliably across all time? <br> Exclude: Studies with missing/incomplete data for some or all time points. | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |
| :---: | :---: |
| 3. Were enough years of data before and after the intervention analyzed to provide stable estimates of both the pre- and post-intervention periods? <br> Note: For annual data 'Yes' is at least 2 years pre and 2 years post. For quarterly data 'Yes' is at least 8 quarters pre and 8 quarters post. For monthly data 'Yes' is 24 months pre and 24 months post. Studies that used only a single year of pre- or post-intervention data (mono-operation bias) are always 'No.' Do not include "skipped" time points. | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |

## Construct Validity for GDL Program/Components (IV)

4. Did the GDL program/component exposure variable(s) apply to all teens licensed in the state?

Exclude: Studies that only include components that apply to a subset of the 16 or 17 -year-old population (e.g., only a nighttime driving restriction that could be avoided if the teens enrolled in driver education), unless it was demonstrated that the majority of 16- or 17 -year-olds were indeed subject to the intervention.

## Internal Validity

5. Were crash counts adjusted for changes in the underlying population of teens?

Exclude: Studies of raw crash counts not adjusted for population-level exposure using age-specific population, licensed drivers, or mileage either by computing rates or including exposure estimates in the statistical models.
6. Were the effects of teens transitioning into the program explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the study design?
Note: Include effects of both increased numbers of teens rushing to be licensed prior to the intervention as well as allowing adequate numbers of teens to transition into the treatment. If effects averaged over multiple years were
$\square$ Yes
$\square$ No
$\square$ Maybe
$\square$ N/A
$\square$ Yes
$\square$ No
$\square$ Maybe
$\square$ N/A analyzed such that temporary transition effects should be minimized, code 'Yes.'
7. Were the effects of age-specific trends explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for trend, was evidence given that the surrogate was likely adequate to remove it?

Note: If crash rates of another unaffected age group or another state were used to implicitly remove trend, but it
$\square$ Yes
$\square$ No
$\square$ Maybe
$\square$ N/A was not demonstrated that the trends were similar prior to the intervention, code Maybe.
8. Were the effects of seasonality explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for seasonality, was evidence given that the surrogate was likely adequate to remove it?
Note: If crash rates of another unaffected age group or another state were used to implicitly remove these effects, but it was not demonstrated that the crash rates were correlated prior to the intervention, code Maybe.
9. Were the effects of other historical factors (e.g., other traffic safety laws, weather, gas prices, etc.) explicitly modeled, reasonably shown to not be a threat to validity, or accounted for in the design? If a surrogate was used to implicitly adjust for unmeasured historical factors, was evidence given that the surrogate was likely adequate to remove these effects?

Note: If crash rates of another unaffected age group or another state were used to implicitly remove these effects, but it was not demonstrated that the crash rates were correlated prior to the intervention, code Maybe.

| 10a. Was an appropriate analysis conducted using statistical testing or confidence estimates? <br> Exclude: Studies lacking appropriate statistical analysis or without estimates of precision. | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |
| :---: | :---: |
| 10b. Are estimates of variability/dispersion included or calculable from the data? <br> Note. Estimates of variability/dispersion include confidence intervals, actual $p$ values for the effect estimates, standard deviations, or standard errors. | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |
| 10c. Are estimates of effect size included or calculable from the data? <br> Note. Examples include adjusted crash rates, rate ratios, rate differences, hazard ratios, odds ratios, or model parameters. | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |
| 10d. Are sample sizes for weighting effect sizes included or easily obtainable? <br> Note. Examples include population sizes, numbers of licensed drivers, and crash counts. |  |
| Other Serious Threats to Research Validity |  |
| 11. Is the study free from other serious threats to research validity not otherwise named above? <br> If "No" or "Maybe," describe the threat(s) to research validity: | $\square$ Yes $\square$ No $\square$ Maybe $\square$ N/A |
| Overall Quality Rating |  |
| High: Yes to all Items 1-11 <br> Good: Yes to all Items 1-7 \& 10-11 <br> Moderate: Yes to all Items 1-4 \& 10-11, but Maybe to at least one Item 5-7 <br> Low: No to any Item 1-7 or 10-11 | $\square$ High $\square$ Good $\square$ Moderate $\square$ Low $\square$ N/A |
| Overall Research Validity Relative to a True Experiment |  |

Confidence in study validity and freedom of ambiguity compared to a perfect randomized design:

| $\square<50 \%$ | $\square 50-59 \%$ | $\square 60-74 \%$ | $\square 75-89 \%$ | $\square \geq 90 \%$ |
| :--- | :--- | :--- | :--- | :--- |
|  | Comments and Explanations | $\square$ N/A |  |  |
|  |  |  |  |  |

Meta-Analysis Coding Sheet (MACS)
Source Study Descriptive Information


## Overall GDL or Component Represented

$\square$ Overall GDL, or
$\square$ Only a one or more GDL components with the calibration checked below


## Overall Outcome Metric

| Major Crash Type: | $\square$ Total crashes (Fatal+ Injury + PDO)$\square$ Injury crashes (excluding fatal)$\square$ Other (describe): |  | $\square$ Injury crashes (including fatal)$\square$ Fatal crashes only |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Crash Subtype: | Nighttime crashes during restricted hours $\quad \square$ C |  |  | Crash-involvements (driver-based events) <br> Other Road Users (e.g., pedestrians or bicyclists) |  |  |
| Numerator: | Crashes (unique events) <br> Victims (check): $\square$ <br> Drivers Passengers $\square$ Other: |  |  |  |  |  |
| Denominator: | $\square$ Population $\square$ Licensed Drivers $\square$ VMT $\square$ Other: |  |  |  |  |  |
| Notes. Code as close as possible single effect sizes for total crashes, all injury crashes, and fatal-only crashes. |  |  |  |  |  |  |
| Comments: |  |  |  |  |  |  |
| Rates, Rate Ratios, \& Standard Errors (SE) |  |  |  |  |  |  |
| Teen Group | Year(s): <br> $\square$ Pre $\square$ Post | Year(s): <br> $\square$ Pre <br> $\square$ Post | Year(s): <br> $\square$ Pre $\square$ Post | Year(s): <br> $\square$ Pre $\square$ Post | Year(s): <br> $\square$ Pre <br> $\square$ Post | Year(s): <br> $\square$ Pre <br> Post |
| Denominator \# |  |  |  |  |  |  |
| Crashes \# |  |  |  |  |  |  |



## Appendix B

## Relevance and Quality Ratings by Study

The 157 GDL Studies Identified During the Literature Search, and Relevance Screening Ratings, and States Studied

| Ref\# | Relevance rating | Authors | Year | Title |
| :---: | :--- | :--- | :--- | :--- | :--- |
| 001 | Exclude (Pre-2001) |  <br> Albrecht | 1983 | Youth License Control Demonstration Project (Report No. DOT-806-616). <br> NHTSA. |
| 002 | Exclude (Pre-2001) | McKnight | 1986 | Safety Effects of Provisional Driver Licensing Programs. In Provisional Licensing <br> Programs for Young Drivers. NHTSA. |
| 003 | Exclude (Pre-2001) | Hagge \& Marsh | PA, LA, NY, <br> MD, CA, MI |  |
| 004 | Exclude (Pre-2001) | Jones | 1988 | The traffic safety impact of provisional licensing (Report 116). California <br> Department of Motor Vehicles. |
| 005 | Exclude Not Empirical <br> Study of GDL | Foss \& Evenson | 1999 | The Effectiveness of Provisional Licensing in Oregon: An Analysis of Traffic <br> Safety Benefits. Journal of Safety Research, 25, 33-46. |
| 006 | Exclude (Pre-2001) | Effectiveness of Graduated Driver Licensing in Reducing Motor Vehicle Crashes. <br> American Journal of Preventative Medicine, 16(1S). |  |  |
| \& Williams |  |  |  |  |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 019 | Include but Combine with Ref\# 145 | Atkins, Cooper, \& Gillen | 2002 | Measuring the impact of changes in graduated licensing laws: The case of California | CA |
| 020 | Exclude Not Empirical Study of GDL | Ballesteros, \& Dischinger | 2002 | Characteristics of traffic crashes in Maryland (1996-1998): Differences among the youngest drivers. Accident Analysis and Prevention, 34, 279-284. | MD |
| 021 | Include | Bloch, Shin, \& Labin | 2002 | Does graduated driver licensing reduce drinking and driving?: An examination of California's teen driving restrictions. Paper presented at the 16th International Conference on Alcohol, Drugs and Traffic Safety. Montreal, Canada. | CA |
| 022 | Include but Combine with Ref\# 150 | Friedlander, Raleigh, \& Joyce | 2002 | Assessment of early effects of graduated driver licensing in Maryland: Impact on crashes of 16-year-old drivers. Glen Burnie, MD: Motor Vehicle Administration. | MD |
| 023 | Exclude Not Empirical Study of GDL | McKnight, Peck, \& Foss | 2002 | Graduated driver licensing: What works? Injury Prevention, 8(Supp. II), ii32-ii36. | Multiple |
| 024 | Exclude Not Empirical Study of GDL | Williams \& Ferguson | 2002 | Rationale for graduated licensing and risks it should address. Injury Prevention, 8(Supplement II), ii9-ii16. | U.S. |
| 025 | Exclude Not Crash Study | Williams, Nelson, \& Leaf | 2002 | Responses of teenagers and their parents to California's graduated licensing system. Accident Analysis and Prevention, 34, 835-842. | CA |
| 026 | Include | Coben \& McKay | 2003 | Evaluation of the effectiveness of Pennsylvania's graduated driver licensing program. Pennsylvania Department of Transportation. | PA |
| 027 | Include | Elliott \& Shope | 2003 | Use of a Bayesian change point model to estimate effects of a graduated driver's licensing program. Journal of Data Science, 1, 43-63. | MI |
| 028 | Exclude Not Empirical Study of GDL | Ferguson | 2003 | Other high-risk factors for young drivers-how graduated licensing does, doesn't, or could address them. Journal of Safety Research, 34, 71-77. |  |
| 029 | Exclude Not Empirical Study of GDL | Foss \& Goodwin | 2003 | Enhancing the effectiveness of graduated driver licensing legislation. Journal of Safety Research, 34, 79-84. |  |
| 030 | Exclude Not Empirical Study of GDL | Lin \& Fearn | 2003 | The provisional license: Nighttime and passenger restrictions—a literature review. Journal of Safety Research, 34, 51-61. |  |
| 031 | Include but Combine with Ref\# 045 | Masten \& Hagge | 2003 | Evaluation of California’s Graduated Driver Licensing Program (Report No. 205). California Department of Motor Vehicles. | CA |
| 032 | Exclude Not Empirical Study of GDL | Mayhew | 2003 | The learner's permit. Journal of Safety Research, 34, 35-43. |  |
| 033 | Exclude Not Empirical Study of GDL | Mayhew, Simpson, \& Pak | 2003 | Changes in collision rates among novice drivers during the first months of driving. Accident Analysis and Prevention, 35, 683-691. |  |
| 034 | Exclude Not Empirical Study of GDL | McCartt, Shabanova, \& Leaf | 2003 | Driving experience, crashes, and traffic citations of teenage beginning drivers. Accident Analysis and Prevention, 35, 311-320. |  |
| 035 | Exclude Not Empirical Study of GDL | McKnight \& McKnight | 2003 | Young novice drivers: Careless or clueless? Accident Analysis and Prevention, 35, 921-925. | CA \& MD |
| 036 | Exclude Not Empirical Study of GDL | McKnight \& Peck | 2003 | Graduated driver licensing and safer driving. Journal of Safety Research, 34, 85-89. | Multiple |
| 037 | Exclude Not Empirical Study of GDL | Rice, Peek-Asa, \& Kraus | 2003 | Nighttime driving, passenger transport, and injury crash rates of young drivers. Injury Prevention, 9, 245-250. | CA |
| 038 | Exclude Not Empirical Study of GDL | Shope \& Molnar | 2003 | Graduated driver licensing in the United States: evaluation results from the early programs. Journal of Safety Research, 34, 63-69. | Multiple |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 039 | Exclude Not Empirical Study of GDL | Simpson | 2003 | The evolution and effectiveness of graduated licensing. Journal of Safety Research, 34, 25-34. |  |
| 040 | Exclude Not Empirical Study of GDL | Waller | 2003 | The genesis of GDL. Journal of Safety Research, 34, 17-23. |  |
| 041 | Exclude Not Empirical Study of GDL | Williams | 2003 | Teenage drivers: Patterns of risk. Journal of Safety Research, 34, 5-15. |  |
| 042 | Exclude Not Empirical Study of GDL | Williams, Ferguson, \& Wells | 2003 | Sixteen-Year-Old Drivers in Fatal Crashes, United States, 2003, Traffic Injury Prevention, 6, 202-206. | U.S. |
| 043 | Exclude Not Crash Study | Goodwin \& Foss | 2004 | Graduated driver licensing restrictions: Awareness, compliance, and enforcement in North Carolina. Journal of Safety Research, 35, 367-374. | NC |
| 044 | Exclude Not Empirical Study of GDL | Hedlund \& Compton | 2004 | Graduated driver licensing research in 2003 and beyond. Journal of Safety Research, 35, 5-11. |  |
| 045 | Include | Masten \& Hagge | 2004 | Evaluation of California's graduated driver licensing program. Journal of Safety Research, 35, 523-535. | CA |
| 046 | Include but Combine with Ref\# 093 | Raymond, Johns, Golembiewski, Seifert, Nichols, \& Knoblauch | 2004 | Evaluation of Oregon's Graduated Driver Licensing Program: Final Report. DTNH22-99-D-15099 | OR |
| 047 | Include | Rice, Peek-Asa, \& Kraus | 2004 | Effects of the California graduated licensing program. Journal of Safety Research, 35, 375-381. | CA |
| 048 | Include | Shope \& Molnar | 2004 | Michigan's graduated driver licensing program: Evaluation of the first four years, Journal of Safety Research, 35, 337-344. | MI |
| 049 | Exclude Not U.S. Study | Adams | 2005 | Probationary and non-probationary drivers' nighttime crashes in Western Australia, 1996-2000, Journal of Safety Research, 36, 33-37. |  |
| 050 | Include | Dee, Grabowski, \& Morrisey | 2005 | Graduated driver licensing and teen traffic fatalities. Journal of Health Economics, 24, 571-589. | U.S. |
| 051 | Include | Falb | 2005 | Graduated Driver License: Iowa's Experience Since the Law’s Inception. Iowa Department of Transportation. | IA |
| 052 | Include | Fohr, Layde, \& Guse | 2005 | Graduated driver licensing in Wisconsin: Does it create safer drivers? Wisconsin Medical Journal, 104(7), 31-36. | WI |
| 053 | Exclude Not Crash Study | Hartos, Simons-Morton, Beck, \& Leaf | 2005 | Parent-imposed limits on high-risk adolescent driving: Are they stricter with graduated driver licensing? Accident Analysis and Prevention, 37, 557-562. | MD, CT |
| 054 | Exclude Not Empirical Study of GDL | Hedlund \& Compton | 2005 | Graduated driver licensing research in 2004 and 2005. Journal of Safety Research, 36, 4-14. | Multiple |
| 055 | Include | Hyde, Cook, Knight, \& Olson | 2005 | Graduated driver licensing in Utah: Is it effective? Annals of Emergency Medicine, 45(2), 147-154. | UT |
| 056 | Exclude Not Crash Study | McIntosh | 2005 | Wisconsin's Experience with the graduated driver licensing law. Wisconsin Medical Journal, 104(1), 52-56. | WI |
| 057 | Exclude Not Empirical Study of GDL | Simons-Morton, Lerner, \& Singer | 2005 | The observed effects of teenage passengers on the risk driving behavior of teenage drivers. Accident Analysis and Prevention, 37, 973-982. |  |
| 058 | Include | Chen, Baker, \& Li | 2006 | Graduated driver licensing programs and fatal crashes of 16-year-old drivers: A national evaluation. Pediatrics, 118, 56-62. | U.S. |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 059 | Exclude Not Empirical Study of GDL | Chen, Durbin, Elliot, Senserrick, \& Winston | 2006 | Child passenger injury risk in motor vehicle crashes: A comparison of nighttime and daytime driving by teenage and adult drivers. Journal of Safety Research, 37, 299-306. |  |
| 060 | Exclude Not Crash Study | Goodwin, Wells, Foss, \& Williams | 2006 | Encouraging compliance with graduated driver licensing restrictions. Journal of Safety Research, 37, 343-351. |  |
| 061 | Include but Combine with Ref\# 089 | Hallmark, Witt, \& Veneziano | 2006 | Evaluation of Iowa's graduated driver's licensing program. Iowa DOT and MTC. | IA |
| 062 | Exclude Not Empirical Study of GDL | Hedlund, Shults, \& Compton | 2006 | Graduated driver licensing and teenage driver research in 2006. Journal of Safety Research, 37, 107-121. | Multiple |
| 063 | Include but Combine with Ref\# 077 | Males | 2006 | California's graduated driver license law: Effects on older teenagers. Californian Journal of Health Promotion, 4, 207-221. | CA |
| 064 | Include but Combine with Ref\# 050 | Morrisey, Grabowski, Dee, \& Campbell | 2006 | The strength of graduated drivers license programs and fatalities among teen drivers and passengers. Accident Analysis and Prevention, 38, 135-141. | U.S. |
| 065 | Exclude; not valid citation |  |  |  |  |
| 066 | Include | Rios, Wald, Nelson, Dark, Price, \& Kellerman | 2006 | Impact of Georgia's teenage and adult driver responsibility act. Annals of Emergency Medicine, 47, 361-369. | GA |
| 067 | Exclude Not Empirical Study of GDL | Simons-Morton \& Ouimet | 2006 | Parent involvement in novice teen driving: A review of the literature. Injury Prevention, 12(Suppl I), i30-i37. |  |
| 068 | Exclude Not Empirical Study of GDL | Simons-Morton, Hartos, Leaf, \& Preusser | 2006 | The effect on teen driving outcomes of the Checkpoints Program in a statewide trial, Accident Analysis and Prevention, 38, 907-912. |  |
| 069 | Include | Willis | 2006 | Fatal Crashes Involving 16-year-old Texas Drivers Pre- and Post-GDL: Who, When, Where, and Why? College Station, TX: Texas Transportation Institute. | TX |
| 070 | Include | Zwicker, Williams, Chaudhary, \& Farmer | 2006 | Evaluation of California’s Graduated Licensing System. Arlington, VA: IIHS. | CA |
| 071 | Include but Combine with Ref\# 058 | Baker, Chen, \& Li | 2007 | Nationwide review of graduated driver licensing. Washington, DC: AAAFST | U.S. |
| 072 | Include | Chaudhary, Williams, \& Nissen | 2007 | Evaluation and compliance of passenger restrictions in a graduated drivers licensing program. Report No. DOT HS 810 781. NHTSA | CA, MA, VA |
| 073 | Exclude Not Empirical <br> Study of GDL | Ferguson, Teoh, \& McCartt | 2007 | Progress in teenage crash risk during the last decade. Journal of Safety Research, 38, 137-145. | U.S. |
| 074 | Include | Foss, Masten, \& Goodwin | 2007 | Long-term effects of Graduated Driver Licensing in North Carolina (Working Paper). Chapel Hill, NC: HSRC | NC |
| 075 | Exclude Not Empirical Study of GDL | Foss | 2007 | Improving graduated driver licensing systems: a conceptual approach and its implications. Journal of Safety Research, 38, 185-192. |  |
| 076 | Include but Combine with Ref\# 066 | Kellermann, Rios, Wald, Nelson, Dark, \& Price | 2007 | Graduated driver licensing in Georgia: The impact of the teenage and adult driver responsibility act (TADRA) (Report No. DOT HS 810 715). NHTSA. | GA |
| 077 | Include | Males | 2007 | California's graduated driver license law: Effect on teenage drivers' deaths through 2005. Journal of Safety Research, 38, 651-659. | CA |
| 078 | Exclude Not Crash Study | Margolis, Masten, \& Foss | 2007 | The effects of graduated driver licensing on hospitalization rates and charges for 16- and 17-year-olds in North Carolina. Traffic Injury Prevention, 8, 35-38. | NC |


| Ref\# | Relevance rating | Authors | Year | Title |
| :---: | :--- | :--- | :--- | :--- | :--- |
| 079 | Exclude Not Crash <br> Study | McCartt, Hellinga, \& Haire | 2007 | Age of licensure and monitoring teenagers' driving: Survey of parents of novice <br> teenage drivers. Journal of Safety Research, 38, 697-706. |
| 080 | Include |  <br> Ellis | 2007 | Effect of a graduated licensing system on motor vehicle crashes and associated <br> injuries involving drivers less than 18 years of age. Prehospital Emergency Care, <br> 11(4), 389-393. |
| 081 | Exclude Not Empirical <br> Study of GDL |  <br> Morris | 2007 | Parent-Taught Driver Education in Texas: A Comparative Evaluation (Report No. <br> DOT HS 810 760). NHTSA |
| 082 | Include | Preusser \& Tison | 2007 | GDL then and now. Journal of Safety Research, l38, 159-163. |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 097 | Exclude Not Crash Study | Campbell, Chaudhary, Saleheen, Borrup, \& Lapidus | 2009 | Does Knowledge of Teen Driving Risks and Awareness of Current Law Translate into Support for Stronger GDL Provisions? Lessons Learned from One State. Traffic Injury Prevention, 10, 320-324. | CT |
| 098 | Exclude Not Crash Study | Fell, Fisher, Voas, Blackman, \& Tippetts | 2009 | The Impact of Underage Drinking Laws on Alcohol-Related Fatal Crashes of Young Drivers. Alcoholism: Clinical and Experimental Research, 33(7), 1-12. | Multiple |
| 099 | Include | Henk \& Fette | 2009 | After GDL, what's next? The role of peer influence in reducing car crashes among young drivers. College Station, TX: Texas Transportation Institute. | TX |
| 100 | Exclude Other Reason | Karaca-Mandic \& Ridgeway | 2009 | Behavioral impact of graduated driver licensing on teenage driving risk and exposure. Journal of Health Economics, 29, 48-61. | Multiple |
| 101 | Include | Kim, Anton, \& Shearer | 2009 | Impacts of Public Policy on Safety - Graduated Driver’s License. Jefferson City, MO: Missouri Department of Transportation. | MO |
| 102 | Exclude Not Empirical Study of GDL | McCartt, Mayhew, Braitman, Ferguson, \& Simpson | 2009 | Effects of Age and Experience on Young Driver Crashes: Review of Recent Literature. Traffic Injury Prevention, 10, 209-219. |  |
| 103 | Include but Combine with Ref\# 117 | McCartt, Teoh, Fields, Braitman, \& Helinga | 2009 | Graduated Licensing Laws and Fatal Crashes of Teenage Drivers: A National Study. Arlington, VA: IIHS. | U.S. |
| 104 | Exclude Not Crash Study | Pressley, Benedicto, Trieu, Kendig, \& Barlow | 2009 | Motor Vehicle Injury, Mortality, and Hospital Charges by Strength of Graduated Driver Licensing Laws in 36 States. Journal of Trauma, Injury, Infection, and Critical Care, 67(1), S43-S53. | Multiple |
| 105 | Exclude Not Correct Ages | Traynor | 2009 | The impact of State level behavioral regulations on traffic fatality rates. Journal of Safety Research, 40, 421-426. | U.S. |
| 106 | Exclude Not Crash Study | Trempel | 2009 | Graduated driver licensing laws and insurance collision claim frequencies of teenage drivers. Arlington, VA: Highway Loss Data Institute. |  |
| 107 | Include but Combine with Ref\# 152 | UNC Highway Safety Research Center | 2009 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program. Chapel Hill, NC: HSRC |  |
| 108 | Include | Williams | 2009 | Licensing Age and Teenage Driver Crashes: A Review of the Evidence. Traffic Injury Prevention, 10, 9-15. | NJ |
| 109 | Include | Zhu, Chu, \& Li | 2009 | Effects of graduated driver licensing on licensure and traffic injury rates in Upstate New York. Accident Analysis and Prevention, 41, 531-535. | NY |
| 110 | Exclude Not Empirical Study of GDL | Centers for Disease Control \& Prevention | 2010 | Drivers Aged 16 or 17 Years Involved in Fatal Crashes — United States, 20042008. Morbidity and Mortality Weekly Report, 59(41). Available: <br> www.cdc.gov/mmwr/preview/mmwrhtml/mm5941a2.htm?s_cid=mm5941a2_w | U.S. |
| 111 | Exclude Not Empirical Study of GDL | D'Angelo, HalpernFelsher, \& Abraham | 2010 | Adolescents and driving: A position paper of the society for adolescent health and medicine: Society for adolescent health and medicine. Journal of Adolescent Health, 47(2), 212-214. |  |
| 112 | Exclude Not U.S. Study | Hinchcliff, Chapman, Ivers, Senserrick, \& Du | 2010 | Media framing of graduated licensing policy debates. Accident Analysis and Prevention, 42(4), 1283-1287. doi: 10.1016/j.aap.2010.02.005 |  |
| 113 | Exclude Not Crash Study | Hinchcliff, Ivers, Poulos, \& Senserrick | 2010 | Utilization of Research in Policymaking for Graduated Driver Licensing. American Journal of Public Health, 100(11), 2052-2058. doi: 10.2105/ajph.2009.184713 |  |
| 114 | Include but Combine with Ref\# 130 | Jiang \& Lyles | 2010 | Exposure-based assessment of the effectiveness of Michigan's graduated driver licensing nighttime driving restriction. Submitted for presentation at the 90th Annual Meeting of the Transportation Research Board and publication | MI |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 115 | Include | Karaca-Mandic \& Ridgeway | 2010 | Behavioral impact of graduated driver licensing on teenage driving risk and exposure. Journal of Health Economics, 29(1), 48-61. doi: <br> 10.1016/j.jhealeco.2009.10.002 | Multiple |
| 116 | Include | Masten \& Foss | 2010 | Long-term effect of the North Carolina graduated driver licensing system on licensed driver crash incidence: A 5-year survival analysis. Accident Analysis and Prevention, 47, 1647-1652. | NC |
| 117 | Include | McCartt, Teoh, Fields, Braitman, \& Hellinga | 2010 | Graduated licensing laws and fatal crashes of teenage drivers: A national study. Traffic Injury Prevention, 11, 240-248. | U.S. |
| 118 | Exclude Not Empirical Study of GDL | Paleti, Eluru, \& Bhat | 2010 | Examining the influence of aggressive driving behavior on driver injury severity in traffic crashes. Accident Analysis and Prevention, 42(6), 1839-1854. |  |
| 119 | Exclude Not Empirical Study of GDL | Prato, Toledo, Lotan, \& Taubman-Ben-Ari | 2010 | Modeling the behavior of novice young drivers during the first year after licensure. Accident Analysis and Prevention, 42(2), 480-486. |  |
| 120 | Exclude Not Empirical Study of GDL | Thor \& Gabler | 2010 | Assessing the residual teen crash risk factors after graduated drivers license implementation, Las Vegas, NV, United States. |  |
| 121 | Exclude Not Empirical Study of GDL | Williams \& Shults | 2010 | Graduated Driver Licensing Research, 2007-Present: A Review and Commentary, Journal of Safety Research, 41, 77-84. |  |
| 122 | Include | Williams, Chaudhary, Tefft, \& Tison | 2010 | Evaluation of New Jersey's graduated driver licensing program. Traffic Injury Prevention, 11, 1-7. | NJ |
| 123 | Exclude Not Empirical Study of GDL | Williams, Ali, \& Shults | 2010 | The Contribution of Fatal Crashes Involving Teens Transporting Teens, Traffic Injury Prevention, 11, 567-572. |  |
| 124 | Include but Combine with Ref\# 122 | Williams, Chaudhary, \& Tison | 2010 | Evaluation of New Jersey’s Graduated Driver Licensing Program. Washington, DC: AAAFTS. | NJ |
| 125 | Exclude Not Crash Study | Bernstein | 2011 | A review of automobile mortality data among novice drivers: Do objective ratings suggest a best practice methodology for graduated drivers license programs. (71), ProQuest Information \& Learning. |  |
| 126 | Exclude Not U.S. Study | Brookland \& Begg | 2011 | Adolescent, and their parents, attitudes towards graduated driver licensing and subsequent risky driving and crashes in young adulthood. Journal of Safety Research, 42, 109-115. |  |
| 127 | Exclude Not Empirical Study of GDL | Curry, Hafetz, Kallan, Winston, \& Durbin | 2011 | Prevalence of teen driver errors leading to serious motor vehicle crashes, Accident Analysis and Prevention, 43, 1285-1290. |  |
| 128 | Include | Fell, Todd, \& Voas | 2011 | A national evaluation of the nighttime and passenger restriction components of graduated driver licensing. Journal of Safety Research, 42(4), 283-290. | U.S. |
| 129 | Include | Fell, Jones, Romano, \& Voas | 2011 | An evaluation of graduated driver licensing effects on fatal crash involvements of young drivers in the United States. Traffic Injury Prevention, 12(5), 423-431. | U.S. |
| 130 | Include | Jiang \& Lyles | 2011 | Exposure-based assessment of the effectiveness of Michigan's graduated driver licensing nighttime driving restriction. Safety Science, 49(3), 484-490. | MI |
| 131 | Exclude Not Empirical Study of GDL | Knapp | 2011 | Reducing crash fatalities on rural roadways. Transportation Research Record (2213), 29-36. |  |
| 132 | Include | Lyon, Pan, \& Li | 2011 | National Evaluation of the Effect of Graduated Driver Licensing Laws on Teenager Fatality and Injury Crashes. Submitted to the Journal of Safety Research. | U.S. |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 133 | Include but Combine with Ref\# 134 | Masten | 2011 | National study of teen driver licensing systems and graduated driver licensing program core components. (Dissertation) Chapel Hill, North Carolina: University of North Carolina. | U.S. |
| 134 | Include | Masten, Foss, \& Marshall | 2011 | Graduated Driver Licensing and Fatal Crashes Involving 16- to 19-Year-Old Drivers. JAMA, 306(10), 1098-1103. | U.S. |
| 135 | Include | Morrisey \& Grabowski | 2011 | Gas prices, beer taxes and GDL programmes: effects on auto fatalities among young adults in the US. Applied Economics, 43(25), 3645-3654. | U.S. |
| 136 | Exclude Not Empirical Study of GDL | National Safety Council | 2011 | Safer Teen Driving is Goal of New Education Initiative Business Wire (English). |  |
| 137 | Exclude Other Reason | O’Brien, Foss, Goodwin, \& Masten | 2011 | Parents' opinions of supervised driving requirements in a graduated driver licensing program. University of North Carolina, Highway Safety Research Center: Chapel Hill, NC. |  |
| 138 | Exclude Not Empirical Study of GDL | Russell, Vandermeer, \& Hartling | 2011 | Graduated driver licensing for reducing motor vehicle crashes among young drivers (Review) |  |
| 139 | Exclude Not Crash Study | Scott, Monroe, Gibbs, Nichols, \& King | 2011 | Teenage driving safety and awareness of graduated drivers license programs: What you don't know can't deter you. Journal of Investigative Medicine, 59(2), 475-475. |  |
| 140 | Exclude Not U.S. Study | Taubman-Ben-Ari | 2011 | The contribution of perceived parental and familial characteristics to attitudes toward accompanied driving among young drivers. Accident Analysis and Prevention, 43(5), 1720-1729. |  |
| 141 | Include | Unknown | 2011 | Evaluating the Relative Effectiveness of Graduated Driver License Provisions: A Focus on Fatalities Involving 16 Year Olds (and 15 \& 18). Paper submitted to the Journal of Safety Research, apparently never published. | U.S. |
| 142 | Exclude Not Empirical Study of GDL | Williams | 2011 | Commentary: Graduated Licensing-Moving Forward or Standing Still? Traffic Injury Prevention, 12(3), 207-209. doi: 10.1080/15389588.2011.573353 |  |
| 143 | Exclude Not Crash Study | Williams, Braitman, \& McCartt | 2011 | Views of Parents of Teenagers About Licensing Policies: A National Survey. Traffic Injury Prevention, 12(1), 1-8. doi: 10.1080/15389588.2010.515631 |  |
| 144 | Include | Cooper, Atkins, \& Gillen | 2005 | Measuring the impact of passenger restrictions on new teenage drivers. Accident Analysis and Prevention, 34, 19-23. | CA |
| 145 | Include but Combine with Ref\# 19 | Cooper \& Gillen | 2005 | Long term impacts of California's graduated licensing law of 1998. Available: www.path.berkeley.edu/PATH/Publications/PDF/PRR 2005/PRR-2005-25.pdf. | CA |
| 146 | Exclude; not valid citation |  |  |  |  |
| 147 | Include | Neyens, Donmez, \& Boyle | 2008b | The Iowa graduated driver licensing program: effectiveness in reducing crashes in teenage drivers. Journal of Safety Research, 39(4), 383-390. | IA |
| 148 | Include but Combine with Ref\# 149 | Vanlaar, Mayhew, Marcoux, Wets, Brijs, \& Shope | 2009a | An evaluation of graduated driver licensing programs in North America: An Analysis of Relative Fatality Risks of 16, 17, 18 and 19 Year Old Drivers Using a Meta-Analytic Approach | U.S. |
| 149 | Include | Vanlaar, Mayhew, Marcoux, Wets, Brijs, \& Shope | 2009b | An evaluation of graduated driver licensing programs in North America using a meta-analytic approach. Accident Analysis and Prevention, 41, 1104-1111. | U.S. |


| Ref\# | Relevance rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 150 | Include | Freidlander, Kane, Raleigh, \& Joyce | 2004 | Reduction in young fatal and disabling driver crashes in Maryland: Assessment of three years experience with graduated licensing for 16-year-olds. Glen Burnie, MD: Motor Vehicle Administration. | OR |
| 151 | Include | Massachusetts Traffic Safety Research Program | Unkn <br> own | Review and recommendations for young driver licensing practices in Massachusetts. Amherst, MA: University of Massachusetts at Amherst. | MA |
| 152 | Include | Foss, Masten, Goodwin, \& O'Brien | 2012 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program | U.S.; Multiple |
| 153 | Exclude (Pre-2001) | Levy | 1988 | The effects of driving age, driver education, and curfew laws on traffic fatalities of 15-17 year olds. Risk Analysis, 8(4), 569-574. | 47 States (Excl. AL, MA, VT, \& DC) |
| 154 | Exclude (Pre-2001) | Preusser, Williams, Zador, \& Blomberg | 1984 | The effect of curfew laws on motor vehicle crashes. Law \& Policy, 6(1), 115-128. | $\begin{aligned} & \text { PA, LA, NY, } \\ & \text { MD } \end{aligned}$ |
| 155 | Include | Mayhew, Simpson, Singhal, \& Desmond | 2006 | Reducing the crash risk for young drivers. AAA Foundation for Traffic Safety: Washington D.C. | OR, ONT |
| 156 | Include | Eisenberg | 2003 | Evaluating the effectiveness of policies related to drunk driving. Journal of Policy Analysis and Management, 22(2), 249-274. | U.S. |
| 157 | Include | Rogers, Bentley, Campbell, Borrup, Saleheen, Wang, \& Lapidus | 2011 | Impact of Connecticut's Graduated Driver Licensing System on Teenage Motor Vehicle Crash Rates, Journal of Trauma, 71(5), S527-S530. | CT |

The 75 Studies Deemed Relevant for Possible Inclusion in the Meta-Analysis, Quality Screening Ratings, and States Studied

| Ref\# | Quality rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 010 | Low | Agent, Steenbergen, Pigman, Kidd, McCoy, \& Pollack | 2001 | Impact of the partial graduated driver's license program on teen motor vehicle crashes in Kentucky. Transportation Research Record, 1779, 54-61. Washington, D.C.: Transportation Research Board. | KY |
| 013-1 | Low | Foss, Feaganes, \& Rodgman | 2001 | Initial effects of graduated driver licensing on 16-year-old driver crashes in North Carolina. JAMA 286, 1588-1592. | NC |
| 013-2 |  | Foss | 2000 | Preliminary Evaluation of the North Carolina Graduated Driver Licensing System: Effects on Young Driver Crashes. Chapel Hill, NC: HSRC |  |
| 015 | Low | Ohio Department of Public Safety | 2001 | An Evaluation of Ohio’s Graduated Driver License <br> Law. Columbus, OH: Office of the Governor's Highway Safety Representative, Ohio Department of Public Safety. | OH |
| 017 | Low | Smith, Pierce, Ray, \& Murrin | 2001 | Motor vehicle occupant crashes among teens: Impact of the graduated licensing law in San Diego. 45th Annual Proceedings of the Association for the Advancement of Automotive Medicine, pp. 379-385. San Antonio, TX: AAAM. | CA |
| 018 | Low | Ulmer, Ferguson, Williams, \& Preusser | 2001 | Teenage crash reduction associated with delayed licensure in Connecticut. Journal of Safety Research, 32, 31-41. | CT |
| 021 | Low | Bloch, Shin, \& Labin | 2002 | Does graduated driver licensing reduce drinking and driving?: An examination of California's teen driving restrictions. Paper presented at the 16th International Conference on Alcohol, Drugs and Traffic Safety. Montreal, Canada. | CA |
| 026 | Low | Coben, \& McKay | 2003 | Evaluation of the effectiveness of Pennsylvania's graduated driver licensing program. Harrisburg, PA: Pennsylvania Department of Transportation. | PA |
| 045-1 | High | Masten \& Hagge | 2004 | Evaluation of California's graduated driver licensing program. Journal of Safety Research, 35, 523-535. | CA |
| 045-2 |  | Masten \& Hagge | 2003 | Evaluation of California’s Graduated Driver Licensing Program (Report No. 205). California Department of Motor Vehicles. |  |
| 047 | Low | Rice, Peek-Asa, \& Kraus | 2004 | Effects of the California graduated licensing program. Journal of Safety Research, 35, 375-381. | CA |
| 048-1 | Moderate | Shope \& Molnar | 2004 | Michigan's graduated driver licensing program: Evaluation of the first four years, Journal of Safety Research, 35, 337-344. | MI |
| 048-2 |  | Shope, Molnar, Elliott, \& Waller | 2001 | Graduated driver licensing in Michigan: Early impact on motor vehicle crashes among 16-year-old drivers. JAMA, 286, 1593-1598. |  |
| 048-3 |  | Elliott \& Shope | 2003 | Use of a Bayesian change point model to estimate effects of a graduated driver's licensing program. Journal of Data Science, 1, 43-63. |  |
| 050-1 | High | Dee, Grabowski, \& Morrisey | 2005 | Graduated driver licensing and teen traffic fatalities. Journal of Health Economics, 24, 571-589. | U.S. |
| 050-2 |  | Morrisey, Grabowski, Dee, Campbell | 2006 | The strength of graduated drivers license programs and fatalities among teen drivers and passengers. Accident Analysis and Prevention, 38, 135-141. |  |
| 050-3 |  | Morrisey \& Grabowski | 2011 | Gas prices, beer taxes and GDL programmes: effects on auto fatalities among young adults in the US. Applied Economics, 43(25), 3645-3654. |  |


| Ref\# | Quality rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 052 | Low | Fohr, Layde, \& Guse | 2005 | Graduated driver licensing in Wisconsin: Does it create safer drivers? Wisconsin Medical Journal, 104(7), 31-36. | WI |
| 055 | N/A | Hyde, Cook, Knight, \& Olson | 2005 | Graduated driver licensing in Utah: Is it effective? Annals of Emergency Medicine, 45(2), 147-154. | UT |
| 058-1 | Moderate | Chen, Baker, \& Li | 2006 | Graduated driver licensing programs and fatal crashes of 16-year-old drivers: A national evaluation. Pediatrics, 118, 56-62. | U.S. |
| 058-2 |  | Baker, Chen, \& Li | 2007 | Nationwide review of graduated driver licensing. AAAFST. |  |
| 066-1 | Moderate | Rios, Wald, Nelson, Dark, Price, \& Kellerman | 2006 | Impact of Georgia's teenage and adult driver responsibility act. Annals of Emergency Medicine, 47, 361-369. | GA |
| 066-2 |  | Kellermann, Rios, Wald, Nelson, Dark, \& Price | 2007 | Graduated driver licensing in Georgia: The impact of the teenage and adult driver responsibility act (TADRA) (Report No. DOT HS 810 715). NHTSA. |  |
| 069 | Low | Willis | 2006 | Fatal Crashes Involving 16-year-old Texas Drivers Pre- and Post-GDL: Who, When, Where, and Why? College Station, TX: Texas Transportation Institute. | TX |
| 070 | Low | Zwicker, Williams, Chaudhary, \& Farmer | 2006 | Evaluation of California's Graduated Licensing System. Arlington, VA: IIHS. | CA |
| 072 | High | Chaudhary, Williams, \& Nissen | 2007 | Evaluation and compliance of passenger restrictions in a graduated drivers licensing program. Report No. DOT HS 810 781. NHTSA | CA, MA, VA |
| 074 | Good | Foss, Masten, \& Goodwin | 2007 | Long-term effects of Graduated Driver Licensing in North Carolina (Working Paper). Chapel Hill, NC: HSRC | NC |
| 077-1 | Good | Males | 2007 | California's graduated driver license law: Effect on teenage drivers' deaths through 2005. Journal of Safety Research, 38, 651-659. | CA |
| 077-2 |  | Males | 2006 | California's graduated driver license law: Effects on older teenagers. Californian Journal of Health Promotion, 4, 207-221. |  |
| 080 | Low | O'Connor, Tinkoff, \& Ellis | 2007 | Effect of a graduated licensing system on motor vehicle crashes and associated injuries involving drivers less than 18 years-of-age. Prehospital Emergency Care, 11(4), 389-393. | DE |
| 082 | Low | Preusser \& Tison | 2007 | GDL then and now. Journal of Safety Research, 138, 159-163. | U.S. |
| 083-1 | Low | Raymond, Johns, Golembiewski, Seifert, Nichols, \& Knoblauch | 2007 | Evaluation of Oregon's Graduated Driver Licensing Program: Final Report (Report No. DOT HS 810 830). | OR |
| 083-2 |  | Raymond, Johns, Golembiewski, Seifert, Nichols, \& Knoblauch | 2004 | Evaluation of Oregon's Graduated Driver Licensing Program: Final Report. DTNH22-99-D-15099 |  |
| 084 | Low | René Ewing \& Associates | 2007 | Teenage Driving Study: Final Report with Executive Summary. Washington StateJoint Transportation Committee. | WA |
| 089-1 | Low | Hallmark, Veneziano, Falb, Pawlovich, \& Witt | 2008 | Evaluation of Iowa's graduated driver's licensing program. Accident Analysis and Prevention, 40, 1401-1405. | IA |
| 089-2 |  | Hallmark, Witt, \& Veneziano | 2006 | Evaluation of Iowa's graduated driver's licensing program. Iowa DOT and MTC. |  |
| 089-3 |  | Falb | 2005 | Graduated Driver License: Iowa’s Experience Since the Law’s Inception. Iowa Department of Transportation. |  |


| Ref\# | Quality rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 091 | Moderate | Kirley, Feller, Braver, \& Langenberg | 2008 | Does the Maryland graduated driver licensing law affect both 16-year-old drivers and those who share the road with them? Journal of Safety Research, 39, 295-301. | MD |
| 093 | Low | Ross | 2008 | Effectiveness of Oregon's teen licensing program. Oregon Department of Transportation. | OR |
| 099 | Low | Henk \& Fette | 2009 | After GDL, what's next? The role of peer influence in reducing car crashes among young drivers. College Station, TX: Texas Transportation Institute. | TX |
| 101 | Low | Kim, Anton, \& Shearer | 2009 | Impacts of Public Policy on Safety - Graduated Driver’s License. Jefferson City, MO: Missouri Department of Transportation. | MO |
| 108-1 | Low | Williams | 2009 | Licensing Age and Teenage Driver Crashes: A Review of the Evidence. Traffic Injury Prevention, 10, 9-15. | $\begin{aligned} & \text { NJ } \\ & \text { NJ } \end{aligned}$ |
| 108-2 |  | Williams | 2008 | Licensing Age and Teenage Driver Crashes: A Review of the Evidence, Arlington, VA: IIHS. |  |
| 109 | Low | Zhu, Chu, \& Li | 2009 | Effects of graduated driver licensing on licensure and traffic injury rates in Upstate New York. Accident Analysis and Prevention, 41, 531-535. | NY |
| 115-1 | Low | Karaca-Mandic \& Ridgeway | 2010 | Behavioral impact of graduated driver licensing on teenage driving risk and exposure. Journal of Health Economics, 29(1), 48-61. doi: 10.1016/j.jhealeco.2009.10.002 | Multiple |
| 115-2 |  | Karaca-Mandic \& Ridgeway | 2008 | Behavioral impact of graduated driver licensing on teenage driving risk and exposure. |  |
| 116 | Low | Masten \& Foss | 2010 | Long-term effect of the North Carolina graduated driver licensing system on licensed driver crash incidence: A 5-year survival analysis. Accident Analysis and Prevention, 47, 1647-1652. | NC |
| 117-1 | Low | McCartt, Teoh, Fields, Braitman, \& Hellinga | 2010 | Graduated licensing laws and fatal crashes of teenage drivers: A national study. Traffic Injury Prevention, 11, 240-248. | U.S. |
| 117-2 |  | McCartt, Teoh, Fields, Braitman, \& Hellinga | 2009 | Graduated Licensing Laws and Fatal Crashes of Teenage Drivers: A National Study. Arlington, VA: IIHS. |  |
| 122-1 | Moderate | Williams, Chaudhary, Tefft, \& Tison | 2010 | Evaluation of New Jersey's graduated driver licensing program. Traffic Injury Prevention, 11, 1-7. | NJ |
| 122-2 |  | Williams, Chaudhary, \& Tison | 2010 | Evaluation of New Jersey’s Graduated Driver Licensing Program. Washington, DC: AAAFTS. |  |
| 129-1 | Low | Fell, Jones, Romano, \& Voas | 2011 | An evaluation of graduated driver licensing effects on fatal crash involvements of young drivers in the United States. Traffic Injury Prevention, 12(5), 423-431. | U.S. |
| 129-2 |  | Fell, Todd, \& Voas | 2011 | A national evaluation of the nighttime and passenger restriction components of graduated driver licensing. Journal of Safety Research, 42(4), 283-290. |  |
| 130-1 | N/A | Jiang \& Lyles | 2011 | Exposure-based assessment of the effectiveness of Michigan's graduated driver licensing nighttime driving restriction. Safety Science, 49(3), 484-490. | MI |
| 130-2 |  | Jiang \& Lyles | 2010 | Exposure-based assessment of the effectiveness of Michigan's graduated driver licensing nighttime driving restriction. Submitted for presentation at the 90th Annual Meeting of the Transportation Research Board and publication |  |
| 132 | Low | Lyon, Pan, \& Li | 2011 | National Evaluation of the Effect of Graduated Driver Licensing Laws on Teenager Fatality and Injury Crashes. Submitted to the Journal of Safety Research. | U.S. |


| Ref\# | Quality rating | Authors | Year | Title | States studied |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 134-1 | High | Masten, Foss, \& Marshall | 2011 | Graduated Driver Licensing and Fatal Crashes Involving 16- to 19-Year-Old Drivers. JAMA, 306(10), 1098-1103. | U.S. |
| 134-2 |  | Masten | 2011 | National study of teen driver licensing systems and graduated driver licensing program core components. (Dissertation) Chapel Hill, North Carolina: University of North Carolina. |  |
| 141 | Low | Unknown | 2011 | Evaluating the Relative Effectiveness of Graduated Driver License Provisions: A Focus on Fatalities Involving 16-Year-Olds (and 15 \& 18). Unpublished. | U.S. |
| 144-1 | Low | Cooper, Atkins, \& Gillen | 2005 | Measuring the impact of passenger restrictions on new teenage drivers. Accident Analysis and Prevention, 34, 19-23. | CA |
| 144-2 |  | Atkins, Cooper, \& Gillen | 2002 | Measuring the impact of changes in graduated licensing laws: The case of California |  |
| 144-3 |  | Cooper \& Gillen | 2005 | Long term impacts of California’s graduated licensing law of 1998. Available: www.path.berkeley.edu/PATH/Publications/PDF/ PRR/2005/PRR -2005-25.pdf. |  |
| 147 | High | Neyens, Donmez, \& Boyle | 2008 | The Iowa graduated driver licensing program: effectiveness in reducing crashes in teenage drivers. Journal of Safety Research, 39(4), 383-390. | IA |
| 149-1 | Low | Vanlaar, Mayhew, Marcoux, Wets, Brijs, \& Shope | $\begin{aligned} & 2009 \\ & \text { b } \end{aligned}$ | An evaluation of graduated driver licensing programs in North America using a meta-analytic approach. Accident Analysis and Prevention, 41, 1104-1111. | U.S. |
| 149-2 |  | Vanlaar, Mayhew, Marcoux, Wets, Brijs, \& Shope | 2009a | An evaluation of graduated driver licensing programs in North America: An Analysis of Relative Fatality Risks of 16-, 17-, 18- and 19-Year-Old Drivers Using a Meta-Analytic Approach |  |
| 150-1 | Moderate | Friedlander, Kane, Raleigh, \& Joyce | 2004 | Reduction in young fatal and disabling driver crashes in Maryland: Assessment of three years experience with graduated licensing for 16-year-olds. Glen Burnie, MD: Motor Vehicle Administration. | MD |
| 150-2 |  | Friedlander, B., Raleigh, R., \& Joyce, J. | 2002 | Assessment of early effects of graduated driver licensing in Maryland: Impact on crashes of 16-year-old drivers. Glen Burnie, MD: Motor Vehicle Administration. |  |
| 151 | Low | Massachusetts Traffic Safety Research Program | Unkn own | Review and recommendations for young driver licensing practices in Massachusetts. Amherst, MA: University of Massachusetts at Amherst. | MA |
| 152-1 | High | Foss, Masten, Goodwin, \& O'Brien | 2012 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program | U.S.; Multiple |
| 152-2 |  | UNC Highway Safety Research Center | 2009 | The Role of Supervised Driving Requirements In a Graduated Driver Licensing Program. Chapel Hill, NC: HSRC |  |
| 155 | Low | Mayhew, Simpson, Singhal, \& Desmond | 2006 | Reducing the crash risk for young drivers. AAAFTS | OR, ONT |
| 156 | N/A | Eisenberg | 2003 | Evaluating the effectiveness of policies related to drunk driving. Journal of Policy Analysis and Management, 22(2), 249-274. | U.S. |
| 157 | Moderate | Rogers, Bentley, Campbell, Borrup, Saleheen, Wang, \& Lapidus | 2011 | Impact of Connecticut's Graduated Driver Licensing System on Teenage Motor Vehicle Crash Rates, Journal of Trauma, 71(5), S527-S530. | CT |

## Appendix C

## Coding Outcomes for 14 Studies Included in the Meta-Analysis

## Effect Sizes, Descriptions of GDL Program Components, and Coding Outcomes for Selected Moderator Variables for the 14 Studies Included in the Meta-Analysis

| Study authors and moderator coding | State | Age | Comparison | Outcome | Follow-up time (years) | $\begin{aligned} & \text { Rate } \\ & \text { ratio } \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Standard } \\ \text { error } \end{gathered}$ | $\begin{gathered} \text { Learner entry } \\ \text { age } \end{gathered}$ | Learner permit holding period | Supervised driving hours required | $\begin{aligned} & \text { Intermediate } \\ & \text { license entry } \\ & \text { age } \end{aligned}$ | Nighttime driving restriction | Passenger driving restriction | Unrestricted licensure entry age | Contingent advancement | $\begin{gathered} \text { Teen Licensing } \\ \text { System } \end{gathered}$ | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1. Chaudhary, Williams, \& Nissen, 2007 ${ }^{\text {a }}$ | CA | 16 | Overall GDL | F/I Driver Involvements PC | 5.5 | 0.88 | 0.06 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. |  | 17 | Yes | GDL2R | Good |
| Peer review: No |  |  |  | F/, Driver finolvements PC |  |  |  |  |  |  | 16 |  | 0 pass, $\geq 6$ mo. |  |  |  |  |
| Author: Public/Government Design: Time series |  | 16 | Overall GDL | Traffic Fatalities/Injuries PC | 5.5 | 0.90 | 0.05 | 15 | 6 mo . | 50 hr . | 16 | 12 am . | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
| Control: Other States, see note Quality: High |  | 16 | Passenger Restriction | F/I Driver Involvements PI | 5.5 | 0.93 | 0.13 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
| Validity: 75-89\% |  | 15-17 | Overall GDL | F/I Driver Involvements PC | 5.5 | 0.94 | 0.04 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 15-17 | Overall GDL | Traffic Fatalities/Injuries PC | 5.5 | 0.99 | 0.03 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
|  |  | 18-19 | Overall GDL | F/I Driver Involvements PC | 5.5 | 1.03 | 0.02 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  | MA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 16 | Overall GDL | F/I Driver Involvements PC | 5.1 | 0.82 | 0.06 | 16 | 6 mo . | 12 hr . | 16, 6 mo. | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 16 | Overall GDL | Traffic Fatalities/Injuries PC | 5.1 | 0.90 | 0.02 | 16 | 6 mo . | 12 hr . | 16, 6 mo. | 12 am . | 0 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 16 | Passenger Restriction | F/I Driver Involvements PI | 5.1 | 0.61 | 0.20 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 15-17 | Overall GDL | F/I Driver Involvements PC | 5.1 | 0.87 | 0.05 | 16 | 6 mo . | 12 hr . | 16, 6 mo. | 12 am . | 0 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 15-17 | Overall GDL | Traffic Fatalities/Injuries PC | 5.1 | 0.84 | 0.04 | 16 | 6 mo . | 12 hr . | 16, 6 mo. | 12 a.m. | 0 pass, $\geq 6$ mo. | 18 | Yes | GDL2R | Good |
|  |  | 18-19 | Overall GDL | F/I Driver Involvements PC | 5.1 | 0.95 | 0.02 | 16 | 6 mo . | 12 hr . | 16, 6 mo. | 12 a.m. | 0 pass, $\geq 6$ mo. | 18 | Yes | GDL2R | Good |
|  | VA | 16 | Overall GDL | F/I Driver Involvements PC | 2.5 | 0.73 | 0.08 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 16 | Overall GDL | Traffic Fatalities/Injuries PC | 2.5 | 0.81 | 0.06 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 am . | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 16 | Passenger Restriction | F/I Driver Involvements PI | 2.5 | 1.27 | 0.08 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 15-17 | Overall GDL | F/I Driver Involvements PC | 2.5 | 0.83 | 0.06 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 15-17 | Overall GDL | Traffic Fatalities/Injuries PC | 2.5 | 0.89 | 0.02 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 am . | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 18-19 | Overall GDL | F/I Driver Involvements PC | 2.5 | 0.99 | 0.04 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
| 2. Chen, Baker, \& Li, 2006 ${ }^{\text {b }}$ | US |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Some, see note Author: Academic |  | 16 | Learner Holding Period | Fatal Driver Involvements PC | Varies | 1.04 | 0.07 |  | $\geq 3 \mathrm{mo}$. |  |  |  |  |  |  |  |  |
| Design: Time series Control: Other States \& Ages |  | 16 | Nighttime Restriction | Fatal Driver Involvements PC | Varies | 1.02 | 0.07 |  |  |  |  | Any |  |  |  |  |  |
| Quality: Moderate <br> Validity: 60-74\% |  | 16 | Overall GDL | F/I Driver Involvements PC | Varies | 0.84 | 0.06 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
|  |  | 16 | Overall GDL | Fatal Driver Involvements PC | Varies | 0.92 | 0.05 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| 3. Dee, Grabowski, \& Morrisey, $2005^{c}$ <br> Peer review: Yes | US | 15-17 | Overall GDL | Driver Fatalities PC | Varies | 0.93 | 0.02 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| Author: Academic |  | 15-17 | Overall GDL | Traffic Fatalities PC | Varies | 0.91 | 0.02 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |


| Study authors and moderator coding | State | Age | Comparison | Outcome | Follow-up time (years) | $\begin{aligned} & \text { Rat } \\ & \text { ratio } \end{aligned}$ | $\begin{gathered} \text { Standard } \\ \text { error } \end{gathered}$ | $\underset{\text { Learner entry }}{\text { age }}$ | Learner permit holding period | Supervised driving hours required | Intermediate license entry age | Nighttime <br> driving <br> restriction | Passenger driving restriction | Unrestricted licensure entry age | Contingent advancement | Teen Licensing System | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Design: Time Series Control: Other States Quality: High |  | 18-20 | Overall GDL | Driver Fatalities PC | Varies | 0.97 | 0.02 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| Validity: $\geq 90 \%$ |  | 18-20 | Overall GDL | Traffic Fatalities PC | Varies | 0.96 | 0.02 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| 4. Foss, Masten, \& Goodwin, $2007^{\text {d }}$ | NC | 16 | Nighttime Restriction | Total Driver Involvements PI | 6.1 | 0.68 | 0.02 |  |  |  |  | 9 p.m. |  |  |  |  |  |
| Author: Academic Design: Time Series |  | 16 | Overall GDL | Total Driver Involvements PC | 6.1 | 0.61 | 0.03 | 15 | 12 mo . | None | 16 | 9 p.m. | None | 16, 6 mo. | Yes | GDL1R | Fair |
| Control: Other Age Group, see note Quality: Good |  | 16 | Overall GDL | Total Driver Involvements PLD | 6.1 | 0.93 | 0.03 | 15 | 12 mo . | None | 16 | 9 p.m. | None | 16, 6 mo. | Yes | GDL1R | Fair |
| Validity: 75-89\% |  | 16 | Overall GDL | F/I Driver Involvements PC | 6.1 | 0.53 | 0.11 | 15 | 12 mo . | None | 16 | $9 \mathrm{p} . \mathrm{m}$. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 16 | Passenger Restriction | Total Driver Involvements PC | 2.1 | 0.94 | 0.03 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 16 | Passenger Restriction | Total Driver Involvements PLD | 2.1 | 0.94 | 0.03 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 16 | Passenger Restriction | Total Driver Involvements PI | 2.1 | 0.68 | 0.03 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 16 | Passenger Restriction | F/I Driver Involvements PC | 2.1 | 1.14 | 0.06 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Nighttime Restriction | Total Driver Involvements PI | 6.1 | 0.95 | 0.02 |  |  |  |  | 9 p.m. |  |  |  |  |  |
|  |  | 17 | Overall GDL | Total Driver Involvements PC | 5.1 | 0.80 | 0.03 | 15 | 12 mo . | None | 16 | 9 p.m. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 17 | Overall GDL | Total Driver Involvements PLD | 6.1 | 0.95 | 0.02 | 15 | 12 mo . | None | 16 | $9 \mathrm{p} . \mathrm{m}$. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 17 | Overall GDL | F/I Driver Involvements PC | 5.1 | 0.63 | 0.13 | 15 | 12 mo . | None | 16 | 9 p.m. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 17 | Passenger Restriction | Total Driver Involvements PC | 2.1 | 0.96 | 0.02 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Passenger Restriction | Total Driver Involvements PLD | 2.1 | 0.95 | 0.02 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Passenger Restriction | Total Driver Involvements PI | 2.1 | 0.85 | 0.03 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Passenger Restriction | F/I Driver Involvements PC | 2.1 | 1.00 | 0.06 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 18 | Overall GDL | Total Driver Involvements PC | 4.1 | 1.00 | 0.01 | 15 | 12 mo . | None | 16 | $9 \mathrm{p} . \mathrm{m}$. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 18 | Passenger Restriction | Total Driver Involvements PC | 2.1 | 0.94 | 0.01 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 19 | Overall GDL | Total Driver Involvements PC | 4.1 | 0.97 | 0.01 | 15 | 12 mo . | None | 16 | $9 \mathrm{p} . \mathrm{m}$. | None | 16, 6 mo. | Yes | GDL1R | Fair |
|  |  | 19 | Passenger Restriction | Total Driver Involvements PC | 2.1 | 0.93 | 0.01 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
| 5. Foss, Masten, Goodwin, \& | FL |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| O'Brien, 2012 ${ }^{\text {e }}$ <br> Peer review: Some see note |  | 16 | Overall GDL | F/I Driver Involvements PC | 9.5 | 0.90 | 0.03 | 15 | 6 mo . | None | 16 | 11 p.m. | None | 18 | Yes | GDL1R | Fair |
| Author: Academic Design: Time Series |  | 17 | Overall GDL | F/I Driver Involvements PC | 9.5 | 0.98 | 0.02 | 15 | 6 mo . | None | 16 | 11 p.m. | None | 18 | Yes | GDL1R | Fair |
| Control: Other Age Group Quality: High |  | 18 | Overall GDL | F/I Driver Involvements PC | 9.5 | 0.99 | 0.02 | 15 | 6 mo . | None | 16 | 11 p.m. | None | 18 | Yes | GDL1R | Fair |
| , |  | 19 | Overall GDL | F/I Driver Involvements PC | 9.5 | 0.99 | 0.02 | 15 | 6 mo . | None | 16 | 11 p.m. | None | 18 | Yes | GDL1R | Fair |
|  | MN |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 16 | Learner Holding Period | F/I Driver Involvements PC | 5.8 | 1.00 | 0.11 |  | 6 mo . |  |  |  |  |  |  |  |  |
|  |  | 16 | Supervised Driving Hours | F/I Driver Involvements PC | 4 | 1.12 | 0.10 |  |  | 40 hr . |  |  |  |  |  |  |  |


| Study authors and moderator coding | State | Age | Comparison | Outcome | $\begin{aligned} & \text { Follow-up } \\ & \text { time } \\ & \text { (years) } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { Rate } \\ & \text { ratio } \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Standard } \\ \text { error } \end{gathered}$ | $\begin{gathered} \text { Learner entry } \\ \text { age } \end{gathered}$ | Learner permit holding period | Supervised driving hours required | Intermediate license entry age | Nighttime driving restriction | Passenger driving restriction | Unrestricted <br> licensure entry age | Contingent advancement | $\begin{gathered} \text { Teen Licensing } \\ \text { System } \end{gathered}$ | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 17 | Learner Holding Period | F/I Driver Involvements PC | 5.8 | 0.85 | 0.07 |  | 6 mo . |  |  |  |  |  |  |  |  |
|  |  | 17 | Supervised Driving Hours | F/I Driver Involvements PC | 4 | 1.08 | 0.07 |  |  | 40 hr . |  |  |  |  |  |  |  |
|  | SC | 16 | Learner Holding Period | F/I Driver Involvements PC | 6.5 | 0.93 | 0.08 |  | 3 mo . |  |  |  |  |  |  |  |  |
|  |  | 16 | Overall GDL | F/I Driver Involvements PC | 2.8 | 0.88 | 0.07 | 15 | 6 mo . | 40 hr . | 15, 6 mo. | 6 p.m. | 2 pass, $\geq 6 \mathrm{mo}$ | 16, 6 mo. | No | GDL1R | Marginal |
|  |  | 17 | Learner Holding Period | F/I Driver Involvements PC | 6.5 | 0.97 | 0.05 |  | 3 mo. |  |  |  |  |  |  |  |  |
|  |  | 17 | Overall GDL | F/I Driver Involvements PC | 2.8 | 0.97 | 0.04 | 15 | 6 mo . | 40 hr . | 15, 6 mo. | 6 p.m. | 2 pass, $\geq 6 \mathrm{mo}$ | 16, 6 mo. | No | GDL1R | Marginal |
|  |  | 18 | Overall GDL | F/I Driver Involvements PC | 2.8 | 0.98 | 0.06 | 15 | 6 mo . | 40 hr . | 15, 6 mo. | 6 p.m. | 2 pass, $\geq 6 \mathrm{mo}$ | 16, 6 mo. | No | GDL1R | Marginal |
|  |  | 19 | Overall GDL | F/I Driver Involvements PC | 2.8 | 1.00 | 0.06 | 15 | 6 mo . | 40 hr . | 15, 6 mo. | 6 p.m. | 2 pass, $\geq 6$ mo | 16, 6 mo. | No | GDL1R | Marginal |
|  | VA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 16 | Learner Entry Age | F/I Driver Involvements PC | 9.5 | 0.88 | 0.10 | 15 |  |  |  |  |  |  |  |  |  |
|  |  | 16 | Learner Holding Period | F/I Driver Involvements PC | 8.5 | 1.01 | 0.09 |  | 6 mo . |  |  |  |  |  |  |  |  |
|  |  | 16 | Overall GDL | F/I Driver Involvements PC | 3.5 | 0.78 | 0.07 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 16 | Passenger Restriction | F/I Driver Involvements PC | 6.5 | 0.98 | 0.07 |  |  |  |  |  | 3 pass, $\geq 6 \mathrm{mo}$ |  |  |  |  |
|  |  | 16 | Passenger Restriction | F/I Driver Involvements PC | 1.5 | 0.78 | 0.09 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Learner Entry Age | F/I Driver Involvements PC | 9.5 | 0.92 | 0.04 | 15 |  |  |  |  |  |  |  |  |  |
|  |  | 17 | Learner Holding Period | F/I Driver Involvements PC | 8.5 | 1.00 | 0.04 |  | 6 mo . |  |  |  |  |  |  |  |  |
|  |  | 17 | Overall GDL | F/I Driver Involvements PC | 3.5 | 0.98 | 0.03 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 17 | Passenger Restriction | F/I Driver Involvements PC | 6.5 | 1.02 | 0.03 |  |  |  |  |  | 3 pass, $\geq 6 \mathrm{mo}$ |  |  |  |  |
|  |  | 17 | Passenger Restriction | F/I Driver Involvements PC | 1.5 | 0.93 | 0.04 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 18 | Overall GDL | F/I Driver Involvements PC | 3.5 | 0.99 | 0.04 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo. | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
|  |  | 19 | Overall GDL | F/I Driver Involvements PC | 3.5 | 1.05 | 0.03 | 15, 6 mo. | 9 mo . | 40 hr . | 16, 3 mo . | $12 \mathrm{a} . \mathrm{m}$. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | No | GDL2R | Good |
| 6. Kirley, Feller, Braver, \& | MD |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Langenberg, 2008 ${ }^{f}$ <br> Peer review: Yes |  | 16 | Overall GDL | Total Driver Involvements PC | 3 | 0.82 | 0.08 | 15, 9 mo. | 4 mo . | 40 hr . | 16, 1 mo. | 12 a.m. | None | 17, 7 mo. | Yes | GDL1R | Fair |
| Author: Academic Design: Pre-Post |  | 16 | Overall GDL | Total Driver Involvements PLD | 3 | 1.09 | 0.08 | 15, 9 mo. | 4 mo. | 40 hr . | 16, 1 mo. | 12 a.m. | None | 17, 7 mo. | Yes | GDL1R | Fair |
| Control: Other Age Group Quality: Moderate |  | 16 | Overall GDL | Driver Serious Injuries PC | 3 | 0.63 | 0.22 | 15, 9 mo. | 4 mo . | 40 hr . | 16, 1 mo. | 12 am . | None | 17, 7 mo . | Yes | GDL1R | Fair |
| Validity: 60-74\% |  | 16 | Overall GDL | Driver Serious Injuries PLD | 3 | 0.83 | 0.24 | 15, 9 mo. | 4 mo. | 40 hr . | 16, 1 mo. | 12 a.m. | None | 17, 7 mo . | Yes | GDL1R | Fair |
| 7. Males, $\mathbf{2 0 0 7}^{\text {g }}$ | CA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Yes Author: Academic |  | 16 | Overall GDL | Driver Fatalities PC | 6.5 | 0.80 | 0.14 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
| Design: Time Series \& Pre-Post, see note |  | 16 | Passenger Restriction | Driver Fatalities PI | 6.5 | 0.91 | 0.21 |  |  |  |  |  | 0 pass, $\geq 6$ mo. |  |  |  |  |
| Control: Other Age Group, see note <br> Quality: Good <br> Validity: 75-89\% |  | 17 | Overall GDL | Driver Fatalities PC | 5.5 | 0.96 | 0.14 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
|  |  | 17 | Passenger Restriction | Driver Fatalities PI | 5.5 | 0.93 | 0.15 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |


| Study authors and moderator coding | State | Age | Comparison | Outcome | $\begin{gathered} \text { Follow-up } \\ \text { time } \\ \text { (years) } \\ \hline \end{gathered}$ | Rate ratio | Standard | Learner entry age | Learner permit holding period | Supervised driving hours required | Intermediate <br> license entry age | Nighttime driving restriction | Passenger driving restriction | Unrestricted licensure entry age | Contingent advancement | Teen Licensing System | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 18 | Overall GDL | Driver Fatalities PC | 4.5 | 1.24 | 0.08 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 18 | Passenger Restriction | Driver Fatalities PI | 4.5 | 0.97 | 0.11 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 19 | Overall GDL | Driver Fatalities PC | 3.5 | 1.10 | 0.06 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
|  |  | 19 | Passenger Restriction | Driver Fatalities PI | 3.5 | 1.12 | 0.12 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 16-17 | Overall GDL | Fatal Driver Involvements PC | 4.5-5.5 | 0.86 | 0.03 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 16-19 | Overall GDL | Driver Fatalities PC | 3.5 | 1.08 | 0.05 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
|  |  | 16-19 | Overall GDL | Driver Fatalities PLD | 3.5 | 1.09 | 0.06 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 18-19 | Overall GDL | Fatal Driver Involvements PC | 2.5-3.5 | 1.10 | 0.02 | 15 | 6 mo . | 50 hr . | 16 | $12 \mathrm{a} . \mathrm{m}$. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
| 8. Masten \& Hagge, 2004 ${ }^{\text {h }}$ | CA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Some, see note |  | 16 | Overall GDL | Total Driver Involvements PC | 3 | 1.02 | 0.07 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
| Design: Time Series <br> Control: Other Age Group, see note |  | 16 | Overall GDL | F/I Driver Involvements PC | 3 | 0.90 | 0.12 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
| $\begin{aligned} & \text { Quality: High } \\ & \text { Validity: } \geq 90 \% \end{aligned}$ |  | 15-17 | Nighttime Restriction | Total Driver Involvements PI | 2.5 | 0.91 | 0.05 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 15-17 | Nighttime Restriction | F/I Driver Involvements PI | 2.5 | 0.90 | 0.05 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 15-17 | Overall GDL | Total Driver Involvements PC | 3 | 1.06 | 0.05 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 15-17 | Overall GDL | F/I Driver Involvements PC | 3 | 1.05 | 0.06 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6$ mo. | 17 | Yes | GDL2R | Good |
|  |  | 15-17 | Passenger Restriction | Total Driver Involvements PI | 2.5 | 0.93 | 0.04 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 15-17 | Passenger Restriction | F/I Driver Involvements PI | 2.5 | 0.86 | 0.07 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 16-17 | Nighttime Restriction | F/I Driver Involvements PI | 2.5 | 0.90 | 0.05 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 16-17 | Passenger Restriction | F/I Driver Involvements PI | 2.5 | 0.86 | 0.07 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 18-19 | Overall GDL | Total Driver Involvements PC | 3 | 1.19 | 0.05 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
|  |  | 18-19 | Overall GDL | F/I Driver Involvements PC | 3 | 1.18 | 0.06 | 15 | 6 mo . | 50 hr . | 16 | 12 a.m. | 0 pass, $\geq 6 \mathrm{mo}$. | 17 | Yes | GDL2R | Good |
| 9. Masten, Foss, \& Marshall, $2011^{i}$ | US | 16 | Nighttime Restriction | Fatal Driver Involvements PC | Varies | 1.04 | 0.11 |  |  |  |  | 12 a.m. |  |  |  |  |  |
| Author: Academic <br> Design: Time Series |  | 16 | Overall GDL | Fatal Driver Involvements PC | Varies | 0.80 | 0.04 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| Control: Other States \& Ages Quality: High |  | 16 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 0.91 | 0.09 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
| Validity: 75-89\% |  | 16 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 0.80 | 0.05 |  |  |  |  |  | 1 pass, $\geq 6$ mo. |  |  |  |  |
|  |  | 17 | Nighttime Restriction | Fatal Driver Involvements PC | Varies | 1.02 | 0.18 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 17 | Overall GDL | Fatal Driver Involvements PC | Varies | 0.96 | 0.03 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
|  |  | 17 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 0.98 | 0.07 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 0.88 | 0.06 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |


| Study authors and moderator coding | State | Age | Comparison | Outcome | Follow-up time (years) | $\begin{aligned} & \text { Rate } \\ & \text { ratio } \end{aligned}$ | $\begin{gathered} \text { Standard } \\ \text { error } \end{gathered}$ | $\begin{gathered} \text { Learner entry } \\ \text { age } \end{gathered}$ | Learner permit holding period | Supervised driving hours required | $\begin{gathered} \text { Intermediate } \\ \text { license entry } \\ \text { age } \end{gathered}$ | Nighttime driving restriction | Passenger driving restriction | Unrestricted licensure entry age | Contingent advancement | $\begin{gathered} \text { Teen Licensing } \\ \text { System } \end{gathered}$ | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 18 | Nighttime Restriction | Fatal Driver Involvements PC | Varies | 0.81 | 0.16 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 18 | Overall GDL | Fatal Driver Involvements PC | Varies | 1.11 | 0.03 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
|  |  | 18 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 1.00 | 0.08 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 18 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 0.99 | 0.09 |  |  |  |  |  | 1 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 19 | Nighttime Restriction | Fatal Driver Involvements PC | Varies | 1.10 | 0.11 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 19 | Overall GDL | Fatal Driver Involvements PC | Varies | 1.03 | 0.03 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
|  |  | 19 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 1.01 | 0.05 |  |  |  |  |  | 0 pass, $\geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 19 | Passenger Restriction | Fatal Driver Involvements PC | Varies | 1.03 | 0.05 |  |  |  |  |  | 1 pass, $\geq 6$ mo. |  |  |  |  |
|  |  | 16-19 | Overall GDL | Fatal Driver Involvements PC | Varies | 0.98 | 0.02 | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies | Varies |
| 10. Neyens, Donmez, \& Boyle, $2008^{j}$ | IA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Yes <br> Author: Academic |  | 16 | Overall GDL | Total Driver Involvements PLD | 6.5 | 0.96 | 0.01 | 14 | 6 mo . | 30 hr . | 16 | 12:30 a.m. | None | 17 | Yes | GDL1R | Fair |
| Design: Time Series Control: Other Age Group Quality: High |  | 17 | Overall GDL | Total Driver Involvements PLD | 5.5 | 0.97 | 0.01 | 14 | 6 mo . | 30 hr . | 16 | 12:30 a.m. | None | 17 | Yes | GDL1R | Fair |
| Validity: $\geq 90 \%$ |  | 18 | Overall GDL | Total Driver Involvements PLD | 4.5 | 0.99 | 0.01 | 14 | 6 mo . | 30 hr . | 16 | 12:30 a.m. | None | 17 | Yes | GDL1R | Fair |
| 11. Rios et al., 2006 ${ }^{\text {k }}$ | GA |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Some, see note Author: Academic |  | 16 | Nighttime Restriction | Fatal Driver Involvements PI | 5.5 | 1.04 | 0.28 |  |  |  |  | 1 a.m. |  |  |  |  |  |
| Design: Pre-Post <br> Control: Other States \& Ages, see |  | 16 | Overall GDL | Fatal Driver Involvements PC | 5.5 | 0.64 | 0.09 | 15 | 12 mo . | None | 16 | 1 a.m. | 3 pass, $\geq 6 \mathrm{mo}$ | 18 | Yes | 2STAGE0R | Fair |
| note <br> Quality: Moderate |  | 17 | Overall GDL | Fatal Driver Involvements PC | 5.5 | 0.82 | 0.08 | 15 | 12 mo . | None | 16 | 1 a.m. | $3 \mathrm{pass}, \geq 6 \mathrm{mo}$ | 18 | Yes | 2STAGE0R | Fair |
| Validity: 60-74\% |  | 18 | Overall GDL | Fatal Driver Involvements PC | 5.5 | 0.98 | 0.08 | 15 | 12 mo . | None | 16 | 1 a.m. | 3 pass, $\geq 6 \mathrm{mo}$ | 18 | Yes | 2Stageor | Fair |
|  |  | 19 | Overall GDL | Fatal Driver Involvements PC | 5.5 | 0.93 | 0.08 | 15 | 12 mo . | None | 16 | 1 a.m. | $3 \mathrm{pass}, \geq 6 \mathrm{mo}$ | 18 | Yes | 2STAGE0R | Fair |
|  |  | 20 | Overall GDL | Fatal Driver Involvements PC | 5.5 | 0.94 | 0.08 | 15 | 12 mo . | None | 16 | $1 \mathrm{a} . \mathrm{m}$. | 3 pass, $\geq 6 \mathrm{mo}$ | 18 | Yes | 2STAGE0R | Fair |
| 12. Rogers et al., 2011 ${ }^{1}$ | CT |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Yes Author: Academic |  | 16 | Overall GDL | Total Driver Involvements PLD | 4 | 0.81 | 0.04 | 16 | 4 mo. | None | 16, 4 mo. | None | 0 pass, $<6$ mo. | 16, 7 mo. | No | GDL1R | Fair |
| Design: Pre-Post, see note Control: Other Age Group |  | 17 | Overall GDL | Total Driver Involvements PLD | 4 | 0.91 | 0.02 | 16 | 4 mo. | None | 16, 4 mo. | None | 0 pass, < 6 mo. | 16, 7 mo. | No | GDL1R | Fair |
| Quality: Moderate <br> Validity: 50-59\% |  | 18 | Overall GDL | Total Driver Involvements PLD | 4 | 1.02 | 0.03 | 16 | 4 mo . | None | 16, 4 mo. | None | 0 pass, < 6 mo. | 16, 7 mo. | No | GDL1R | Fair |
|  |  | 19 | Overall GDL | Total Driver Involvements PLD | 4 | 1.03 | 0.01 | 16 | 4 mo . | None | 16, 4 mo. | None | 0 pass, $<6$ mo. | 16, 7 mo. | No | GDL1R | Fair |
| 13. Shope \& Molnar, 2004 ${ }^{\text {m }}$ | MI |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Peer review: Yes <br> Author: Academic |  | 16 | Nighttime Restriction | Total Driver Involvements PI | 4 | 0.67 | 0.03 |  |  |  |  | 12 a.m. |  |  |  |  |  |
| Design: Pre-Post <br> Control: Other Age Group, see note |  | 16 | Overall GDL | Total Driver Involvements PC | 4 | 0.78 | 0.00 | 14, 9 mo. | 6 mo . | 50 hr . | 16 | 12 a.m. | None | 17 | Yes | GDL1R | Fair |
| Quality: Moderate <br> Validity: 50-59\% |  | 16 | Overall GDL | F/I Driver Involvements PC | 4 | 0.77 | 0.01 | 14, 9 mo. | 6 mo . | 50 hr . | 16 | 12 a.m. | None | 17 | Yes | GDL1R | Fair |
|  |  | 16 | Overall GDL | F/I Driver Involvements PC | 4 | 0.77 | 0.01 | 14, 9 mo. | 6 mo . | 50 hr . | 16 | 12 a.m. | None | 17 | Yes | GDL1R | Fair |
|  |  | 16 | Overall GDL | Fatal Driver Involvements PC | 4 | 0.82 | 0.11 | 14, 9 mo. | 6 mo . | 50 hr . | 16 | 12 am . | None | 17 | Yes | GDL1R | Fair |


| Study authors and moderator coding | State | Age | Comparison | Outcome | Follow-up time (years) | $\begin{aligned} & \text { Rat } \\ & \text { ratio } \end{aligned}$ | $\begin{gathered} \text { Standard } \\ \text { error } \end{gathered}$ | Learner entry age | Learner permit holding period | Supervised driving hours required | Intermediate license entry age | Nighttime driving restriction | Passenger driving restriction | Unrestricted licensure entry age | Contingent advancement | Teen Licensing System | IIHS rating |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 14. Williams, Chaudhary, Tefft, \& Tison, 2010 ${ }^{\text {n }}$ | NJ | 16 | Nighttime Restriction | Total Driver Involvements PI | 4 | 1.00 | 0.17 |  |  |  |  | 12 a.m. |  |  |  |  |  |
| Author: Public/Government |  | 16 | Nighttime Restriction | Fatal Driver Involvements PI | 6 | 0.17 | 0.88 |  |  |  |  | 12 a.m. |  |  |  |  |  |
| Control: Other Age Group, see note Quality: Moderate |  | 16 | Overall GDL | Total Driver Involvements PC | 4 | 1.00 | 0.03 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 16 | Overall GDL | F/I Driver Involvements PC | 4 | 0.91 | 0.06 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 16 | Overall GDL | Fatal Driver Involvements PC | 6 | 0.57 | 0.35 | 16 | 6 mo. | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 16 | Passenger Restriction | Fatal Driver Involvements PI | 6 | 1.49 | 0.71 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 17 | Nighttime Restriction | Total Driver Involvements PI | 4 | 0.70 | 0.03 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 17 | Nighttime Restriction | Fatal Driver Involvements PI | 6 | 0.72 | 0.35 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 17 | Overall GDL | Total Driver Involvements PC | 4 | 0.84 | 0.01 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 17 | Overall GDL | F/I Driver Involvements PC | 4 | 0.86 | 0.01 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 17 | Overall GDL | Fatal Driver Involvements PC | 6 | 0.75 | 0.11 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 17 | Passenger Restriction | Fatal Driver Involvements PI | 6 | 1.06 | 0.25 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 18 | Nighttime Restriction | Total Driver Involvements PI | 4 | 0.92 | 0.03 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 18 | Nighttime Restriction | Fatal Driver Involvements PI | 6 | 0.72 | 0.26 |  |  |  |  | 12 a.m. |  |  |  |  |  |
|  |  | 18 | Overall GDL | Total Driver Involvements PC | 4 | 0.90 | 0.01 | 16 | 6 mo . | None | 17 | 12 a.m. | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 18 | Overall GDL | F/I Driver Involvements PC | 4 | 0.90 | 0.01 | 16 | 6 mo . | None | 17 | 12 a.m. | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 18 | Overall GDL | Fatal Driver Involvements PC | 6 | 0.96 | 0.11 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 18 | Passenger Restriction | Fatal Driver Involvements PI | 6 | 0.75 | 0.25 |  |  |  |  |  | $1 \mathrm{pass}, \geq 6 \mathrm{mo}$. |  |  |  |  |
|  |  | 19 | Overall GDL | Total Driver Involvements PC | 4 | 0.99 | 0.01 | 16 | 6 mo . | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 19 | Overall GDL | F/I Driver Involvements PC | 4 | 0.99 | 0.01 | 16 | 6 mo. | None | 17 | 12 a.m. | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |
|  |  | 19 | Overall GDL | Fatal Driver Involvements PC | 6 | 1.15 | 0.11 | 16 | 6 mo . | None | 17 | 12 am . | 1 pass, $\geq 6 \mathrm{mo}$. | 18 | Yes | GDL2R | Good |







 both drivers licensed under GDL and others who never had GDL licenses.
 based on single-vehicle crashes due to data limitations.
 were obtained from the authors. Component effects estimated through Modeling/Design.
 weighted effect sizes.

 obtained from the authors in order to calculate some of the effect sizes. Component effects estimated through Modeling/Design.
${ }^{\mathrm{f}}$ Fatal involvements (alone) were not coded.
 order to calculate some of the effect sizes. Component effects estimated through proportional incidence calculated as Restricted/Non-Restricted Rate.
 the non-peer-reviewed monograph; the other effects are from the peer reviewed article. Component effects estimated through proportional incidence calculated as Restricted/Total Rate.


 meta-analysis.
${ }^{\mathrm{j}}$ Includes Neyens, Donmez, \& Boyle (2008). Additional data were obtained from the authors in order to calculate the effect sizes.
 through proportional incidence calculated as Restricted/Total Rate

 not realized until after the quality ratings were completed; in retrospect this study would have been excluded altogether.
 incidence calculated as Restricted/Total Rate.
 Restricted/Non-Restricted Rate.

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U.S. Department of Transportation
National Highway Traffic Safety Administration

